

EE 325

Two-Variable Regression Model: The Problem of Estimation

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

The Two-Variable PRF:

$$Y_i = \beta_1 + \beta_2 X_i + u_i$$

The Two-Variable SRF:

$$Y_i = \hat{\beta}_1 + \hat{\beta}_2 X_i + \hat{u}_i \\ = \hat{Y}_i + \hat{u}_i$$

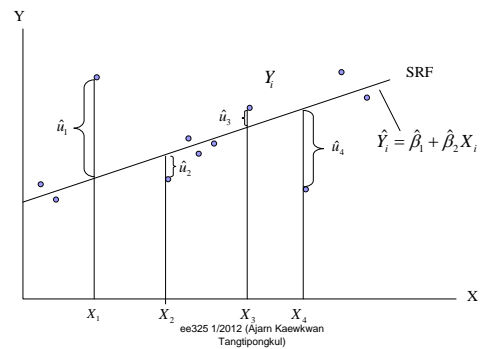
\hat{Y}_i is the estimated (conditional mean) value of Y_i

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

$$\hat{u}_i = Y_i - \hat{Y}_i \\ = Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i$$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Ordinary Least Squares (OLS)



ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Ordinary Least Squares (OLS)

$$\sum \hat{u}_i^2 = \sum (Y_i - \hat{Y}_i)^2 \\ = \sum (Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i)^2$$

The principle or the method of least squares chooses $\hat{\beta}_1$ and $\hat{\beta}_2$ in such a manner that, for a given sample or set of data, $\sum \hat{u}_i^2$ is as small as possible

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

$$\frac{\partial (\sum \hat{u}_i^2)}{\partial \hat{\beta}_1} = -2 \sum (Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i) = -2 \sum \hat{u}_i \\ \frac{\partial (\sum \hat{u}_i^2)}{\partial \hat{\beta}_2} = -2 \sum (Y_i - \hat{\beta}_1 - \hat{\beta}_2 X_i) X_i = -2 \sum \hat{u}_i X_i$$

Setting these equation to zero

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

$$\sum Y_i = n\hat{\beta}_1 + \hat{\beta}_2 \sum X_i$$

$$\sum Y_i X_i = \hat{\beta}_1 \sum X_i + \hat{\beta}_2 \sum X_i^2$$

Where n is the sample size. These simultaneous equations are known as the **Normal Equation**

ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)

$$\begin{aligned} \hat{\beta}_2 &= \frac{n \sum X_i Y_i - \sum X_i \sum Y_i}{n \sum X_i^2 - (\sum X_i)^2} \\ &= \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2} \\ &= \frac{\sum x_i y_i}{\sum x_i^2} \end{aligned}$$

ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)

$$\begin{aligned} \hat{\beta}_2 &= \frac{\sum x_i y_i}{\sum x_i^2} \\ &= \frac{\sum x_i Y_i}{\sum X_i^2 - n\bar{X}^2} \\ &= \frac{\sum X_i y_i}{\sum X_i^2 - n\bar{X}^2} \end{aligned}$$

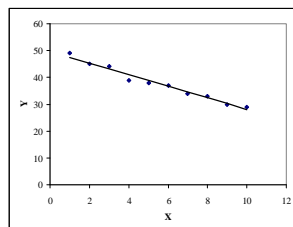
ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)

$$\begin{aligned} \hat{\beta}_1 &= \frac{\sum X_i^2 \sum Y_i - \sum X_i \sum X_i Y_i}{n \sum X_i^2 - (\sum X_i)^2} \\ &= \bar{Y} - \hat{\beta}_2 \bar{X} \end{aligned}$$

ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)

Example ☺

| Y | X |
|----|----|
| 49 | 1 |
| 45 | 2 |
| 44 | 3 |
| 39 | 4 |
| 38 | 5 |
| 37 | 6 |
| 34 | 7 |
| 33 | 8 |
| 30 | 9 |
| 29 | 10 |



ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)

$$\hat{\beta}_2 = \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2} = \frac{-178}{82.5} \approx -2.1576$$

$$\hat{\beta}_1 = \bar{Y} - \hat{\beta}_2 \bar{X} = 37.8 - (-2.1576)(5.5) = 49.667$$

