

1. Which of the following can cause the usual OLS  $t$  statistics to be invalid (that is, not to have  $t$  distributions under  $H_0$ )?

- Heteroskedasticity.
- A sample correlation coefficient of .95 between two independent variables that are in the model.
- Omitting an important explanatory variable.

Ans: (i) and (ii) cause the  $t$ -statistics not to have a distribution under  $H_0$   
 Homoskedasticity is one of the classical linear model assumpt. OLS doesn't work  
 (iii) important omitted variable violates assumption MLR3

2. Consider an equation to explain salaries of CEOs in terms of annual firm sales, return on equity ( $roe$ , in percentage form), and return on the firm's stock ( $ros$ , in percentage form):

$$\log(\text{salary}) = \beta_0 + \beta_1 \log(\text{sales}) + \beta_2 roe + \beta_3 ros + u.$$

- i. In terms of the model parameters, state the null hypothesis that, after controlling for  $sales$  and  $roe$ ,  $ros$  has no effect on CEO salary. State the alternative that better stock market performance increases a CEO's salary.

$$H_0: \beta_3 = 0$$

$$H_a: \beta_3 > 0$$

- ii. Using the data in CEOSAL1, the following equation was obtained by OLS:

$$\widehat{\log(\text{salary})} = 4.32 + .280 \log(\text{sales}) + .0174 roe + .00024 ros$$

$$(.32) \quad (.035) \quad (.0041) \quad (.00054)$$

$$n = 209, R^2 = .283.$$

By what percentage is  $salary$  predicted to increase if  $ros$  increases by 50 points? Does  $ros$  have a practically large effect on  $salary$ ?

The proportionate effect on salary =  $0.00024(50) = 0.012 = 1.2\%$

therefore, a 50 point ceteris paribus increase in  $ros$  is predicted to rise salary by 1.2%.

- iii. Test the null hypothesis that  $ros$  has no effect on  $salary$  against the alternative that  $ros$  has a positive effect. Carry out the test at the 10% significance level.

$$H_0: \beta_3 = 0$$

$$H_a: \beta_3 > 0$$

sig level 10% = 0.1

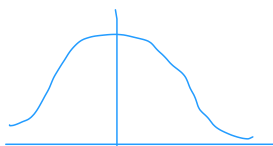
d.f. =  $209 - 5 - 1 = 205 > 30$  use Z-scores

$$t_{crit} = 1.282$$

$$t_{cal} = \frac{\hat{\beta}_3 - 0}{SE(\hat{\beta}_3)} = .44$$

$$0.44 < 1.282$$

we can't reject  $H_0$  at 10% sig level



C1. The following model can be used to study whether campaign expenditures affect election outcomes:

$$\text{vote}_A = \beta_0 + \beta_1 \log(\text{expend}_A) + \beta_2 \log(\text{expend}_B) + \beta_3 \text{prtystr}_A + u,$$

where  $\text{vote}_A$  is the percentage of the vote received by Candidate A,  $\text{expend}_A$  and  $\text{expend}_B$  are campaign expenditures by Candidates A and B, and  $\text{prtystr}_A$  is a measure of party strength for Candidate A (the percentage of the most recent presidential vote that went to A's party).

- i. What is the interpretation of  $\beta_1$ ?

$$\Delta \text{vote}_A = \beta_1 \Delta \log(\text{Expend}_A)$$

$$= (\beta_1 / 100) (100 \times \Delta \log(\text{Expend}_A))$$

$$\approx (\beta_1 / 100) [\% \Delta \log(\text{Expend}_A)]$$

$\therefore \beta_1 / 100$  is ceteris paribus percentage point change of vote received when campaign expenditure by candidate A increases by 1%.

- ii. In terms of the parameters, state the null hypothesis that a 1% increase in A's expenditures is offset by a 1% increase in B's expenditures.

$$H_0: \beta_2 = -\beta_1$$

$$H_1: \beta_2 \neq -\beta_1$$

- iii. Estimate the given model using the data in VOTE1 and report the results in usual form. Do A's expenditures affect the outcome? What about B's expenditures? Can you use these results to test the hypothesis in part (ii)?

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. reg voteA lexpendA lexpendB prtystraA
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| Source   | SS         | df  | MS         | Number of obs | = | 173    |
|----------|------------|-----|------------|---------------|---|--------|
| Model    | 38405.1096 | 3   | 12801.7032 | F(3, 169)     | = | 215.23 |
| Residual | 10052.1389 | 169 | 59.480112  | Prob > F      | = | 0.0000 |
|          |            |     |            | R-squared     | = | 0.7926 |
|          |            |     |            | Adj R-squared | = | 0.7889 |
| Total    | 48457.2486 | 172 | 281.728189 | Root MSE      | = | 7.7123 |

| voteA     | Coef.     | Std. Err. | t      | P> t  | [95% Conf. Interval] |
|-----------|-----------|-----------|--------|-------|----------------------|
| lexpendA  | 6.083316  | .38215    | 15.92  | 0.000 | 5.328914 6.837719    |
| lexpendB  | -6.615417 | .3788203  | -17.46 | 0.000 | -7.363246 -5.867588  |
| prtystraA | .1519574  | .0620181  | 2.45   | 0.015 | .0295274 .2743873    |
| _cons     | 45.07893  | 3.926305  | 11.48  | 0.000 | 37.32801 52.82985    |

regression model in usual form:  
 $voteA = (45.1) + 6.08 \log(Expand A) - 6.62 \log(Expand B) + 0.152 (prtystraA)$

