

2.1 Breusch-Pagan test (BP test)

To test the above hypothesis, we follow these steps.

- 1) We need to find out whether u^2 is correlated with X . This can be done by regressing

population parameter u^2 = $\delta_0 + \delta_1 X_1 + \dots + \delta_K X_K + \text{error}$

blc Then perform a joint-hypothesis testing for $u = \text{real error}$

$H_0: \delta_1 = \delta_2 = \delta_3 = \dots = \delta_K = 0$ (homosk.)

against $H_a: H_0$ is not true. So, if at least one of the $\delta \neq 0$, we already have heterosk.

- 2) We don't know the true value of u^2 . So, we use $\hat{u}^2$ instead.

BP

To perform the Breusch-Pagan Test in STATA

STATA commands (in case $k = 4$):

regress y x_1 x_2 x_3 x_4 ← estimate the main specification first.

predict u_hat , residual ← in order to predict $\hat{u}$.

generate $u_hat_sq = u_hat^2$

regress u_hat_sq x_1 x_2 x_3 x_4

** Then, check the F-statistic on the top right-hand corner of the result table.

Example: Finding the determinants of GPA.

y X_1 X_2 X_3 X_4 X_5

. regress termgpa attend priGPA final frosh soph

post-estimation command following "regress"

predict u_hat , residual

new variable name

an option to specify that we want to predict $\hat{u}$.

BP-test for heterosk.

Source	SS	df	MS			
Model	226.077541	5	45.2155081			
Residual	142.319996	674	.211157264			
Total	368.397537	679	.542558964			

Number of obs =	680
F(5, 674) =	214.13
Prob > F =	0.0000
R-squared =	0.6137
Adj R-squared =	0.6108
Root MSE =	.45952

termgpa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
attend	.046594	.0036082	12.91	0.000	.0395093	.0536787
priGPA	.5329307	.0403281	13.21	0.000	.4537468	.6121146
final	.0503197	.0040339	12.47	0.000	.0423992	.0582403
frosh	.0974307	.0560211	1.74	0.082	-.0125662	.2074276
soph	.0689273	.0467006	1.48	0.140	-.0227689	.1606236
_cons	-1.361077	.1316861	-10.34	0.000	-1.619642	-1.102513

```

. predict u_hat, residual

. generate u_hat_sq = u_hat^2

. regress u_hat_sq attend priGPA final frosh soph

. regress u_hat_sq attend priGPA final frosh soph

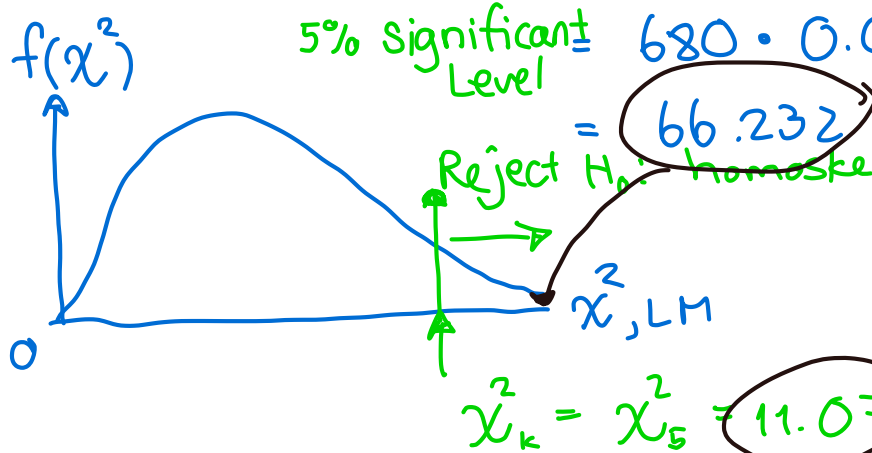
```