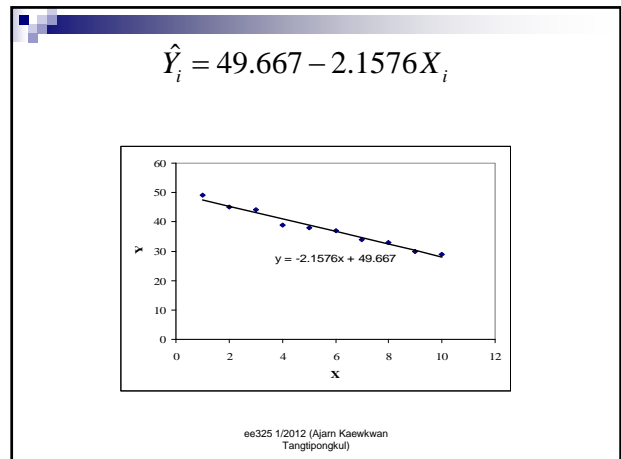
$$\bar{X} = 5.5$$

$$\bar{Y} = 37.8$$

ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)

| Y | X | $(X_i - \bar{X})$ | $(X_i - \bar{X})^2$ | $Y_i - \bar{Y}$ | $(X_i - \bar{X})(Y_i - \bar{Y})$ |
|----|----|-------------------|---------------------|-----------------|----------------------------------|
| 49 | 1 | -4.5 | 20.25 | 11.2 | -50.4 |
| 45 | 2 | -3.5 | 12.25 | 7.2 | -25.2 |
| 44 | 3 | -2.5 | 6.25 | 6.2 | -15.5 |
| 39 | 4 | -1.5 | 2.25 | 1.2 | -1.8 |
| 38 | 5 | -0.5 | 0.25 | 0.2 | -0.1 |
| 37 | 6 | 0.5 | 0.25 | -0.8 | -0.4 |
| 34 | 7 | 1.5 | 2.25 | -3.8 | -5.7 |
| 33 | 8 | 2.5 | 6.25 | -4.8 | -12 |
| 30 | 9 | 3.5 | 12.25 | -7.8 | -27.3 |
| 29 | 10 | 4.5 | 20.25 | -8.8 | -39.6 |
| | | 82.5 | | -178 | |

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)



Practice I

A Random Sample from the Population of Table 2.1

| Weekly consumption expenditure S (Y) | Weekly income \$(X) |
|--------------------------------------|---------------------|
| 70 | 80 |
| 65 | 100 |
| 90 | 120 |
| 95 | 140 |
| 110 | 160 |
| 115 | 180 |
| 120 | 200 |
| 140 | 220 |
| 155 | 240 |
| 150 | 260 |

$\hat{Y} = 0.5091X_i + 24.455$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

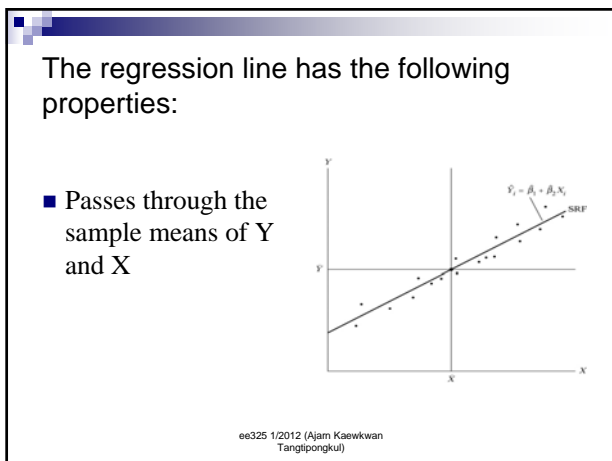
Practice II

Another Random Sample from the Population of Table 2.1

| Weekly consumption expenditure S (Y) | Weekly income \$(X) |
|--------------------------------------|---------------------|
| 55 | 80 |
| 88 | 100 |
| 90 | 120 |
| 80 | 140 |
| 118 | 160 |
| 120 | 180 |
| 145 | 200 |
| 135 | 220 |
| 145 | 240 |
| 175 | 260 |

$\hat{Y} = 0.5761X_i + 17.17$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)



