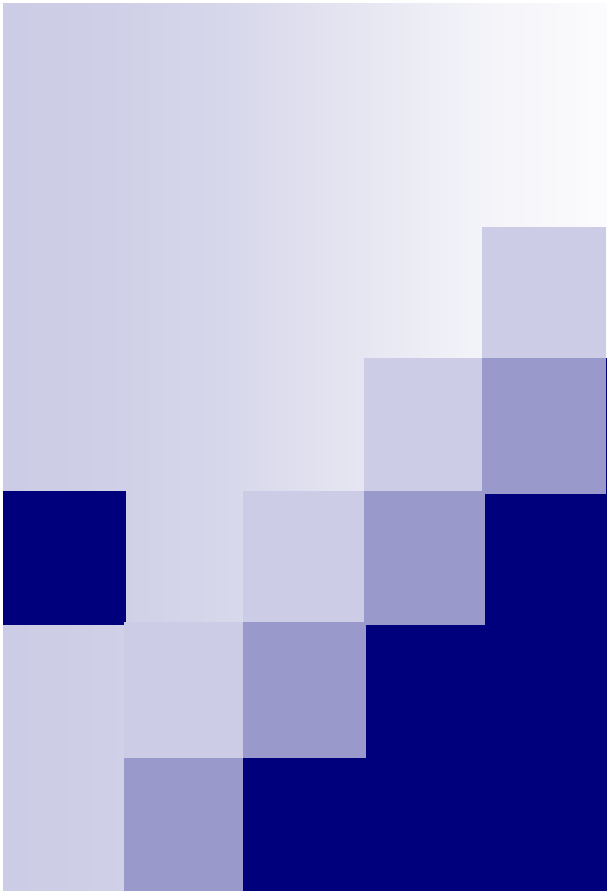
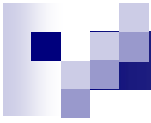




EE 325

# Two-Variable Regression Model: The Problem of Estimation



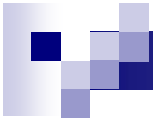
The Two-Variable PRF:

$$Y_i = \beta_1 + \beta_2 X_i + u_i$$

The Two-Variable SRF:

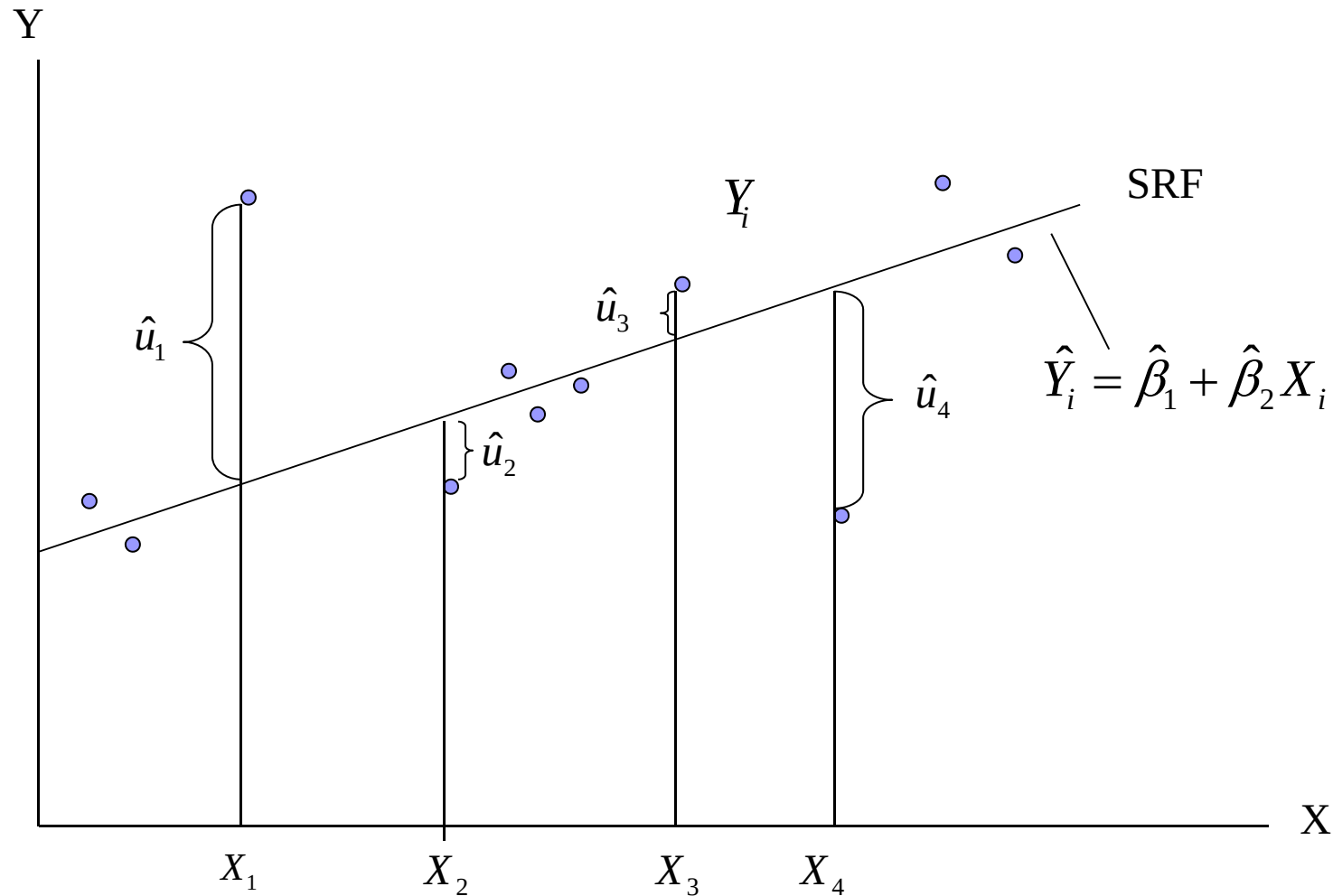
$$\begin{aligned} Y_i &= \hat{\beta}_1 + \hat{\beta}_2 X_i + \hat{u}_i \\ &= \hat{Y}_i + \hat{u}_i \end{aligned}$$

$\hat{Y}_i$  is the estimated (conditional mean) value of  $Y_i$



$$\begin{aligned}\hat{u}_i &= Y_i - \hat{Y}_i \\ &= Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i\end{aligned}$$

# Ordinary Least Squares (OLS)

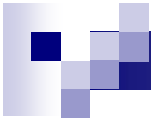




## Ordinary Least Squares (OLS)

$$\begin{aligned}\sum \hat{u}_i^2 &= \sum (Y_i - \hat{Y}_i)^2 \\ &= \sum (Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i)^2\end{aligned}$$

The principle or the method of least squares chooses  $\hat{\beta}_1$  and  $\hat{\beta}_2$  in such a manner that, for a given sample or set of data,  $\sum \hat{u}_i^2$  is as small as possible



$$\frac{\partial \left( \sum \hat{u}_i^2 \right)}{\partial \hat{\beta}_1} = -2 \sum \left( Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i \right) = -2 \sum \hat{u}_i$$

$$\frac{\partial \left( \sum \hat{u}_i^2 \right)}{\partial \hat{\beta}_2} = -2 \sum \left( Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i \right) X_i = -2 \sum \hat{u}_i X_i$$