- v. Estimate a model that directly gives the t statistic for testing the hypothesis in part (ii). What do you conclude? (Use a two-sided alternative.)

rewrite hypothesis  $\theta = \beta_1 + \beta_2 \Rightarrow H_0: \theta_1 = 0 \quad H_a: \theta_1 \neq 0$   
 rearrange equation  $voteA = \beta_0 + \theta \log(ExpandA) + \beta_2 [\log(ExpandB) - \log(ExpandA)] + \beta_3 prtystraA$   
 when estimate equation we obtain  $\hat{\beta}_1 \approx .533$   
 the t-stat  $\frac{-.532 - 0}{.533} \approx -1 \therefore$  can't reject  $H_0: \beta_2 = -\beta_1$

C6. Use the data in WAGE2 for this exercise.

- i. Consider the standard wage equation

$$\log(wage) = \beta_0 + \beta_1 educ + \beta_2 exper + \beta_3 tenure + u.$$

State the null hypothesis that another year of general workforce experience has the same effect on  $\log(wage)$  as another year of tenure with the current employer.

- ii. Test the null hypothesis in part (i) against a two-sided alternative, at the 5% significance level, by constructing a 95% confidence interval. What do you conclude?

```
. reg lwage educ exper tenure
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| Source   | SS         | df  | MS         | Number of obs | = | 935    |
|----------|------------|-----|------------|---------------|---|--------|
| Model    | 25.6953242 | 3   | 8.56510806 | F(3, 931)     | = | 56.97  |
| Residual | 139.960959 | 931 | .150334005 | Prob > F      | = | 0.0000 |
|          |            |     |            | R-squared     | = | 0.1551 |
|          |            |     |            | Adj R-squared | = | 0.1524 |
| Total    | 165.656283 | 934 | .177362188 | Root MSE      | = | .38773 |

| lwage  | Coef.    | Std. Err. | t     | P> t  | [95% Conf. Interval] |
|--------|----------|-----------|-------|-------|----------------------|
| educ   | .0748638 | .0065124  | 11.50 | 0.000 | .062083 .0876446     |
| exper  | .0153285 | .0033696  | 4.55  | 0.000 | .0087156 .0219413    |
| tenure | .0133748 | .0025872  | 5.17  | 0.000 | .0082974 .0184522    |
| _cons  | 5.496696 | .1105282  | 49.73 | 0.000 | 5.279782 5.713609    |

C8. The data set 401KSUBS contains information on net financial wealth (*nettfa*), age of the survey respondent (*age*), annual family income (*inc*), family size (*fsize*), and participation in certain pension plans for people in the United States. The wealth and income variables are both recorded in thousands of dollars. For this question, use only the data for single-person households (so *fsize* = 1 ).

- i. How many single-person households are there in the data set?
- ii. Use OLS to estimate the model

$$nettfa = \beta_0 + \beta_1 inc + \beta_2 age + u,$$

and report the results using the usual format. Be sure to use only the single-person households in the sample. Interpret the slope coefficients.

Are there any surprises in the slope estimates?

- iii. Does the intercept from the regression in part (ii) have an interesting meaning? Explain.
- iv. Find the *p*-value for the test  $H_0: \beta_2 = 1$  against  $H_1: \beta_2 < 1$  . Do you reject  $H_0$  at the 1% significance level?
- v. If you do a simple regression of *nettfa* on *inc*, is the estimated coefficient on *inc* much different from the estimate in part (ii)? Why or why not?

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. reg nettfa inc age if fsize ==1
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| Source   | SS         | df    | MS         | Number of obs | = | 2,017  |
|----------|------------|-------|------------|---------------|---|--------|
| Model    | 544916.989 | 2     | 272458.495 | F(2, 2014)    | = | 136.46 |
| Residual | 4021048.06 | 2,014 | 1996.54819 | Prob > F      | = | 0.0000 |
|          |            |       |            | R-squared     | = | 0.1193 |
|          |            |       |            | Adj R-squared | = | 0.1185 |
| Total    | 4565965.05 | 2,016 | 2264.86361 | Root MSE      | = | 44.683 |

| nettfa | Coef.     | Std. Err. | t      | P> t  | [95% Conf. Interval] |
|--------|-----------|-----------|--------|-------|----------------------|
| inc    | .7993167  | .0597307  | 13.38  | 0.000 | .6821762 .9164572    |
| age    | .8426563  | .0920169  | 9.16   | 0.000 | .6621982 1.023115    |
| _cons  | -43.03981 | 4.080393  | -10.55 | 0.000 | -51.04204 -35.03758  |