- The mean value of the estimated $Y = \hat{Y}_i$ is equal to the mean value of the actual Y for

$$\hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

- The mean value of the residuals \hat{u}_i is zero
- The residuals \hat{u}_i are uncorrelated with the predicted \hat{Y}
- The residuals \hat{u}_i are uncorrelated with X_i

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Classical Linear regression model (CLRM)

The assumptions underlying the method of least squares:

1. Linear regression model
2. Fixed X values or X values independent of the error term
3. Zero mean value of disturbance
4. Homoscedasticity or Constant Variance of u_i
5. No autocorrelation between the disturbances u_i
6. The number of observations n must be greater than the number of parameters to be estimated
7. The nature of X variables

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Linear regression model

$$Y_i = \beta_1 + \beta_2 X_i + u_i$$

- Linear in the parameters
- May or may not be linear in the variables

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

X values independent of the error term

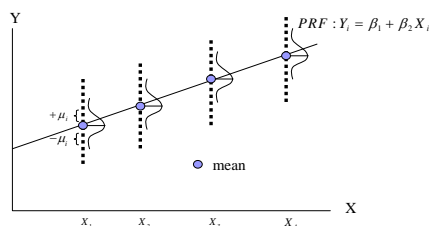
$$Cov(X_i, u_i) = 0$$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Zero mean value of disturbance term

$$E(u_i | X_i) = 0$$

$$E(u_i) = 0$$



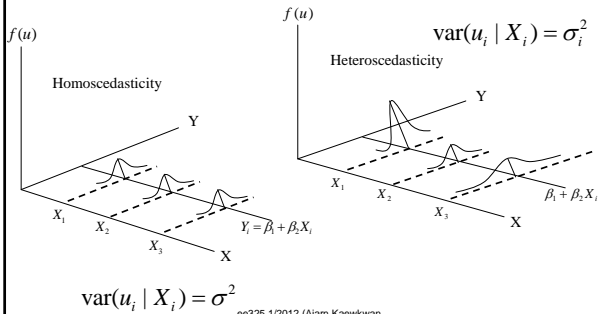
ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Homoscedasticity

$$\begin{aligned} \text{var}(u_i) &= E[u_i - E(u_i | X_i)]^2 \\ &= E(u_i^2 | X_i), \text{ because of assumption 3} \\ &= E(u_i^2), \text{ if } X_i \text{ are nonstochastic} \\ &= \sigma^2 \end{aligned}$$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Homoscedasticity vs. Heteroscedasticity



Homoscedasticity v.s. Heteroscedasticity

Homoscedasticity

- Equal variance
- The variation around the regression line is the same across the X values

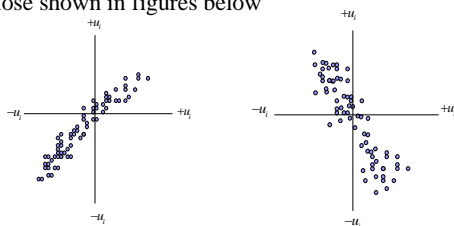
Heteroscedasticity

- Unequal variance

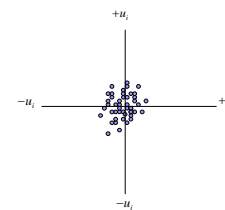
ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

No autocorrelation between the disturbances

Given X_i , the deviations of any two Y values from their mean value do not exhibit patterns such as those shown in figures below



No autocorrelation between the disturbances



No autocorrelation between the disturbances

Given any two X values, X_i and $X_j (i \neq j)$, the correlation between any two u_i and $u_j (i \neq j)$ is zero.

$$\text{cov}(u_i, u_j | X_i, X_j) = 0$$

$$\text{cov}(u_i, u_j) = 0, \text{ if } X \text{ is nonstochastic}$$

Where i and j are two different observation and where cov means covariance

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

The number of observations n must be greater than the number of parameters to be estimated

- The number of observations must be greater than the number of explanatory variables

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

The nature of X variables

- The X values in a given sample must not all be the same
- No outliers in the X values