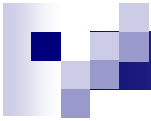
Setting these equation to zero



$$\sum Y_i = n\hat{\beta}_1 + \hat{\beta}_2 \sum X_i$$

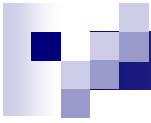
$$\sum Y_i X_i = \hat{\beta}_1 \sum X_i + \hat{\beta}_2 \sum X_i^2$$

Where n is the sample size. These simultaneous equations are known as the **Normal Equation**

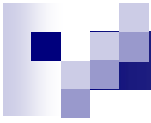


$$\begin{aligned}\hat{\beta}_2 &= \frac{n \sum X_i Y_i - \sum X_i \sum Y_i}{n \sum X_i^2 - (\sum X_i)^2} \\ &= \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2} \\ &= \frac{\sum x_i y_i}{\sum x_i^2}\end{aligned}$$





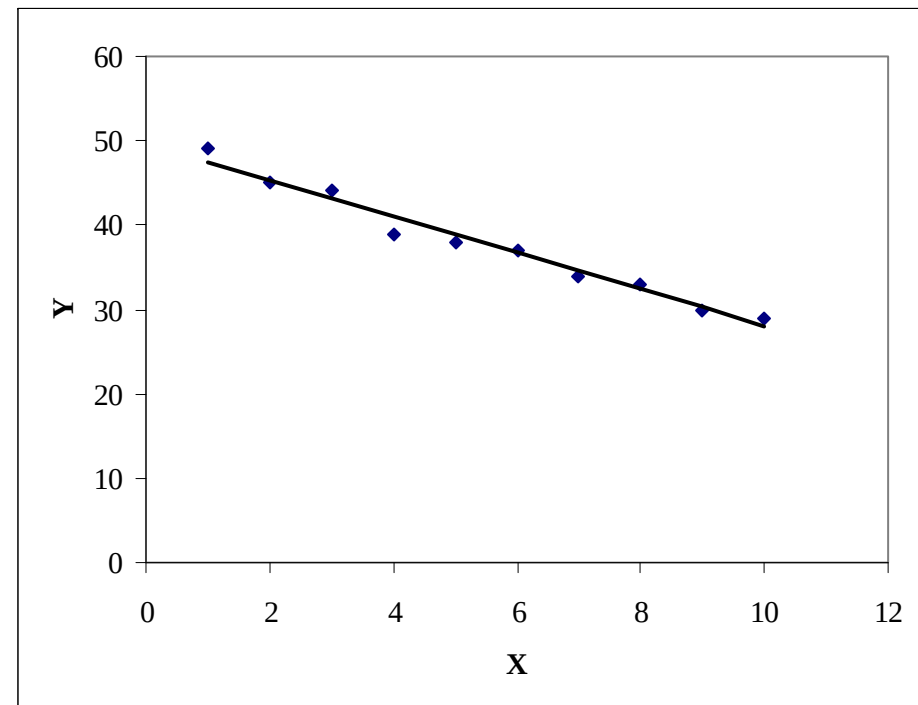
$$\begin{aligned}\hat{\beta}_2 &= \frac{\sum x_i y_i}{\sum x_i^2} \\ &= \frac{\sum x_i Y_i}{\sum X_i^2 - n\bar{X}^2} \\ &= \frac{\sum X_i y_i}{\sum X_i^2 - n\bar{X}^2}\end{aligned}$$

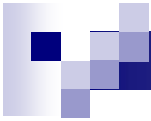


$$\begin{aligned}\hat{\beta}_1 &= \frac{\sum X_i^2 \sum Y_i - \sum X_i \sum X_i Y_i}{n \sum X_i^2 - (\sum X_i)^2} \\ &= \bar{Y} - \hat{\beta}_2 \bar{X}\end{aligned}$$

# Example 😊

| Y  | X |
|----|---|
| 49 | 1 |
| 45 | 2 |
| 44 | 3 |
| 39 | 4 |
| 38 | 5 |
| 37 | 6 |
| 34 | 7 |
| 33 | 8 |
| 30 | 9 |





$$\hat{\beta}_2 = \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2} = \frac{-178}{82.5} \approx -2.1576$$

$$\hat{\beta}_1 = \bar{Y} - \hat{\beta}_2 \bar{X} = 37.8 - (-2.1576)(5.5) = 49.667$$

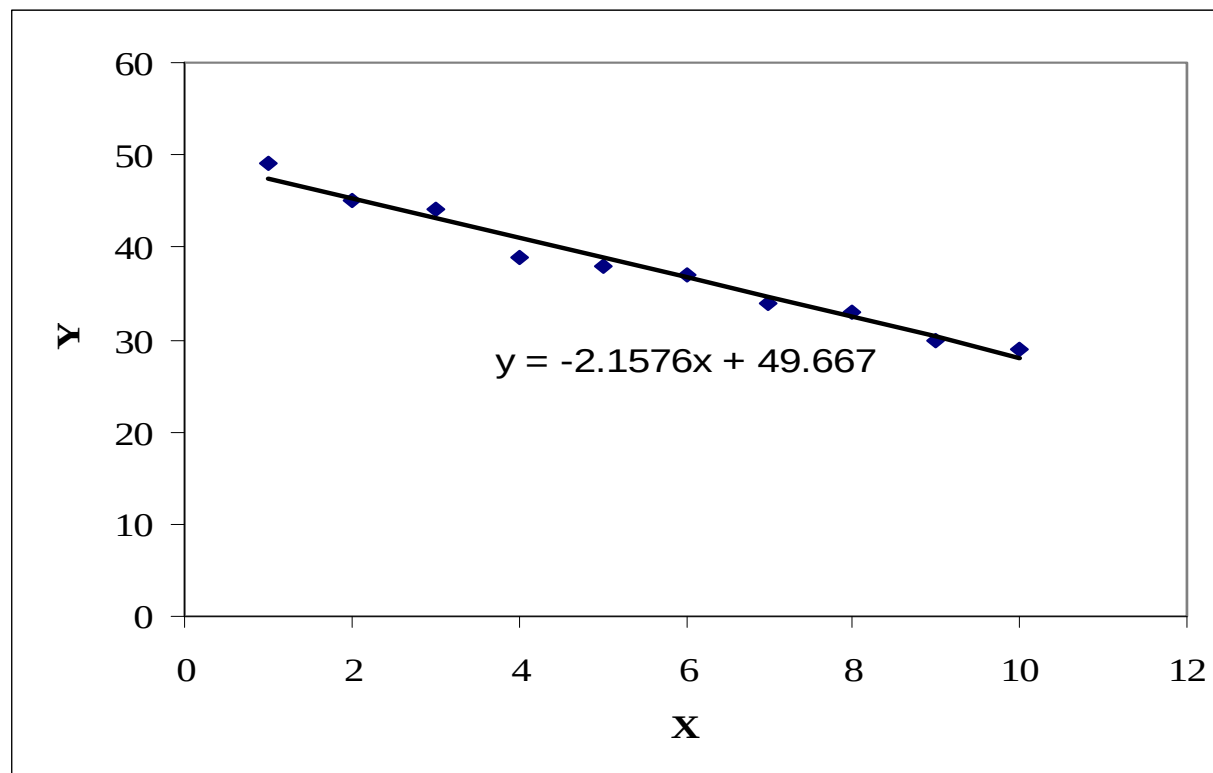
$$\bar{X} = 5.5$$

$$\bar{Y} = 37.8$$



| Y  | X  | $(X_i - \bar{X})$ | $(X_i - \bar{X})^2$ | $Y_i - \bar{Y}$ | $(X_i - \bar{X})(Y_i - \bar{Y})$ |
|----|----|-------------------|---------------------|-----------------|----------------------------------|
| 49 | 1  |                   |                     |                 |                                  |
| 45 | 2  |                   |                     |                 |                                  |
| 44 | 3  |                   |                     |                 |                                  |
| 39 | 4  |                   |                     |                 |                                  |
| 38 | 5  |                   |                     |                 |                                  |
| 37 | 6  |                   |                     |                 |                                  |
| 34 | 7  |                   |                     |                 |                                  |
| 33 | 8  |                   |                     |                 |                                  |
| 30 | 9  |                   |                     |                 |                                  |
| 29 | 10 |                   |                     |                 |                                  |


