

Limited Dependent Variable Models

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Part 2

Chayanee Chawanote

Binary response models

- ▶ The response probability: $P(y = 1|\mathbf{x}) = P(y = 1|x_1, x_2, \dots, x_k)$
- ▶ In the LPM, we assume that the response probability is linear in a set of parameters.
- ▶ For general setting, define the response probability as a function taking on values between 0 and 1: $0 < G(z) < 1$
 $P(y = 1|\mathbf{x}) = G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k) = G(\beta_0 + \mathbf{x}\beta)$ (1)
- ▶ The two most used function G that ensure the probabilities between 0 and 1:
 - ▶ the logistic function $\rightarrow$ logit model
 - ▶ the standard normal cumulative distribution function $\rightarrow$ probit model

Binary response models

Specifying Logit and Probit models

- ▶ Logit model: G is the logistic function

$$G(z) = \frac{\exp(z)}{1+\exp(z)} = \Lambda(z) \quad (2)$$

- ▶ Probit model: G is the standard normal cumulative distribution function (cdf)

$$G(z) = \Phi(z) \equiv \int_{-\infty}^z \phi(v) dv, \quad (3)$$

where $\phi(z)$ is the standard normal density:

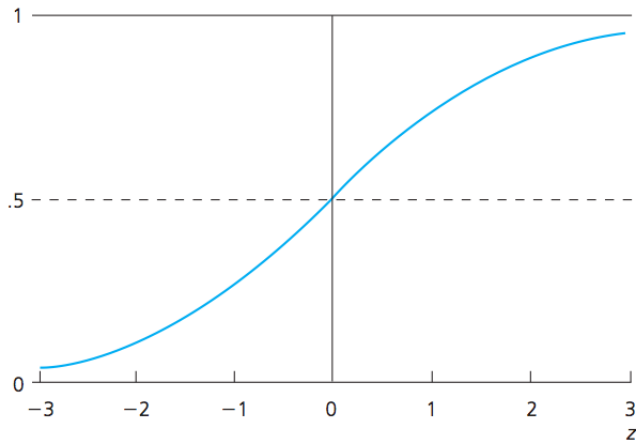
$$\phi(z) = \sqrt{2\pi} \exp\left(-\frac{z^2}{2}\right) \quad (4)$$

- ▶ $G(z)$ in both (2) and (3) are increasing functions with $G(z) \rightarrow 0$ as $z \rightarrow -\infty$, and $G(z) \rightarrow 1$ as $z \rightarrow \infty$.

Logistic function

Graph of the logistic function $G(z) = \exp(z)/[1 + \exp(z)]$.

$$G(z) = \exp(z)/[1 + \exp(z)]$$



Latent variable model

- ▶ Logit and probit models can be derived from an underlying latent variable model.
- ▶ Let y^* be an unobserved, or latent, variable determined by $y^* = \beta_0 + \mathbf{x}\beta + e$, $y = 1[y^* > 0]$ (5)
- ▶ $1[.]$ is an indicator function, which is $= 1$ if the event in the brackets is true, and $= 0$ otherwise.
- ▶ Assume e is independent of $\mathbf{x}$ and e either has the standard logistic distribution or the standard normal distribution.
 - ▶ e is symmetrically distributed about zero
 - ▶ $1 - G(-z) = G(z)$ for all real numbers z
- ▶ Probit is more popular since economists tend to favor the normality assumption for e

Latent variable model

- ▶ From (5), we can derive the response probability for y :
$$P(y = 1|\mathbf{x}) = P(y^* > 0|\mathbf{x}) = P[e > -(\beta_0 + \mathbf{x}\beta)|\mathbf{x}]$$
$$= 1 - G[-(\beta_0 + \mathbf{x}\beta)] = G(\beta_0 + \mathbf{x}\beta) \quad (6)$$
- ▶ The direction of the effect of x_j on $E(y^*|\mathbf{x}) = \beta_0 + \mathbf{x}\beta$ and on $E(y|\mathbf{x}) = P(y = 1|\mathbf{x}) = G(\beta_0 + \mathbf{x}\beta)$ is always the same.
- ▶ For the magnitude, we care about the effect of x_j on the probability of success $P(y = 1|\mathbf{x})$.

Maximum likelihood estimation of logit and probit models

- ▶ Suppose we have a random sample of size n .
- ▶ To obtain MLE, conditional on the explanatory variables, we need the density of y given $\mathbf{x}_j$:
$$f(y|\mathbf{x}_j; \beta) = [G(\mathbf{x}_j\beta)]^y [1 - G(\mathbf{x}_j\beta)]^{1-y}, y = 0, 1 \quad (7)$$
- ▶ Then, the log-likelihood function for observation i is:
$$\ell_i(\beta) = y_i \log[G(\mathbf{x}_i\beta)] + (1 - y_i) \log[1 - G(\mathbf{x}_i\beta)] \quad (8)$$
- ▶ The log-likelihood for a sample size of n : $\ell(\beta) = \sum_{i=1}^n \ell_i(\beta)$ (9)
- ▶ Maximum likelihood estimator (MLE): $\hat{\beta}$ maximizes the log-likelihood in (9)
- ▶ Under general conditions, the MLE is consistent, asymptotically normal, and asymptotically efficient.

Goodness of fit

- ▶ For binary response model, we calculate pseudo R-squared = $1 - \ell_{ur}/\ell_0$
- ▶ ℓ_{ur} = log-likelihood function for the estimated model, ℓ_0 = log-likelihood function with only intercept model.
- ▶ The log-likelihood function is always a negative number, and $|\ell_{ur}| \leq |\ell_0|$.
- ▶ If the covariates have no explanatory power, then $\ell_{ur}/\ell_0 = 1$ and pseudo $R^2 = 0$
- ▶ If $\ell_{ur} = 0$, the pseudo $R^2 = 1$

Interpretation of probits and logits

- ▶ We care about the effect of $\mathbf{x}$ on $P(y=1|\mathbf{x})$, that is $\partial p / \partial \mathbf{x}$.
- ▶ Probit and logit are non-linear, partial effect on $p(\mathbf{x})$ is from $\frac{\partial p(\mathbf{x})}{\partial x_j} = g(\beta_0 + \mathbf{x}\beta)\beta_j$, where $g(z) \equiv \frac{dG}{dz}(z)$
where G is the cdf, g is a probability density function
- ▶ $G(\cdot)$ is a strictly increasing cdf, so $g(z) > 0$ for all z
 - ▶ The partial effect always has the same sign as β_j
- ▶ So, we can only directly compare sign and significance (based on a standard t test) of coefficients

Hypotheses testing

- ▶ For each coefficient, we can construct (asymptotic) t tests as usual: $t = \hat{\beta}_j / se(\hat{\beta}_j)$
- ▶ For multiple hypotheses, or tests of multiple exclusion restrictions, we use the likelihood ratio (LR) test
- ▶ LR test is based on the difference in the log-likelihood functions for the unrestricted and restricted models, whether the log-likelihood becomes significantly smaller (no larger) when dropping variables.
- ▶ LR statistic: $LR = 2(\ell_{ur} - \ell_r)$
- ▶ Under H_0 , if we are testing q exclusion restrictions, $LR \sim \chi^2_q$