Source	SS	df	MS	Number of obs =	680
Model	8.22606613	5	1.64521323	F(5, 674) =	14.54
Residual	76.2624962	674	.113149104	Prob > F =	0.0000
				R-squared =	0.0974
				Adj R-squared =	0.0907
Total	84.4885623	679	.124430872	Root MSE =	.33638

u_hat_sq	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
attend	-.0088079	.0026413	-3.33	0.001	-.0139941 -.0036218
priGPA	-.1454432	.029521	-4.93	0.000	-.2034074 -.0874791
final	.0061879	.0029529	2.10	0.036	.0003899 .0119859
frosh	-.1077493	.0410085	-2.63	0.009	-.1882692 -.0272294
soph	-.0975658	.0341858	-2.85	0.004	-.1646892 -.0304423
_cons	.7368933	.0963968	7.64	0.000	.5476191 .9261674

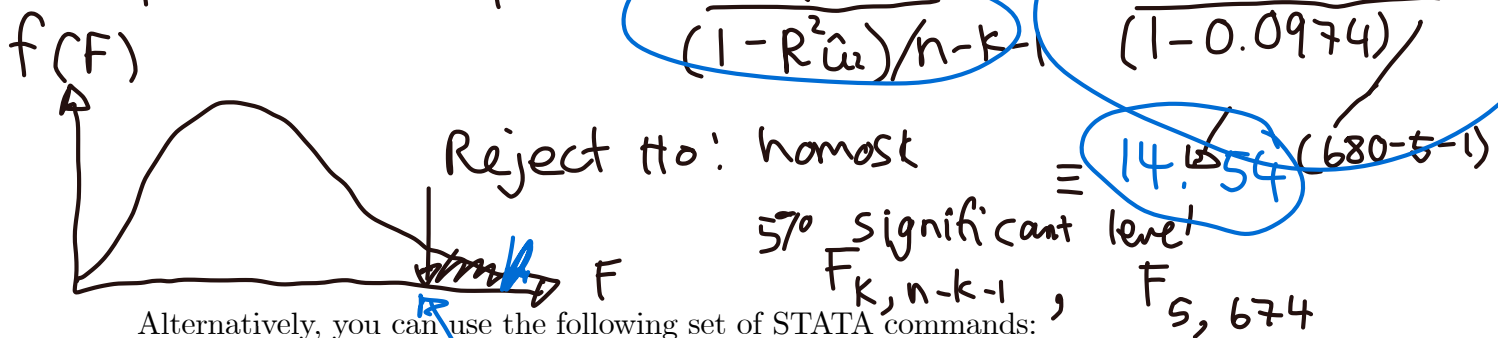
test H_0
of
 $\sigma_1^2 = \sigma_2^2 = \sigma_3^2 = \sigma_4^2 = \sigma_5^2 = 0$
 $\Rightarrow$ Test for joint sig.

LM-test : $Lm = n \cdot R_u^2$



* Since $LM > \chi^2_5$, we reject H_0 in favor of heteroskedasticity.

F-test : $F = \frac{R_u^2 / k}{(1 - R_u^2) / (n - k - 1)} = \frac{0.0974 / 5}{(1 - 0.0974) / (680 - 5 - 1)}$



Alternatively, you can use the following set of STATA commands:

```

. regress y x1 x2 x3 x4
. estat hettest x1 x2 x3 x4

```

2.21

① estimate the main regression (that we want to test for

② estat hettest

Y X₁ X₂ X₃ X₄ X₅

```
. regress termgpa attend priGPA final frosh soph
```

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Residual	142.319996	674	.211157264		Prob > F =	0.0000
					R-squared =	0.6137
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Total	368.397537	679	.542558964			

termgpa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
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priGPA	.5329307	.0403281	13.21	0.000	.4537468 .6121146
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Note:
Heteroske.
std. err.
tend to
be larger
than
homoske.
std. err.
So, we are
to reject H_0 : $\beta = 0$.

Robust

(X₁ X₂ X₃ X₄ X₅)

```
. estat hettest attend priGPA final frosh soph
```

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance → Homoskedasticity

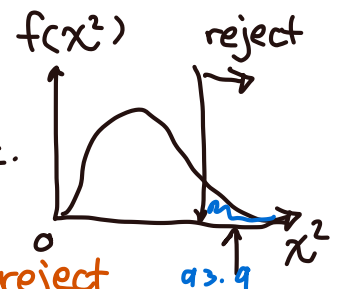
Variables: attend priGPA final frosh soph

 χ^2 d.f.

chi2(5) = 93.90

Prob > chi2 = 0.0000

the test statistic.



if p-value < 0.05, reject H_0 at 5% level of significance. And conclude that we have the heteroskedasticity problem.