$$\hat{Y}_i = 49.667 - 2.1576X_i$$





# Practice I

A Random Sample from the Population  
of Table 2.1

---

| Weekly<br>consumption<br>expenditure \$ (Y) | Weekly income<br>\$(X) |
|---------------------------------------------|------------------------|
| 70                                          | 80                     |
| 65                                          | 100                    |
| 90                                          | 120                    |
| 95                                          | 140                    |
| 110                                         | 160                    |
| 115                                         | 180                    |
| 120                                         | 200                    |
| 140                                         | 220                    |
| 155                                         | 240                    |
| 150                                         | 260                    |

---

$$\hat{Y} = 0.5091X_i + 24.455$$

# Practice II

Another Random Sample from the  
Population of Table 2.1

---

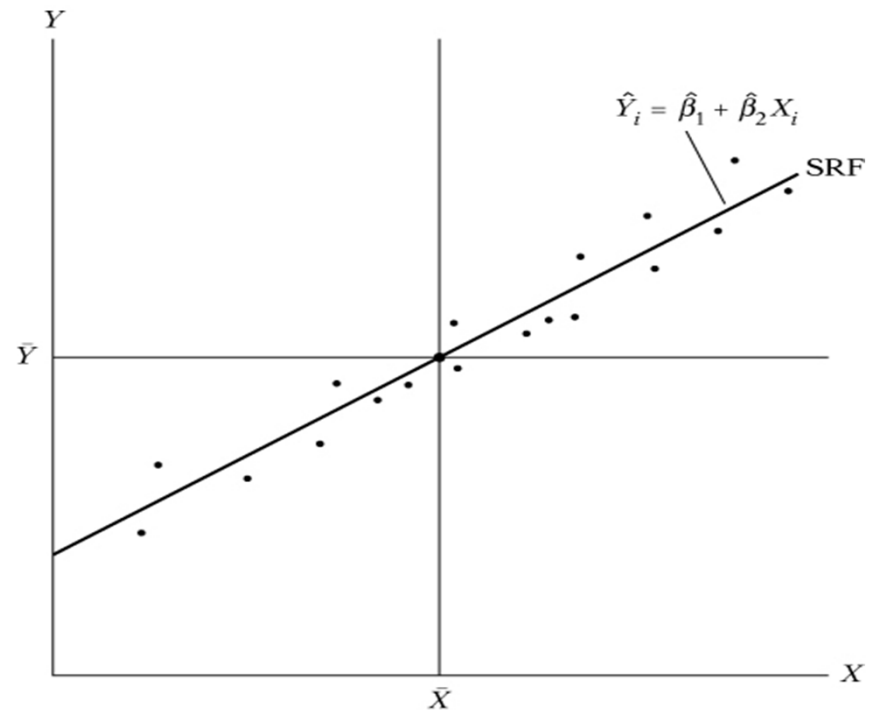
| Weekly<br>consumption<br>expenditure \$ (Y) | Weekly income<br>\$(X) |
|---------------------------------------------|------------------------|
| Y                                           | X                      |
| 55                                          | 80                     |
| 88                                          | 100                    |
| 90                                          | 120                    |
| 80                                          | 140                    |
| 118                                         | 160                    |
| 120                                         | 180                    |
| 145                                         | 200                    |
| 135                                         | 220                    |
| 145                                         | 240                    |
| 175                                         | 260                    |

$$\hat{Y} = 0.5761X_i + 17.17$$



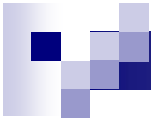
The regression line has the following properties:

- Passes through the sample means of Y and X



- 
- The mean value of the estimated  $Y = \hat{Y}_i$  is equal to the mean value of the actual Y for

$$\hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$$



- The mean value of the residuals  $\hat{u}_i$  is zero
- The residuals  $\hat{u}_i$  are uncorrelated with the predicted  $\hat{y}_i$
- The residuals  $\hat{u}_i$  are uncorrelated with  $X_i$



# Classical Linear regression model (CLRM)

The assumptions underlying the method of least squares:

1. Linear regression model
2. Fixed X values or X values independent of the error term
3. Zero mean value of disturbance
4. Homoscedasticity or Constant Variance of  $u_i$
5. No autocorrelation between the disturbances  $u_i$
6. The number of observations n must be greater than the number of parameters to be estimated
7. The nature of X variables