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Standard Errors of Least-Squares Estimates

$$\begin{aligned} \text{var}(\hat{\beta}_2) &= \frac{\sigma^2}{\sum x_i^2} \\ \text{se}(\hat{\beta}_2) &= \frac{\sigma}{\sqrt{\sum x_i^2}} \\ \text{var}(\hat{\beta}_1) &= \frac{\sum X_i^2}{n \sum x_i^2} \sigma^2 \\ \text{se}(\hat{\beta}_1) &= \sqrt{\frac{\sum X_i^2}{n \sum x_i^2}} \sigma \end{aligned}$$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

$$\hat{\sigma}^2 = \frac{\sum \hat{u}_i^2}{n - 2}$$

$\hat{\sigma}^2$ is the OLS estimator of the true but unknown σ^2

The expression n-2 is known as the number of degrees of freedom

$\sum \hat{u}_i^2$ is the sum of the residuals squared or residual sum of squares (RSS)

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

$$\begin{aligned} \text{COV}(\hat{\beta}_1, \hat{\beta}_2) &= -\bar{X} \text{var}(\hat{\beta}_2) \\ &= -\bar{X} \left(\frac{\sigma^2}{\sum x_i^2} \right) \\ &= \frac{-\bar{X} \sigma^2}{\sum (X_i - \bar{X})^2} \end{aligned}$$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

Example ☺

| Y | X |
|----|----|
| 49 | 1 |
| 45 | 2 |
| 44 | 3 |
| 39 | 4 |
| 38 | 5 |
| 37 | 6 |
| 34 | 7 |
| 33 | 8 |
| 30 | 9 |
| 29 | 10 |

$$\begin{aligned} \hat{\beta}_1 &= 49.667 \\ \hat{\beta}_2 &= -2.1576 \\ \bar{X} &= 5.5 \end{aligned}$$

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

| Y | X | $\hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$ | \hat{u}_i | \hat{u}_i^2 | $(X_i - \bar{X})^2$ | X_i^2 |
|-------|----|---|-------------|---------------|---------------------|---------|
| 49 | 1 | 47.5094 | 1.4906 | 2.2219 | 20.25 | 1 |
| 45 | 2 | 45.3518 | -0.352 | 0.1238 | 12.25 | 4 |
| 44 | 3 | 43.1942 | 0.8058 | 0.6493 | 6.25 | 9 |
| 39 | 4 | 41.0366 | -2.037 | 4.1477 | 2.25 | 16 |
| 38 | 5 | 38.879 | -0.879 | 0.7726 | 0.25 | 25 |
| 37 | 6 | 36.7214 | 0.2786 | 0.0776 | 0.25 | 36 |
| 34 | 7 | 34.5638 | -0.564 | 0.3179 | 2.25 | 49 |
| 33 | 8 | 32.4062 | 0.5938 | 0.3526 | 6.25 | 64 |
| 30 | 9 | 30.2486 | -0.249 | 0.0618 | 12.25 | 81 |
| 29 | 10 | 28.091 | 0.909 | 0.8263 | 20.25 | 100 |
| Total | | | 9.5515 | 82.5 | 385 | |

ee325 1/2012 (Ajarn Kaewkwan Tangtipongkul)

$$\hat{\sigma}^2 = \frac{\sum \hat{u}_i^2}{n-2} = \frac{9.5515}{10-2} = 1.1939$$

$$\text{var}(\hat{\beta}_2) = \frac{\hat{\sigma}^2}{\sum x_i^2} = \frac{\hat{\sigma}^2}{\sum (X_i - \bar{X})^2} = \frac{1.1939}{82.5} = 0.0145$$

$$\text{var}(\hat{\beta}_1) = \frac{\sum X_i^2}{n \sum x_i^2} \hat{\sigma}^2 = \frac{\sum X_i^2}{n \sum (X_i - \bar{X})^2} \hat{\sigma}^2 = \frac{(1.1939)(385)}{(10)(82.5)} = 0.5572$$

$$\text{COV}(\hat{\beta}_1, \hat{\beta}_2) = -\bar{X} \text{var}(\hat{\beta}_2) = -(5.5)(0.0145) = -0.07975$$

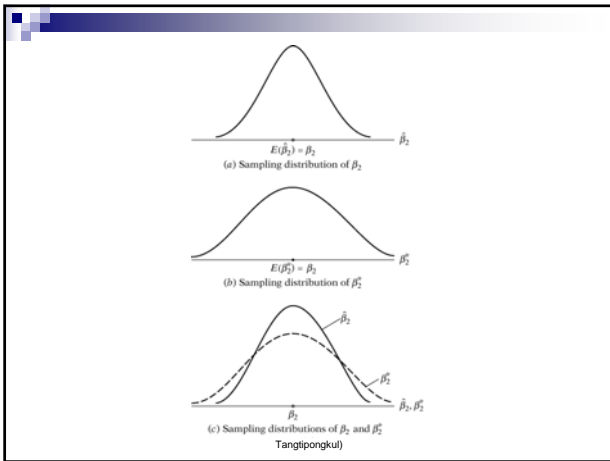
ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)

