

1. Which of the following can cause the usual OLS  $t$  statistics to be invalid (that is, not to have  $t$  distributions under  $H_0$ )?

- i. Heteroskedasticity.
  - ii. A sample correlation coefficient of .95 between two independent variables that are in the model.
  - iii. Omitting an important explanatory variable.
- i. Yes, Violate the assumption of homoskedasticity*  
*ii. No, it only requires the coefficient not to be 1*  
*iii. An important omitted variable violates assumption MLR3*

2. Consider an equation to explain salaries of CEOs in terms of annual firm sales, return on equity ( $roe$ , in percentage form), and return on the firm's stock ( $ros$ , in percentage form):

$$\log(\text{salary}) = \beta_0 + \beta_1 \log(\text{sales}) + \beta_2 roe + \beta_3 ros + u.$$

i. In terms of the model parameters, state the null hypothesis that, after controlling for  $sales$  and  $roe$ ,  $ros$  has no effect on CEO salary. State the alternative that better stock market performance increases a CEO's salary.

*i.  $H_0: \beta_3 = 0$   
 $H_a: \beta_3 > 0$*

ii. Using the data in CEOSAL1, the following equation was obtained by OLS:

$$\widehat{\log(\text{salary})} = 4.32 + .280 \log(\text{sales}) + .0174 roe + .00024 ros$$

(.32) (.035) (.0041) (.00054)  
 $n = 209, R^2 = .283.$

*ii. the proportionate effect on  $\widehat{\text{salary}} = 0.00024(50) = .012 = 1.2\%$   
 therefore, a 50 point ceteris paribus increase in  $ros$  is predicted to increase salary by 1.2%.*

By what percentage is  $salary$  predicted to increase if  $ros$  increases by 50 points? Does  $ros$  have a practically large effect on  $salary$ ?

*iii.  $H_0: \beta_3 = 0$   
 $H_a: \beta_3 > 0$*

*significant level 10% = 0.1  
 $df = 209 - 3 - 1 = 205$  → z-score  
 $t_{crit} = 1.282$*

iii. Test the null hypothesis that  $ros$  has no effect on  $salary$  against the alternative that  $ros$  has a positive effect. Carry out the test at the 10% significance level.

$$t_{cal} = \frac{\hat{\beta}_3 - 0}{SE(\hat{\beta}_3)} = \frac{.0024}{.00054} = .44$$



*$0.44 < 1.282$*

*we cannot reject  $H_0$  at 10% significant level*

C1. The following model can be used to study whether campaign expenditures affect election outcomes:

$$\text{voteA} = \beta_0 + \beta_1 \log(\text{expendA}) + \beta_2 \log(\text{expendB}) + \beta_3 \text{prtystrA} + u,$$

where  $\text{voteA}$  is the percentage of the vote received by Candidate A,  $\text{expendA}$  and  $\text{expendB}$  are campaign expenditures by Candidates A and B, and  $\text{prtystrA}$  is a measure of party strength for Candidate A (the percentage of the most recent presidential vote that went to A's party).

i. What is the interpretation of  $\beta_1$ ?

$$\Delta \text{voteA} = \beta_1 \Delta \log(\text{expendA})$$

$$= (\beta_1 / 100) [100 \times \Delta \log(\text{expendA})]$$

$$\approx (\beta_1 / 100) [\% \Delta \log(\text{expendA})]$$

*∴  $\beta_1 / 100$  is the ceteris paribus percentage point change of vote received when campaign expenditure by candidate A increases by one percent*

ii. In terms of the parameters, state the null hypothesis that a 1% increase in A's expenditures is offset by a 1% increase in B's expenditures.

*$H_0: \beta_2 = -\beta_1$   
 $H_1: \beta_2 \neq -\beta_1$*

iii. Estimate the given model using the data in VOTE1 and report the results in usual form. Do A's expenditures affect the outcome? What about B's expenditures? Can you use these results to test the hypothesis in part (ii)?