# Linear regression model

$$Y_i = \beta_1 + \beta_2 X_i + u_i$$

- Linear in the parameters
- May or may not be linear in the variables



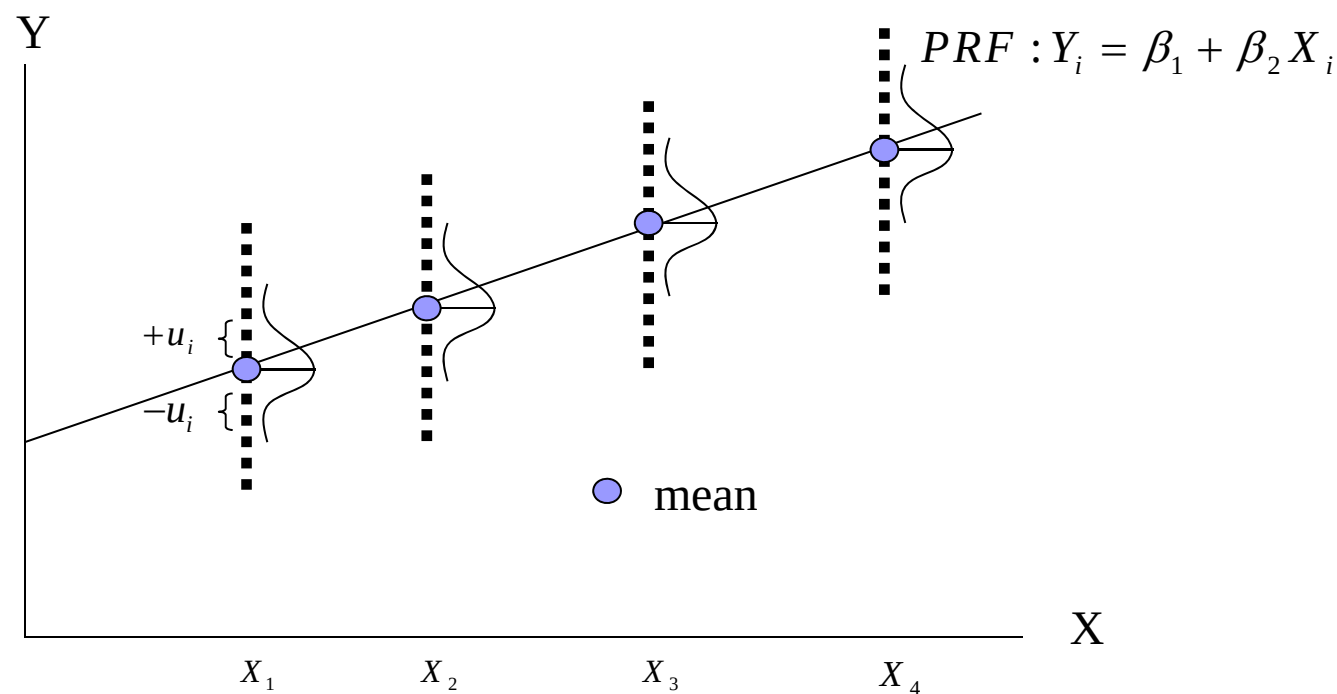
X values independent of the error term

$$Cov(X_i, u_i) = 0$$

# Zero mean value of disturbance term

$$E(u_i | X_i) = 0$$

$$E(u_i) = 0$$



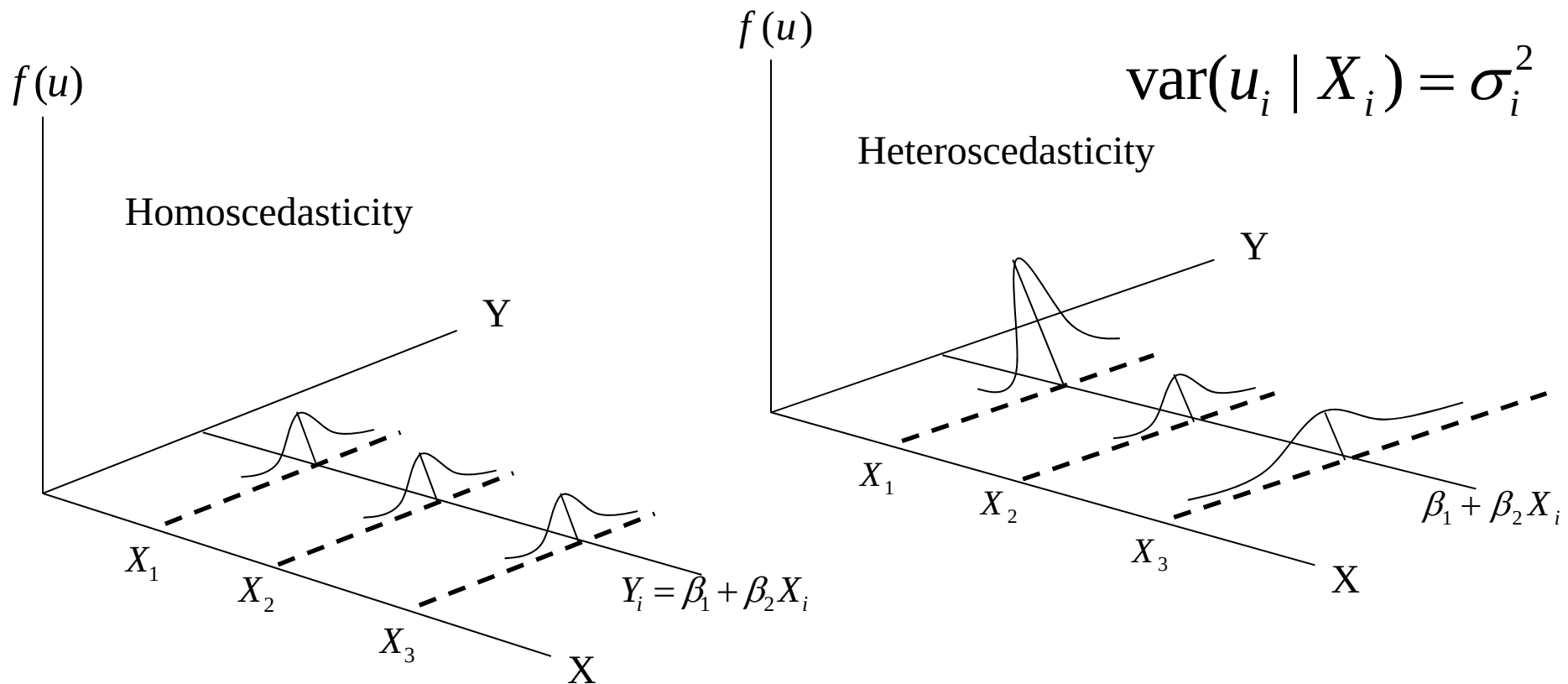


# Homoscedasticity

$$\begin{aligned}\text{var}(u_i) &= E[u_i - E(u_i | X_i)]^2 \\ &= E(u_i^2 | X_i), \text{because of assumption 3} \\ &= E(u_i^2), \text{if } X_i \text{ are nonstochastic} \\ &= \sigma^2\end{aligned}$$



# Homoscedasticity vs. Heteroscedasticity



$$\text{var}(u_i | X_i) = \sigma^2$$



# Homoscedasticity v.s. Heteroscedasticity

## Homoscedasticity

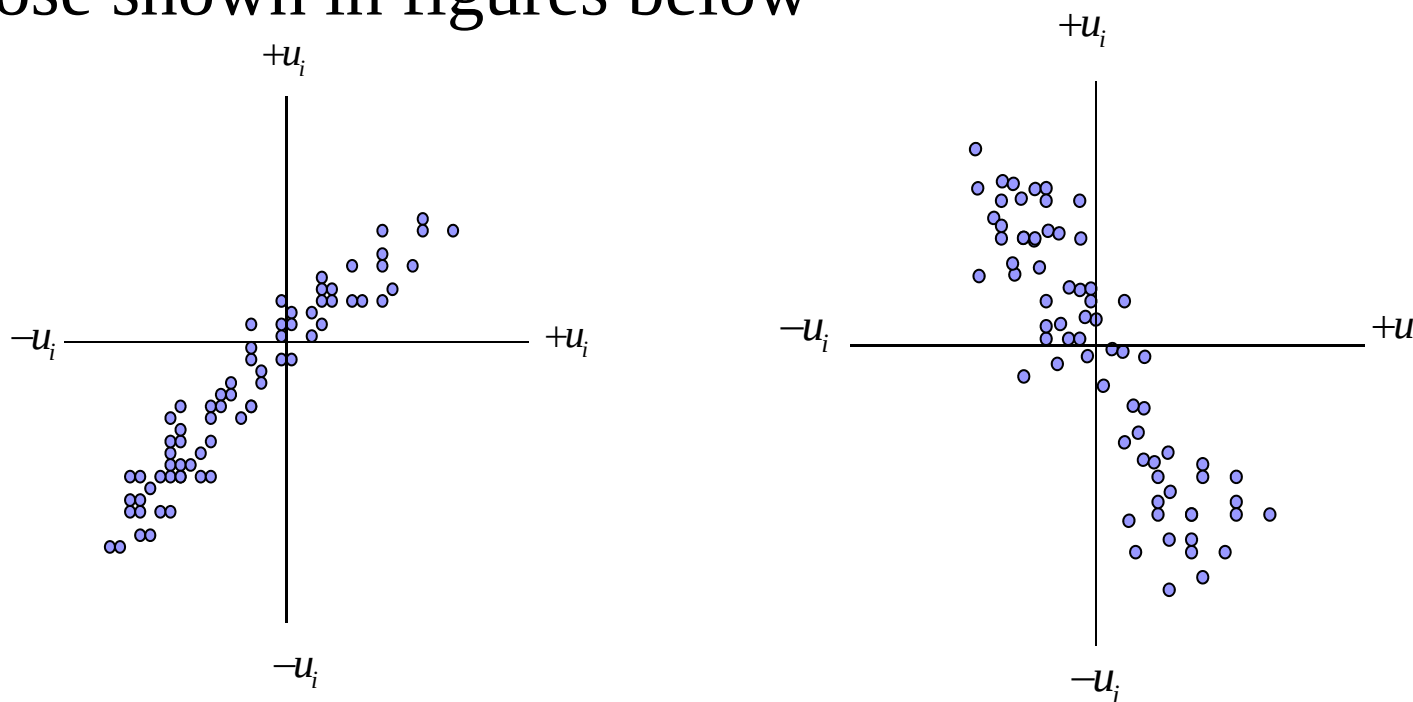
- Equal variance
- The variation around the regression line is the same across the  $X$  values

