

3. Using the data in RDCHEM, the following equation was obtained by OLS:

$$\widehat{rdintens} = 2.613 + .00030 \text{ sales} - .0000000070 \text{ sales}^2$$

(.429) (.00014) (.0000000037)

$n = 32, R^2 = .1484.$

i. At what point does the marginal effect of *sales* on *rdintens* become negative?

$$\frac{\partial \widehat{rdintens}}{\partial \text{sales}} = .0003 - .000000014 \text{ sales} = 0$$

$\text{sales} = \frac{.0003}{.000000014}$
 $\text{sales} = 21,428.57$

∴ The point marginal effect become negative is ↗

ii. Would you keep the quadratic term in the model? Explain.

Yes, because $t_{\text{stat}} = \frac{\hat{\beta}_2 - 0}{\text{SE}(\hat{\beta}_2)} = \frac{-0.0000000070}{0.0000000037} = -1.89$

which is significant against alternative $H_0: \beta_1 < 0$ at 5% level
($CV \approx -1.70, df = 29$)

iii. Define *salesbil* as sales measured in billions of dollars:

***salesbil* = *sales*/1,000.** Rewrite the estimated equation with *salesbil* and *salesbil*² as the independent variables. Be sure to report standard errors and the *R*-squared. [Hint: Note that *salesbil*² = *sales*²/(1,000)².]

$$\widehat{rdintens} = 2.613 + 0.3 \text{ salesbil} - 0.0070 \text{ salesbil}^2$$

(0.429) (0.14) (0.0037)

$n = 32, R^2 = 0.1484$

iv. For the purpose of reporting the results, which equation do you prefer?

The equation from (iii) is easier to read because it contains fewer zeros, the interpretation of 2 equations is the same once the different scales are accounted for.

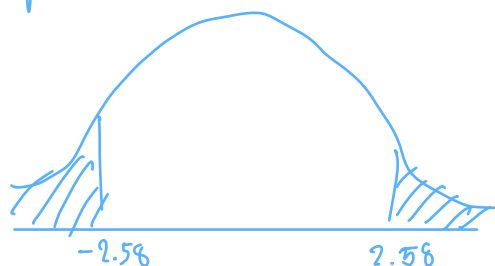
1. Using the data in SLEEP75 (see also Problem 3 in Chapter 3), we obtain the estimated equation

$$\begin{aligned}\widehat{sleep} = & 3,840.83 - .163 \text{ totwrk} - 11.71 \text{ educ} - 8.70 \text{ age} \\ & (235.11) \quad (.018) \quad (5.86) \quad (11.21) \\ & + .128 \text{ age}^2 + 87.75 \text{ male} \\ & (.134) \quad (34.33) \\ n = & 706, R^2 = .123, \bar{R}^2 = .117.\end{aligned}$$

The variable *sleep* is total minutes per week spent sleeping at night, *totwrk* is total weekly minutes spent working, *educ* and *age* are measured in years, and *male* is a gender dummy.

- All other factors being equal, is there evidence that men sleep more than women? How strong is the evidence?
- Is there a statistically significant tradeoff between working and sleeping? What is the estimated tradeoff?
- What other regression do you need to run to test the null hypothesis that, holding other factors fixed, age has no effect on sleeping?

(i) The coefficient on *male* is 87.75 so, man is estimated to sleep more per week comparable to women $t_{\text{male}} = \frac{87.75}{34.33} \approx 2.56$



close to 1% critical value so, the evidence is strong

(ii) $t_{\text{totwrk}} = \frac{-.163}{.018} \approx -9.06$ is statistically significant the coefficient implies that one more hour of work is .163 (60) or 9.8 minutes less sleep.

(iii) the null hypothesis we're testing $H_0: \hat{\beta}_3 = \hat{\beta}_4 = 0$, now run restricted version of the regression where *age*, *age*² are omitted by calculated

$$F = \frac{(R^2_{ur} - R^2_r)}{(1 - R^2_{ur}) / (n - k - 1)}$$

8. Suppose you collect data from a survey on wages, education, experience, and gender. In addition, you ask for information about marijuana usage. The original question is: "On how many separate occasions last month did you smoke marijuana?"