Properties of Least-Squares Estimators: The Gauss-Markov Theorem

Best Linear Unbiased Estimator (BLUE) of β_2 :

- It is linear, that is, a linear function of a random variable, such as the dependent variable Y in the regression model
- It is unbiased $E(\hat{\beta}_2) = \beta_2$
- Efficient estimator – an unbiased estimator with the least variance

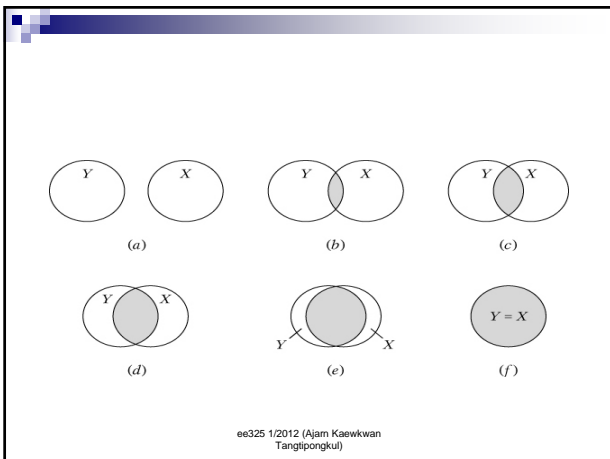
ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)



The Coefficient of Determination r^2

- A measure of goodness of fit
- A summary measure that tells how well the sample regression line fits the data
- Measures the proportion or percentage of the total variation in Y explained by the regression model

ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)



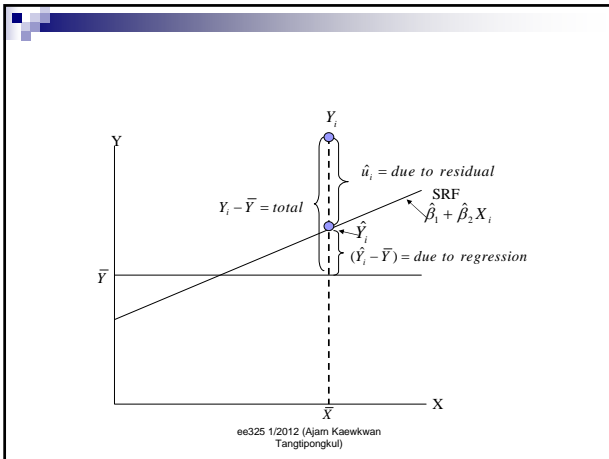
To compute this r^2

$$Y_i = \hat{Y}_i + \hat{u}_i$$

$$y_i = \hat{y}_i + \hat{u}_i$$

$$\begin{aligned} \sum y_i^2 &= \sum \hat{y}_i^2 + \sum \hat{u}_i^2 + 2 \sum \hat{y}_i \hat{u}_i \\ &= \sum \hat{y}_i^2 + \sum \hat{u}_i^2 \\ &= \hat{\beta}_2^2 \sum x_i^2 + \sum \hat{u}_i^2 \end{aligned}$$

ee325 1/2012 (Ajarn Kaewkwan Tanglipongkul)



$$\sum y_i^2 = \sum (Y_i - \bar{Y})^2 \quad \text{Total variation of the actual Y values about their sample mean (Total Sum of Squares, TSS)}$$

$$\sum \hat{y}_i^2 = \sum (\hat{Y}_i - \bar{Y})^2 = \hat{\beta}_2^2 \sum x_i^2 \quad \text{Variation of the estimated Y values about their mean (Explained Sum of Squares, ESS)}$$

$$\sum \hat{u}_i^2 \quad \text{Residual or unexpected variation of the Y values about the regression line (Residual Sum of Squares, RSS)}$$

$$TSS = ESS + RSS$$

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

$$1 = \frac{ESS}{TSS} + \frac{RSS}{TSS}$$

$$= \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} + \frac{\sum \hat{u}_i^2}{\sum (Y_i - \bar{Y})^2}$$

$$r^2 = \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} = \frac{ESS}{TSS}$$