## Heteroscedasticity

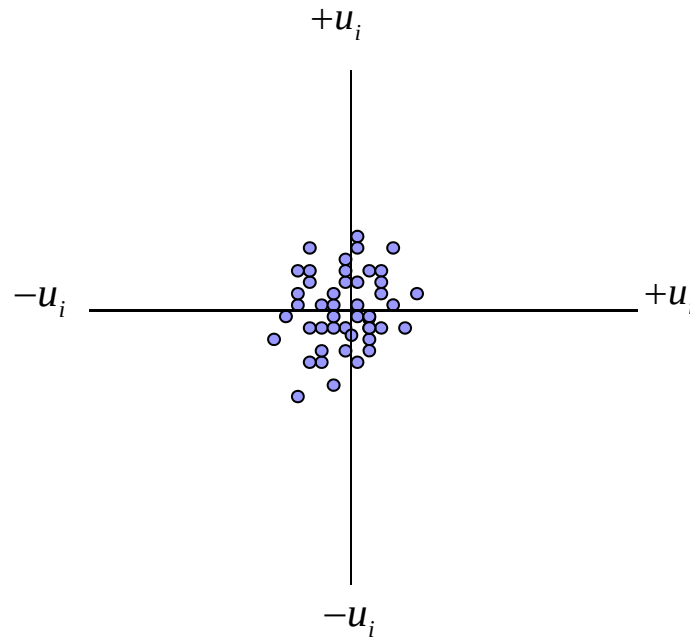
- Unequal variance

# No autocorrelation between the disturbances

Given  $X_i$ , the deviations of any two Y values from their mean value do not exhibit patterns such as those shown in figures below



# No autocorrelation between the disturbances





## No autocorrelation between the disturbances

Given any two  $X$  values,  $X_i$  and  $X_j (i \neq j)$ , the correlation between any two  $u_i$  and  $u_j (i \neq j)$  is zero.

$$\text{cov}(u_i, u_j \mid X_i, X_j) = 0$$

$$\text{cov}(u_i, u_j) = 0, \text{ if } X \text{ is nonstochastic}$$

Where  $i$  and  $j$  are two different observation and where  $\text{cov}$  means covariance



The number of observations  $n$  must be greater than the number of parameters to be estimated

- The number of observations must be greater than the number of explanatory variables



## The nature of $X$ variables

- The  $X$  values in a given sample must not all be the same
- No outliers in the  $X$  values



# Standard Errors of Least-Squares Estimates

$$\text{var}(\hat{\beta}_2) = \frac{\sigma^2}{\sum x_i^2}$$

$$\text{se}(\hat{\beta}_2) = \frac{\sigma}{\sqrt{\sum x_i^2}}$$

$$\text{var}(\hat{\beta}_1) = \frac{\sum X_i^2}{n \sum x_i^2} \sigma^2$$

$$\text{se}(\hat{\beta}_1) = \sqrt{\frac{\sum X_i^2}{n \sum x_i^2}} \sigma$$



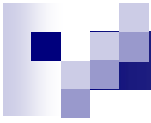


$$\hat{\sigma}^2 = \frac{\sum \hat{u}_i^2}{n - 2}$$

$\hat{\sigma}^2$  is the OLS estimator of the true but unknown  $\sigma^2$

The expression  $n-2$  is known as the number of degrees of freedom

$\sum \hat{u}_i^2$  is the sum of the residuals squared or residual sum of squares (RSS)



$$\begin{aligned} \text{COV}(\hat{\beta}_1, \hat{\beta}_2) &= -\bar{X} \text{var}(\hat{\beta}_2) \\ &= -\bar{X} \left( \frac{\sigma^2}{\sum x_i^2} \right) \\ &= \frac{-\bar{X} \sigma^2}{\sum (X_i - \bar{X})^2} \end{aligned}$$

# Example ☺

| Y  | X |
|----|---|
| 49 | 1 |
| 45 | 2 |
| 44 | 3 |
| 39 | 4 |
| 38 | 5 |
| 37 | 6 |
| 34 | 7 |
| 33 | 8 |
| 30 | 9 |

$$\hat{\beta}_1 = 49.667$$

$$\hat{\beta}_2 = -2.1576$$

$$\bar{X} = 5.5$$



| Y  | X  | $\hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$ | $\hat{u}_i$ | $\hat{u}_i^2$ | $(X_i - \bar{X})^2$ | $X_i^2$ |
|----|----|-------------------------------------------------|-------------|---------------|---------------------|---------|
| 49 | 1  | 47.5094                                         | 1.4906      | 2.2219        | 20.25               | 1       |
| 45 | 2  | 45.3518                                         | -0.352      | 0.1238        | 12.25               | 4       |
| 44 | 3  | 43.1942                                         | 0.8058      | 0.6493        | 6.25                | 9       |
| 39 | 4  | 41.0366                                         | -2.037      | 4.1477        | 2.25                | 16      |
| 38 | 5  | 38.879                                          | -0.879      | 0.7726        | 0.25                | 25      |
| 37 | 6  | 36.7214                                         | 0.2786      | 0.0776        | 0.25                | 36      |
| 34 | 7  | 34.5638                                         | -0.564      | 0.3179        | 2.25                | 49      |
| 33 | 8  | 32.4062                                         | 0.5938      | 0.3526        | 6.25                | 64      |
| 30 | 9  | 30.2486                                         | -0.249      | 0.0618        | 12.25               | 81      |
| 29 | 10 | 28.091                                          | 0.909       | 0.8263        | 20.25               | 100     |
|    |    | Total                                           |             | 9.5515        | 82.5                | 385     |



$$\hat{\sigma}^2 = \frac{\sum \hat{u}_i^2}{n-2} = \frac{9.5515}{10-2} = 1.1939$$

$$\text{var}(\hat{\beta}_2) = \frac{\hat{\sigma}^2}{\sum x_i^2} = \frac{\hat{\sigma}^2}{\sum (X_i - \bar{X})^2} = \frac{1.1939}{82.5} = 0.0145$$

$$\text{var}(\hat{\beta}_1) = \frac{\sum X_i^2}{n \sum x_i^2} \hat{\sigma}^2 = \frac{\sum X_i^2}{n \sum (X_i - \bar{X})^2} \hat{\sigma}^2 = \frac{(1.1939)(385)}{(10)(82.5)} = 0.5572$$