If the null hypothesis is rejected (we have the heteroskedasticity problem), we can use the "robust" option in STATA. This option gives us the correct standard error, or "heteroskedasticity-robust standard error". We can now use the t-statistics in this case.

example, we don't reject H_0 : $\beta_{\text{frosh}} = 0$

```
. regress termgpa attend priGPA final frosh soph, robust
```

Linear regression

termgpa	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
attend	.046594	.0044101	10.57	0.000	.0379348 .0552532
priGPA	.5329307	.0426426	12.50	0.000	.4492023 .616659
final	.0503197	.0041066	12.25	0.000	.0422564 .058383
frosh	.0974307	.0633543	1.54	0.125	-.0269648 .2218262
soph	.0689273	.0520495	1.32	0.186	-.0332713 .1711259
_cons	-1.361077	.1448208	-9.40	0.000	-1.645431 -1.076723

Coef
Robust Std Err.

if we use hetero-robust std. err., but we reject H_0 : $\beta_{\text{frosh}} = 0$

if we use normal homoske. standard error.

So, it is

important to use the appropriate calculation for the std. err. Otherwise ⇒ we can make a wrong conclusion.

2.2 The White Test

(BP) This test is more difficult to pass than BP-test.

Similar to the Breusch-Pagan test, but is stricter because it does not allow $\hat{u}^2$ to be correlated with x^2 or interactions among different x_s .

Suppose

main regression: $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k + u$ (2)

1) like in the BP-test, we first have to obtain $\hat{u}_i$ from regression (2) $\Rightarrow$ generate $\hat{u}_i^2$ from $\hat{u}_i$

2) $\hat{u}^2 = \alpha_0 + \alpha_1 x_1 + \alpha_2 x_2 + \dots + \theta_1 x_1 x_2 + \theta_2 x_1 x_3 + \dots + \theta_{\#} x_k x_{k-1} + \gamma_1 x_1^2 + \gamma_2 x_2^2 + \dots + \gamma_k x_k^2 + \text{error}$ (3)

3) Test for $H_0: \alpha_1 = \alpha_2 = \dots = \theta_1 = \theta_2 = \dots = \gamma_1 = \gamma_2 = \dots = 0$

Homoskedasticity

H_a : H_0 is not true (we have heterosk.)

4) We can use LM-test: LM-statistic = $n \cdot R_{\hat{u}^2}^2$
 R^2 from eq. (3)

5) $LM > \text{critical value} \rightarrow$ reject H_0 : homosk.
 $LM < \text{critical value} \rightarrow$ don't reject $H_0 \rightarrow$ implies ~~heteroskedasticity~~
~~Heteroskedasticity~~
Homo
 $\chi^2_{d.f.} \Rightarrow$ # of slope parameter in (3)

The White Test (special case) (save degree of freedom)

1. Get $\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k$ by OLS. $\rightarrow$ in the main regression (2)

2. Calculate $\hat{u}_i^2 = [y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2 + \dots + \hat{\beta}_k x_k)]^2$

3. Calculate $\hat{y}_i = (\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2 + \dots + \hat{\beta}_k x_k)$

4. Calculate $\hat{y}_i^2$

5. Estimate $\hat{u}_i^2 = \gamma_0 + \gamma_1 \hat{y}_i + \gamma_2 \hat{y}_i^2 + \text{error}$ (keep R^2 of this regression)

6. $LM = nR^2$ $\rightarrow$ from this regression

7. If $p\text{-value} > \text{significance level}$, cannot reject H_0 .

$p\text{-value} < \text{sig. level} \rightarrow$ reject H_0 in favor of heteroskedasticity

3 Remedial measures

As mentioned before, there are 2 types of remedies – passive and active.

- The passive remedies just re-calculate the *std.err.* or $\hat{\beta}$ using the heteroskedasticity-robust standard error formula(s).
- The active remedies include the "weighted least squares (WLS) estimators", "generalized least squares (GLS) estimators", or "feasible GLS estimator".