- i. Write an equation that would allow you to estimate the effects of marijuana usage on wage, while controlling for other factors. You should be able to make statements such as, "Smoking marijuana five more times per month is estimated to change wage by x%."
- ii. Write a model that would allow you to test whether drug usage has different effects on wages for men and women. How would you test that there are no differences in the effects of drug usage for men and women?
- iii. Suppose you think it is better to measure marijuana usage by putting people into one of four categories: nonuser, light user (1 to 5 times per month), moderate user (6 to 10 times per month), and heavy user (more than 10 times per month). Now, write a model that allows you to estimate the effects of marijuana usage on wage.
- iv. Using the model in part (iii), explain in detail how to test the null hypothesis that marijuana usage has no effect on wage. Be very specific and include a careful listing of degrees of freedom.
- v. What are some potential problems with drawing causal inference using the survey data that you collected?

$$(i) \log(\text{wage}) = \beta_0 + \beta_1 \text{usage} + \beta_2 \text{educ} + \beta_3 \text{exper} + \beta_4 \text{female}$$

$$(ii) \log(\text{wage}) = \beta_0 + \beta_1 \text{usage} + \beta_2 \text{educ} + \beta_3 \text{exper} + \beta_4 \text{female} + \beta_5 \text{usage} \cdot \text{female}$$

$$\text{To test } H_0: \beta_5 = 0$$

$$H_a: \beta_5 \neq 0$$

(iii) Assuming that there's no interaction between sex and usage

$$\log(\text{wage}) = \beta_0 + \delta_0 \text{light} + \delta_1 \text{moderate} + \delta_2 \text{heavy} + \beta_2 \text{educ} + \beta_3 \text{exper} + \beta_4 \text{female} + u$$

non-user is the omitted category

(iv) The null hypothesis is $H_0: \beta_1 = \beta_2 = \beta_3 = 0$

Perform F test which $q_3 = df = n - 6 - 1$

(v) respondents may not accurately report their marijuana usage out of fear of legal repercussions or there may be omitted variables which determine both wage and usage.

11. The following equations were estimated using the data in ECONMATH, with standard errors reported under coefficients. The average class score, measured as a percentage, is about 72.2; exactly 50% of the students are male; and the average of *colgpa* (grade point average at the start of the term) is about 2.81.

$$\begin{aligned} \widehat{score} &= 32.31 + 14.32 \text{ colgpa} \\ (2.00) \quad (0.70) \\ n &= 856, R^2 = .329, \bar{R}^2 = .328. \end{aligned}$$

$$\begin{aligned} \widehat{score} &= 29.66 + 3.83 \text{ male} + 14.57 \text{ colgpa} \\ (2.04) \quad (0.74) \quad (0.69) \\ n &= 856, R^2 = .349, \bar{R}^2 = .348. \end{aligned}$$

$$\begin{aligned} \widehat{score} &= 30.36 + 2.47 \text{ male} + 14.33 \text{ colgpa} + 0.479 \text{ male} \cdot \text{colgpa} \\ (2.86) \quad (3.96) \quad (0.98) \quad (1.383) \\ n &= 856, R^2 = .349, \bar{R}^2 = .347. \end{aligned}$$

$$\begin{aligned} \widehat{score} &= 30.36 + 3.82 \text{ male} + 14.33 \text{ colgpa} + 0.479 \text{ male} \cdot (\text{colgpa} - 2.81) \\ (2.86) \quad (0.74) \quad (0.98) \quad (1.383) \\ n &= 856, R^2 = .349, \bar{R}^2 = .347. \end{aligned}$$

i. Interpret the coefficient on *male* in the second equation and construct a 95% confidence interval for β_{male} . Does the confidence interval exclude zero?

ii. In the second equation, how come the estimate on *male* is so imprecise? Should we now conclude that there are no gender differences in *score* after controlling for *colgpa*? [Hint: You might want to compute an *F* statistic for the null hypothesis that there is no gender difference in the model with the interaction.]

iii. Compared with the third equation, how come the coefficient on *male* in the last equation is so much closer to that in the second equation and just as precisely estimated?