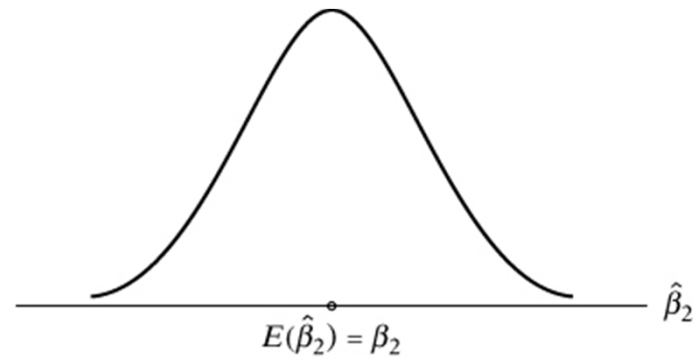
$$\text{COV}(\hat{\beta}_1, \hat{\beta}_2) = -\bar{X} \text{var}(\hat{\beta}_2) = -(5.5)(0.0145) = -0.07975$$



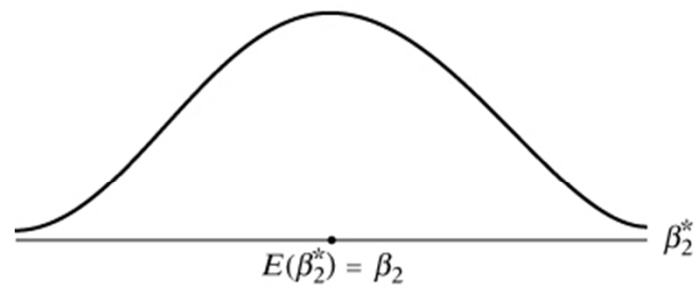
# Properties of Least-Squares Estimators: The Gauss-Markov Theorem

Best Linear Unbiased Estimator (BLUE) of  $\beta_2$  :

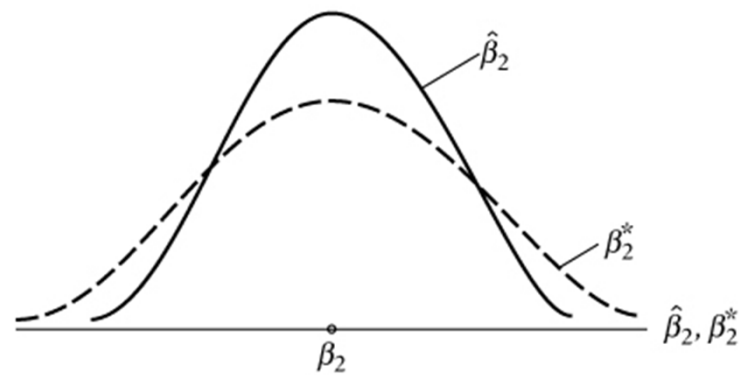
- It is linear, that is, a linear function of a random variable, such as the dependent variable  $Y$  in the regression model
- It is unbiased  $E(\hat{\beta}_2) = \beta_2$
- Efficient estimator – an unbiased estimator with the least variance



(a) Sampling distribution of  $\beta_2$



(b) Sampling distribution of  $\beta_2^*$



(c) Sampling distributions of  $\beta_2$  and  $\beta_2^*$

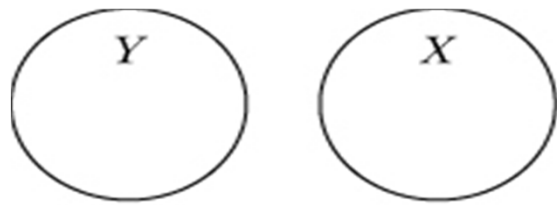
Tangtipongkul)



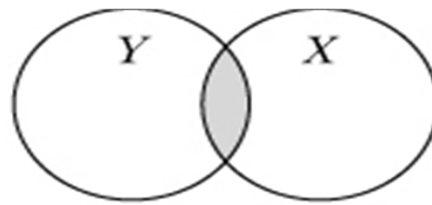
# The Coefficient of Determination $r^2$

- A measure of goodness of fit
- A summary measure that tells how well the sample regression line fits the data
- Measures the proportion or percentage of the total variation in Y explained by the regression model

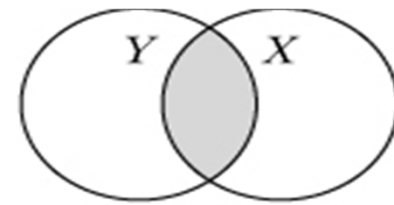




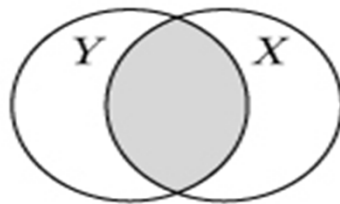
(a)



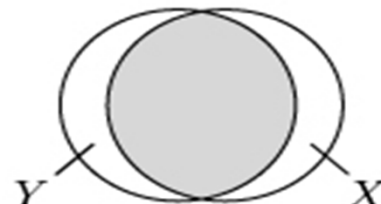
(b)



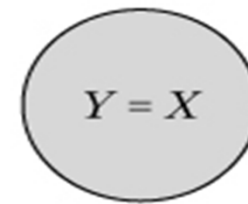
(c)



(d)



(e)



(f)