. reg voteA llexpendA llexpendB prtystra

Source	SS	df	MS	Number of obs	=	173
Model	38405.1096	3	12801.7032	F(3, 169)	=	215.23
Residual	10052.1389	169	59.480112	Prob > F	=	0.0000
				R-squared	=	0.7926
				Adj R-squared	=	0.7889
				Root MSE	=	7.7123

*regression model in usual form:*

$$\text{voteA} = (45.1) + 6.08 \log(\text{ExpendA}) - 6.62 \log(\text{ExpendB}) + 0.152(\text{prtystra})$$

(11.44) (15.92) (-17.46) (2.45)

voteA	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
llexpendA	6.083316	.38215	15.92	0.000	5.328914 6.837719
llexpendB	-6.615417	.3788203	-17.46	0.000	-7.363246 -5.867588
prtystra	.1519574	.0620181	2.45	0.015	.0295274 .2743873
_cons	45.07893	3.926305	11.48	0.000	37.32801 52.82985

iv. Estimate a model that directly gives the  $t$  statistic for testing the hypothesis in part (ii). What do you conclude? (Use a two-sided alternative.)

rewrite hypothesis  $\theta_1 = \beta_1 + \beta_2 \Rightarrow H_0: \theta_1 = 0 \quad H_a: \theta_1 \neq 0$   
 rearrange equation  $\widehat{\text{voteA}} = \beta_0 + \theta \log(\text{ExpendA}) + \beta_2 [\log(\text{ExpendB} - \log(\text{ExpendA}))] + \beta_3 \text{prtystrA}$   
 when estimate equation we obtain  $\hat{\theta}_1 \approx -0.532$   
 $SE(\hat{\theta}_1) \approx 0.533$   
 then  $t_{cal} = \frac{-0.532 - 0}{0.533} \approx -1 \quad \therefore \text{cannot reject } H_0: \beta_2 = -\beta_1$

C6. Use the data in WAGE2 for this exercise.

i. Consider the standard wage equation

$$\log(\text{wage}) = \beta_0 + \beta_1 \text{educ} + \beta_2 \text{exper} + \beta_3 \text{tenure} + u.$$

State the null hypothesis that another year of general workforce experience has the same effect on  $\log(\text{wage})$  as another year of tenure with the current employer.

ii. Test the null hypothesis in part (i) against a two-sided alternative, at the 5% significance level, by constructing a 95% confidence interval. What do you conclude?

. reg lwage educ exper tenure

Source	SS	df	MS	Number of obs	=	935
Model	25.6953242	3	8.56510806	F(3, 931)	=	56.97
Residual	139.960959	931	.150334005	Prob > F	=	0.0000
				R-squared	=	0.1551
				Adj R-squared	=	0.1524
Total	165.656283	934	.177362188	Root MSE	=	.38773

  

	lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ		.0748638	.0065124	11.50	0.000	.062083 .0876446
exper		.0153285	.0033696	4.55	0.000	.0087156 .0219413
tenure		.0133748	.0025872	5.17	0.000	.0082974 .0184522
_cons		5.496696	.1105282	49.73	0.000	5.279782 5.713609

C8. The data set 401KSUBS contains information on net financial wealth (*nettfa*), age of the survey respondent (*age*), annual family income (*inc*), family size (*fsize*), and participation in certain pension plans for people in the United States. The wealth and income variables are both recorded in thousands of dollars. For this question, use only the data for single-person households (so *fsize* = 1).

i. How many single-person households are there in the data set?

ii. Use OLS to estimate the model

$$\text{nettfa} = \beta_0 + \beta_1 \text{inc} + \beta_2 \text{age} + u,$$

and report the results using the usual format. Be sure to use only the single-person households in the sample. Interpret the slope coefficients. Are there any surprises in the slope estimates?

iii. Does the intercept from the regression in part (ii) have an interesting meaning? Explain.

iv. Find the  $p$ -value for the test  $H_0: \beta_2 = 1$  against  $H_1: \beta_2 < 1$ . Do you reject  $H_0$  at the 1% significance level?

v. If you do a simple regression of *nettfa* on *inc*, is the estimated coefficient on *inc* much different from the estimate in part (ii)? Why or why not?

. reg nettfa inc age if fsize ==1

Source	SS	df	MS	Number of obs	=	2,017
Model	544916.989	2	272458.495	F(2, 2014)	=	136.46
Residual	4021048.06	2,014	1996.54819	Prob > F	=	0.0000
				R-squared	=	0.1193
				Adj R-squared	=	0.1185
Total	4565965.05	2,016	2264.86361	Root MSE	=	44.683

  

	nettfa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
inc		.7993167	.0597307	13.38	0.000	.6821762 .9164572
age		.8426563	.0920169	9.16	0.000	.6621982 1.023115
_cons		-43.03981	4.080393	-10.55	0.000	-51.04204 -35.03758