(i) interpretation: as the increase in score by 3.83 when 1 more male is added

CI at 95% confidence = $3.83 \pm 1.96(0.74)$

0 is included because the interval is between (2.329, 5.2804)

(ii) In equation 3 we have an interaction term among the variables so the estimate on *male* has a higher s.e.

Do F-test,

$H_0: \beta_1 = \beta_3 = 0$ in eqn 3

H_1 : otherwise

$$F = \frac{(0.349 - 0.329) / 2}{1 - 0.349 / 856} = 13.08$$

} so, we reject H_0 gender differences are significant

$$F_{(2, 852)} = 3.006$$

(iii) b/c in equation 4 variable $\text{male}(\text{colgpa} - 2.81)$ has been subtract by the mean of *colgpa* (2.81) making it closer to 0 and more precise OLS

C4. Use the data in GPA2 for this exercise.

i. Consider the equation

$$\text{colgpa} = \beta_0 + \beta_1 \text{hsize} + \beta_2 \text{hsize}^2 + \beta_3 \text{hsperc} + \beta_4 \text{sat} + \beta_5 \text{female} + \beta_6 \text{athlete} + u_i$$

where *colgpa* is cumulative college grade point average; *hsize* is size of high school graduating class, in hundreds; *hsperc* is academic percentile in graduating class; *sat* is combined SAT score; *female* is a binary gender variable; and *athlete* is a binary variable, which is one for student-athletes. What are your expectations for the coefficients in this equation? Which ones are you unsure about?

$\beta_3 < 0$ due to high school percentile is defined so that the smaller the number the better the student do.
 $\beta_4 > 0$ because the SAT score can't be negative
 Others coefficients are unclear.

ii. Estimate the equation in part (i) and report the results in the usual form.

What is the estimated GPA differential between athletes and nonathletes? Is it statistically significant?

. reg colgpa hsize hsize^2 hsperc sat female athlete

Source	SS	df	MS	Number of obs	=	4,137
Model	524.819305	6	87.4698842	F(6, 4130)	=	284.59
Residual	1269.37637	4,130	.307355053	Prob > F	=	0.0000
				R-squared	=	0.2925
				Adj R-squared	=	0.2915
Total	1794.19567	4,136	.433799728	Root MSE	=	.5544

colgpa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hsize	-.0568543	.0163513	-3.48	0.001	-.0889117 -.0247968
hsize^2	.0046754	.0022494	2.08	0.038	.0002654 .0090854
hsperc	-.0132126	.0005728	-23.07	0.000	-.0143355 -.0120896
sat	.0016464	.0000668	24.64	0.000	.0015154 .0017774
female	.1548814	.0180047	8.60	0.000	.1195826 .1901802
athlete	.1693064	.0423492	4.00	0.000	.0862791 .2523336
_cons	1.241365	.0794923	15.62	0.000	1.085517 1.397212

$$\hat{\text{colgpa}} = 1.241 - 0.0569 \text{ hsize} + 0.00468 \text{ hsize}^2 - 0.0132 \text{ hsperc} + 0.00165 \text{ sat} + 0.155 \text{ female} + 0.169 \text{ athlete}$$

(0.079) (0.0164) (0.00225) (0.0006) (0.00007) (0.018)

(0.042)

$n = 4,137$ $R^2 = .293$

An athlete is predicted to have a GPA $\approx .169$ points higher than non-athlete CETERIS PARIBUS. The t-stat = $\frac{0.169}{0.042} \approx 4.02$ is very sig.