3.1 Weighted Least Squares (WLS)

We assume that the heteroskedasticity may take the pattern

$$\text{Var}(u | x_1, x_2, \dots, x_k) = \sigma^2 \overbrace{h(x_1, x_2, \dots, x_k)}^{\text{a function which explains the heteros. pattern.}}$$

$\mathbf{x} \leftarrow$ vector of x

From

$$\text{Var}(u_i | \mathbf{x}) = E(u_i^2 | \mathbf{x}) - [E(u_i | \mathbf{x})]^2$$

$= E(u_i^2 | \mathbf{x})$ (since $E(u_i | \mathbf{x}) = 0$)

We get

$$\text{Var}(u_i | \mathbf{x}) = E(u_i^2 | \mathbf{x}) = \sigma^2 h_i(\mathbf{x})$$

$$\frac{E(u_i^2 | \mathbf{x})}{h_i(\mathbf{x})} = \sigma^2$$

$$E\left(\frac{u_i^2}{\sqrt{h_i}^2} | \mathbf{x}\right) = \sigma^2$$

$$E\left(\left(\frac{u_i}{\sqrt{h_i}}\right)^2 | \mathbf{x}\right) = \sigma^2$$

Therefore, even though $\text{Var}(u_i | \mathbf{x})$ is not constant,
 $\hookrightarrow$ the variance of $\frac{u_i}{\sqrt{h_i}}$ is a constant! $\nabla \nabla$

To make the error term become homoskedastic, we weight every term by $\sqrt{h_i}$.

$$\frac{y_i}{\sqrt{h_i}} = \frac{\beta_0}{\sqrt{h_i}} + \beta_1 \frac{x_{1i}}{\sqrt{h_i}} + \beta_2 \frac{x_{2i}}{\sqrt{h_i}} + \dots + \beta_k \frac{x_{ki}}{\sqrt{h_i}} + \frac{u_i}{\sqrt{h_i}}$$

$$y_i^* = \beta_0 x_{0i}^* + \beta_1 x_{1i}^* + \beta_2 x_{2i}^* + \dots + \beta_k x_{ki}^* + u_i^* \quad (1)$$

where $y_i^* = \frac{y_i}{\sqrt{h_i}}$, $x_{0i}^* = \frac{1}{\sqrt{h_i}}$, $x_{1i}^* = \frac{x_{1i}}{\sqrt{h_i}}$, etc.

estimating (1) would give unbiased $\hat{\beta}_{OLS}$
and would not require heteroskedasticity-robust
std. err. because $\underbrace{\text{Var}(u_i^* | x_i^*)}_{\text{homoskedasticity!}} = \sigma^2$

How do we find h_i or $\sqrt{h_i}$, the heteroskedasticity function?

1. If the heteroskedasticity is "known" to be caused by a multiplicative constant, we can adjust using that constant.

suppose $\text{savi}_i = \beta_0 + \beta_1 \text{inc}_i + u_i$

• We expect the heteroskedasticity pattern to vary with income (the more income the person earns the more variations in their savings.) $\rightarrow \text{Var}(u_i | x) = \sigma^2 h(x)$

So, we transform the raw data an estimate $\frac{\text{savi}_i}{\sqrt{\text{inc}_i}} = \beta_0 \frac{1}{\sqrt{\text{inc}_i}} + \beta_1 \frac{\text{inc}_i}{\sqrt{\text{inc}_i}} + \frac{u_i}{\sqrt{\text{inc}_i}}$

$$\text{Var}(u_i | \text{income}_i) = \sigma^2 \text{income}_i$$

2. If the heteroskedasticity pattern is not known, we can estimate it. This procedure is called "Feasible Generalized Least Squares" (also called Feasible GLS or FGLS)

3.2 Feasible GLS (not covered in this class.) $\text{Var}\left(\frac{u_i}{\sqrt{\text{inc}_i}} | x\right) = E\left(\left(\frac{u_i}{\sqrt{\text{inc}_i}}\right)^2 | \text{inc}_i\right) = \sigma^2$

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Since $\text{Var}(\hat{\beta}_j)$ would not be unbiased, we can make valid inferences about β_j , e.g. can use t-test, F-test, etc.