To compute this  $r^2$

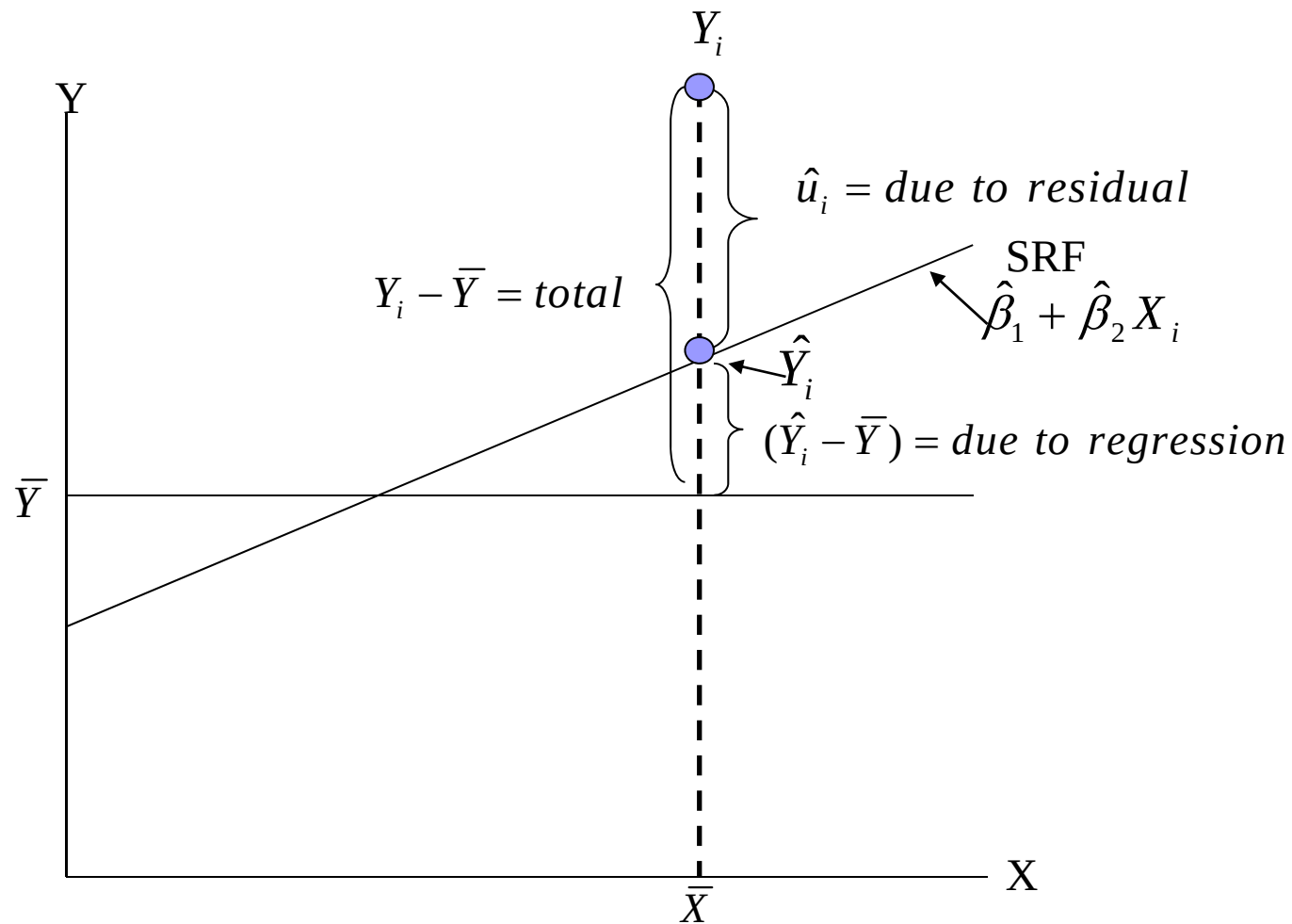
$$Y_i = \hat{Y}_i + \hat{u}_i$$

$$y_i = \hat{y}_i + \hat{u}_i$$

$$\sum y_i^2 = \sum \hat{y}_i^2 + \sum \hat{u}_i^2 + 2\sum \hat{y}_i \hat{u}_i$$

$$= \sum \hat{y}_i^2 + \sum \hat{u}_i^2$$

$$= \hat{\beta}_2^2 \sum x_i^2 + \sum \hat{u}_i^2$$





$$\sum y_i^2 = \sum (Y_i - \bar{Y})^2$$

**Total variation of the actual Y values about their sample mean**

**(Total Sum of Squares, TSS)**

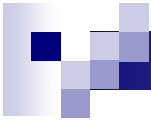
$$\sum \hat{y}_i^2 = \sum (\hat{Y}_i - \bar{\hat{Y}})^2 = \hat{\beta}_2^2 \sum x_i^2$$

**Variation of the estimated Y values about their mean  
(Explained Sum of Squares, ESS)**

$$\sum \hat{u}_i^2$$

**Residual or unexpected variation of the Y values about the regression line (Residual Sum of Squares, RSS)**

$$TSS = ESS + RSS$$



$$\begin{aligned} 1 &= \frac{ESS}{TSS} + \frac{RSS}{TSS} \\ &= \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} + \frac{\sum \hat{u}_i^2}{\sum (Y_i - \bar{Y})^2} \end{aligned}$$

$$r^2 = \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} = \frac{ESS}{TSS}$$

## Example 😊

| Y  | X |
|----|---|
| 49 | 1 |
| 45 | 2 |
| 44 | 3 |
| 39 | 4 |
| 38 | 5 |
| 37 | 6 |
| 34 | 7 |
| 33 | 8 |
| 30 | 9 |

$$\hat{\beta}_1 = 49.667$$

$$\hat{\beta}_2 = -2.1576$$

$$\bar{X} = 5.5$$

$$\bar{Y} = 38$$



| Y  | X  | $\hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$ | $\hat{u}_i$ | $(Y_i - \bar{Y})^2$ | $(\hat{Y}_i - \bar{Y})^2$ | $\hat{u}_i^2$ |
|----|----|-------------------------------------------------|-------------|---------------------|---------------------------|---------------|
| 49 | 1  |                                                 |             |                     |                           |               |
| 45 | 2  |                                                 |             |                     |                           |               |
| 44 | 3  |                                                 |             |                     |                           |               |
| 39 | 4  |                                                 |             |                     |                           |               |
| 38 | 5  |                                                 |             |                     |                           |               |
| 37 | 6  |                                                 |             |                     |                           |               |
| 34 | 7  |                                                 |             |                     |                           |               |
| 33 | 8  |                                                 |             |                     |                           |               |
| 30 | 9  |                                                 |             |                     |                           |               |
| 29 | 10 |                                                 |             |                     |                           |               |

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| Y  | X  | $\hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$ | $\hat{u}_i$ | $(Y_i - \bar{Y})^2$ | $(\hat{Y}_i - \bar{Y})^2$ | $\hat{u}_i^2$ |
|----|----|-------------------------------------------------|-------------|---------------------|---------------------------|---------------|
| 49 | 1  | 47.509                                          |             |                     |                           |               |
| 45 | 2  | 45.352                                          |             |                     |                           |               |
| 44 | 3  | 43.194                                          |             |                     |                           |               |
| 39 | 4  | 41.037                                          |             |                     |                           |               |
| 38 | 5  | 38.879                                          |             |                     |                           |               |
| 37 | 6  | 36.721                                          |             |                     |                           |               |
| 34 | 7  | 34.564                                          |             |                     |                           |               |
| 33 | 8  | 32.406                                          |             |                     |                           |               |
| 30 | 9  | 30.249                                          |             |                     |                           |               |
| 29 | 10 | 28.091                                          |             |                     |                           |               |

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$$r^2 = \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} = \frac{ESS}{TSS}$$

$$r^2 = \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} = \frac{384.06}{393.6} \approx 0.98$$

Approximately 98 percent of the variation in Y is explained by variation in X.