- iii. Drop *sat* from the model and reestimate the equation. Now, what is the estimated effect of being an athlete? Discuss why the estimate is different than that obtained in part (ii).

```
. reg colgpa hsize hsize^2 hspcr female athlete
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Source	SS	df	MS	Number of obs	=	4,137
Model	338.217123	5	67.6434247	F(5, 4131)	=	191.92
Residual	1455.97855	4,131	.35245184	Prob > F	=	0.0000
				R-squared	=	0.1885
				Adj R-squared	=	0.1875
Total	1794.19567	4,136	.433799728	Root MSE	=	.59368

colgpa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hsize	-.0534038	.0175092	-3.05	0.002	-.0877313	-.0190763
hsize^2	.0053228	.0024086	2.21	0.027	.0006007	.010045
hsperc	-.0171365	.0005892	-29.09	0.000	-.0182916	-.0159814
female	.0581231	.0188162	3.09	0.002	.0212333	.095013
athlete	.0054487	.0447871	0.12	0.903	-.0823582	.0932556
_cons	3.047698	.0329148	92.59	0.000	2.983167	3.112229

The coefficient on athlete becomes ≈ 0.0034 which is not as significant as part (ii) bc we don't control SAT score

- iv. In the model from part (i), allow the effect of being an athlete to differ by gender and test the null hypothesis that there is no ceteris paribus difference between women athletes and women nonathletes.

To test hypothesis we choose female non athlete as a group base

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. reg colgpa hsize hsize^2 hspcr sat femath maleath malenonath
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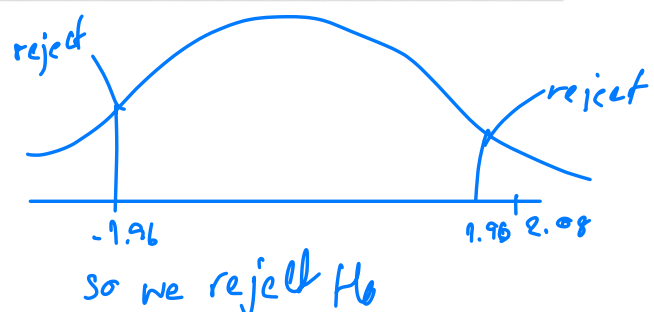
Source	SS	df	MS	Number of obs	=	4,137
Model	524.821272	7	74.9744674	F(7, 4129)	=	243.88
Residual	1269.3744	4,129	.307429015	Prob > F	=	0.0000
				R-squared	=	0.2925
				Adj R-squared	=	0.2913
Total	1794.19567	4,136	.433799728	Root MSE	=	.55446

colgpa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hsize	-.0568006	.0163671	-3.47	0.001	-.0888889	-.0247124
hsize^2	.0046699	.0022507	2.07	0.038	.0002573	.0090825
hsperc	-.0132114	.000573	-23.06	0.000	-.0143349	-.012088
sat	.0016462	.0000669	24.62	0.000	.0015151	.0017773
femath	<u>.1751106</u>	.0840258	2.08	0.037	.0103748	.3398464
maleath	.0128034	.0487395	0.26	0.793	-.0827523	.1083591
malenonath	-.1546151	.0183122	-8.44	0.000	-.1905168	-.1187133
_cons	1.39619	.0755581	18.48	0.000	1.248055	1.544324

$$H_0: \delta_i = 0$$

$$t_{.025, 4129} = 1.96$$

$$t_{cal} = \frac{.175}{.084} = 2.08$$



v. Does the effect of *sat* on *colgpa* differ by gender? Justify your answer.

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. gen femsat=female*sat
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. regress colgpa hsize hsizeq hsperc sat female athlete femsat
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Source	SS	df	MS	Number of obs =	4137
Model	524.867644	7	74.981092	F(7, 4129) =	243.91
Residual	1269.32803	4129	.307417784	Prob > F =	0.0000
Total	1794.19567	4136	.433799728	R-squared =	0.2925
				Adj R-squared =	0.2913
				Root MSE =	.55445

colgpa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hsize	-.0569121	.0163537	-3.48	0.001	-.0889741 -.0248501
hsizeq	.0046864	.0022498	2.08	0.037	.0002757 .0090972
hsperc	-.013225	.0005737	-23.05	0.000	-.0143497 -.0121003
sat	.0016255	.0000852	19.09	0.000	.0014585 .0017924
female	.1023066	.1338023	0.76	0.445	-.1600179 .3646311
athlete	.1677568	.0425334	3.94	0.000	.0843684 .2511452
femsat	.0000512	.0001291	0.40	0.692	-.000202 .0003044
_cons	1.263743	.0974952	12.96	0.000	1.0726 1.454887

So we added *female . sat* to the equation in
iii) its coefficient is about 0.000051 and *t*-stat
 $\approx .4$ there is little evidence that *sat* scores differs
by gender.