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

Example ☺

| Y | X |
|----|----|
| 49 | 1 |
| 45 | 2 |
| 44 | 3 |
| 39 | 4 |
| 38 | 5 |
| 37 | 6 |
| 34 | 7 |
| 33 | 8 |
| 30 | 9 |
| 29 | 10 |

$$\hat{\beta}_1 = 49.667$$

$$\hat{\beta}_2 = -2.1576$$

$$\bar{X} = 5.5$$

$$\bar{Y} = 38$$

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

| Y | X | $\hat{Y}_i = \hat{\beta}_1 + \hat{\beta}_2 X_i$ | \hat{u}_i | $(Y_i - \bar{Y})^2$ | $(\hat{Y}_i - \bar{Y})^2$ | \hat{u}_i^2 |
|----|----|---|-------------|---------------------|---------------------------|---------------|
| 49 | 1 | 47.509 | 1.491 | 125.44 | 94.27 | 2.22 |
| 45 | 2 | 45.352 | -0.352 | 51.84 | 57.03 | 0.12 |
| 44 | 3 | 43.194 | 0.806 | 38.44 | 29.10 | 0.65 |
| 39 | 4 | 41.037 | -2.037 | 1.44 | 10.48 | 4.15 |
| 38 | 5 | 38.879 | -0.879 | 0.04 | 1.16 | 0.77 |
| 37 | 6 | 36.721 | 0.279 | 0.64 | 1.16 | 0.08 |
| 34 | 7 | 34.564 | -0.564 | 14.44 | 10.47 | 0.32 |
| 33 | 8 | 32.406 | 0.594 | 23.04 | 29.09 | 0.35 |
| 30 | 9 | 30.249 | -0.249 | 60.84 | 57.02 | 0.06 |
| 29 | 10 | 28.091 | 0.909 | 77.44 | 94.26 | 0.83 |
| | | | | 393.6 | 384.06 | 9.55 |

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

$$r^2 = \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} = \frac{ESS}{TSS}$$

$$r^2 = \frac{\sum (\hat{Y}_i - \bar{Y})^2}{\sum (Y_i - \bar{Y})^2} = \frac{384.06}{393.6} \approx 0.98$$

Approximately 98 percent of the variation in Y is explained by variation in X.

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

Two properties of r^2

1. Nonnegative quantity
2. Its limits are $0 \leq r^2 \leq 1$

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

The coefficient of correlation: r

- A measure of the degree of association between two variables

$$r = \pm \sqrt{r^2}$$

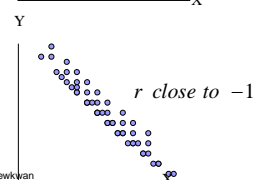
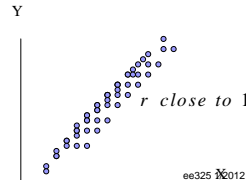
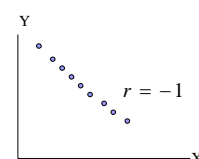
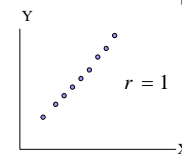
ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

Properties of r

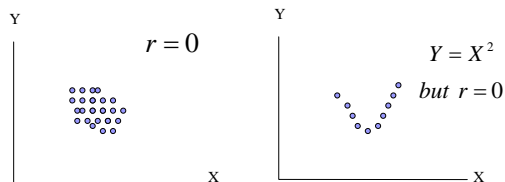
1. Can be positive or negative
2. Lies between the limits of -1 and 1
3. Symmetric in nature
4. Independent of the origin and scale
5. If X and Y are statistically independent, the correlation coefficient between them is zero **but zero correlation does not necessarily imply independence**
6. No meaning for describing nonlinear relations
7. Does not necessarily imply any cause and effect relationship

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

Correlation patterns



ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)



ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)

Source

Gujarati, D.N. (2009) Basic Econometrics. 5th ed. Singapore, McGraw-Hill.

ee325 1/2012 (Ajarn Kaewkwan
Tangtipongkul)