# Two properties of $r^2$

1. Nonnegative quantity
2. Its limits are  $0 \leq r^2 \leq 1$



# The coefficient of correlation: $r$

- A measure of the degree of association between two variables

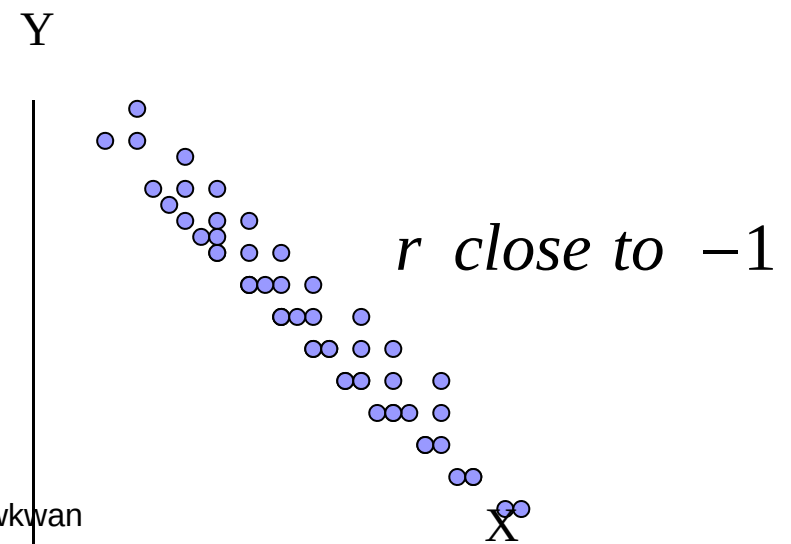
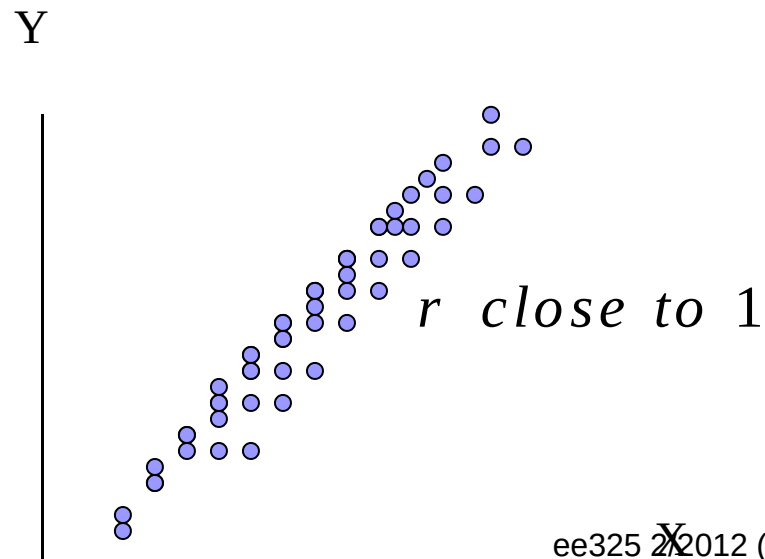
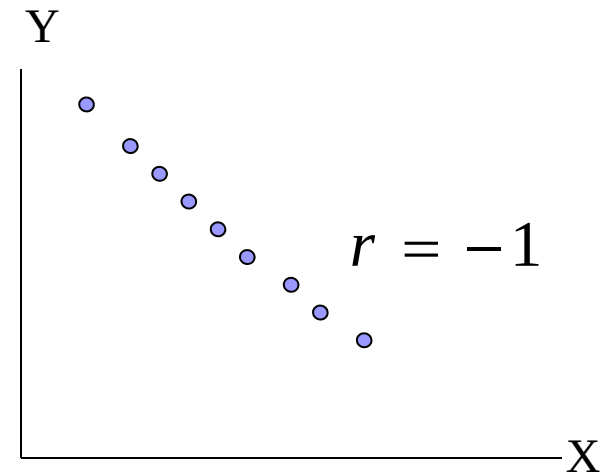
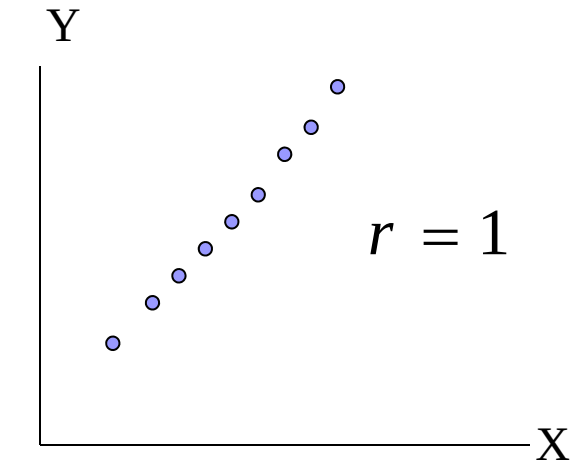
$$r = \pm \sqrt{r^2}$$

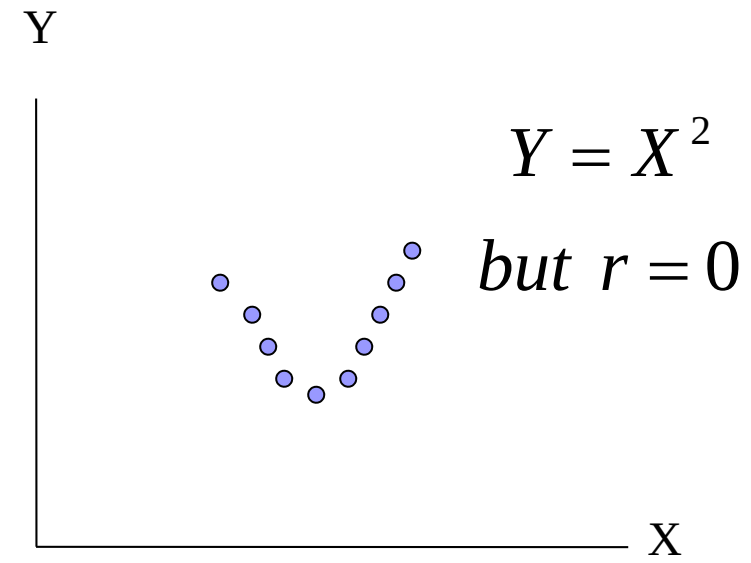
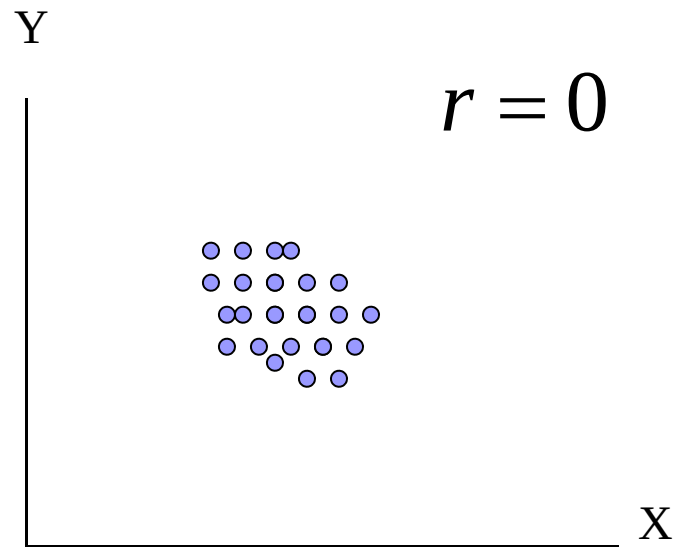
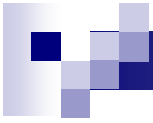


# Properties of $r$

1. Can be positive or negative
2. Lies between the limits of -1 and 1
3. Symmetric in nature
4. Independent of the origin and scale
5. If X and Y are statistically independent, the correlation coefficient between them is zero **but zero correlation does not necessarily imply independence**
6. No meaning for describing nonlinear relations
7. Does not necessarily imply any cause and effect relationship

# Correlation patterns







# Source

Gujarati, D.N. (2009) Basic Econometrics. 5th ed. Singapore, McGraw-Hill.