

# Multiple Regression Analysis with Qualitative Information:

## 1 Outline

- Describing qualitative information
- Using a single dummy independent variable
- Using dummy variables for multiple categories
- Interactions involving dummy variables
- A binary dependent variable (Y variable): The linear probability model

## 2 Describing Qualitative Information

- "Female" and "Married" are qualitative variable.
- We arbitrarily assign a dummy variable to describe them.

$$\begin{aligned}
 female &= \begin{cases} 1 & \text{if female} \\ 0 & \text{otherwise (or if male)} \end{cases} \\
 married &= \begin{cases} 1 & \text{if married} \\ 0 & \text{otherwise (of if single)} \end{cases}
 \end{aligned}$$

TABLE 7.1

A Partial Listing of the Data in WAGE1.RAW

| <i>person</i> | <i>wage</i> | <i>educ</i> | <i>exper</i> | <i>female</i> | <i>married</i> |
|---------------|-------------|-------------|--------------|---------------|----------------|
| 1             | 3.10        | 11          | 2            | 1             | 0              |
| 2             | 3.24        | 12          | 22           | 1             | 1              |
| 3             | 3.00        | 11          | 2            | 0             | 0              |
| 4             | 6.00        | 8           | 44           | 0             | 1              |
| 5             | 5.30        | 12          | 7            | 0             | 1              |
| ⋮             | ⋮           | ⋮           | ⋮            | ⋮             | ⋮              |
| 525           | 11.56       | 16          | 5            | 0             | 1              |
| 526           | 3.50        | 14          | 5            | 1             | 0              |

## 3 Models with a single dummy independent variable

Consider

$$wage = \beta_0 + \delta_0 female + \beta_1 educ + u.$$

where

$$female = \begin{cases} 1 & \text{if female} \\ 0 & \text{otherwise (or if male)} \end{cases}$$

In this case, the  $\delta_0$  notation is used to highlight the interpretation of the parameters multiplying dummy variables. In other cases, we can use any notation that is the most convenient.

4 It is not possible to include all of the dummy alternatives in the same model

- If we include all alternatives of a dummy variable in the same model, we will face the "perfect collinearity" problem.

For example:

$$\begin{aligned} 1 &= female + male \\ female &= male + 1 \end{aligned}$$

or

$$\begin{aligned} 1 &= winter + spring + summer + fall \\ winter &= 1 - spring - summer - fall \end{aligned}$$

- At least one alternative has to be dropped. We treat the dropped alternative as the "BASE GROUP" or "BASELINE" or "BENCHMARK GROUP".

| <pre>. regress lwage female male married educ exper</pre> |            |           |            |                        |                      |          |
|-----------------------------------------------------------|------------|-----------|------------|------------------------|----------------------|----------|
| note: male omitted because of collinearity                |            |           |            |                        |                      |          |
| Source                                                    | SS         | df        | MS         | Number of obs = 526    |                      |          |
| Model                                                     | 54.3265253 | 4         | 13.5816313 | F( 4, 521) = 75.27     |                      |          |
| Residual                                                  | 94.0032262 | 521       | .180428457 | Prob > F = 0.0000      |                      |          |
| Total                                                     | 148.329751 | 525       | .28253286  | R-squared = 0.3663     |                      |          |
|                                                           |            |           |            | Adj R-squared = 0.3614 |                      |          |
|                                                           |            |           |            | Root MSE = .42477      |                      |          |
| lwage                                                     | Coef.      | Std. Err. | t          | P> t                   | [95% Conf. Interval] |          |
| female                                                    | -.3251146  | .0377061  | -8.62      | 0.000                  | -.3991892            | -.25104  |
| male                                                      | 0          | (omitted) |            |                        |                      |          |
| married                                                   | .1380145   | .0411197  | 3.36       | 0.001                  | .0572338             | .2187953 |
| educ                                                      | .0872644   | .0071554  | 12.20      | 0.000                  | .0732075             | .1013213 |
| exper                                                     | .0076213   | .0015314  | 4.98       | 0.000                  | .0046129             | .0106297 |
| _cons                                                     | .4690918   | .1040575  | 4.51       | 0.000                  | .264668              | .6735156 |

## 5 Using dummy variables for multiple categories

**Case 1** We can use many dummy variables in the same model

Consider a model which includes 2 dummy variables– *female* and *married*.

$$\log(wage) = \beta_0 + \delta_0 female + \delta_1 married + \beta_1 educ + \beta_2 exper + \beta_3 exper^2 + \beta_4 tenure + \beta_5 tenure^2 + u.$$

```
regress lwage female married educ exper expersq tenure tenursq
```

| Source   | SS         | df  | MS         | Number of obs = 526    |  |  |
|----------|------------|-----|------------|------------------------|--|--|
| Model    | 65.6482326 | 7   | 9.37831895 | F( 7, 518) = 58.76     |  |  |
| Residual | 82.6815188 | 518 | .159616832 | Prob > F = 0.0000      |  |  |
| Total    | 148.329751 | 525 | .28253286  | R-squared = 0.4426     |  |  |
|          |            |     |            | Adj R-squared = 0.4351 |  |  |
|          |            |     |            | Root MSE = .39952      |  |  |

  

| lwage   | Coef.     | Std. Err. | t     | P> t  | [95% Conf. Interval] |           |
|---------|-----------|-----------|-------|-------|----------------------|-----------|
| female  | -.2901838 | .0361121  | -8.04 | 0.000 | -.3611279            | -.2192396 |
| married | .0529219  | .0407561  | 1.30  | 0.195 | -.0271456            | .1329894  |
| educ    | .0791547  | .0068003  | 11.64 | 0.000 | .0657952             | .0925143  |
| exper   | .0269535  | .0053258  | 5.06  | 0.000 | .0164907             | .0374163  |
| expersq | -.0005399 | .0001122  | -4.81 | 0.000 | -.0007603            | -.0003196 |
| tenure  | .0312962  | .0068482  | 4.57  | 0.000 | .0178426             | .0447499  |
| tenursq | -.0005744 | .0002347  | -2.45 | 0.015 | -.0010355            | -.0001134 |
| _cons   | .4177837  | .0988662  | 4.23  | 0.000 | .2235557             | .6120116  |

Comments:

Consider a model which includes dummy variables for each gender/marital status combination— *marrmale*, *marrfem* and *singfem*.

$$\begin{aligned} \log(wage) = & \beta_0 + \delta_0 marrmale + \delta_1 marrfem + \delta_3 singfem + \beta_1 educ + \beta_2 exper \\ & + \beta_3 exper^2 + \beta_4 tenure + \beta_5 tenure^2 + u. \end{aligned} \quad (8.1)$$

```
regress lwage marrmale marrfem singfem educ exper expersq tenure tenursq
```

| Source   | SS         | df  | MS         | Number of obs = 526    |  |  |
|----------|------------|-----|------------|------------------------|--|--|
| Model    | 68.3617623 | 8   | 8.54522029 | F( 8, 517) = 55.25     |  |  |
| Residual | 79.9679891 | 517 | .154676961 | Prob > F = 0.0000      |  |  |
| Total    | 148.329751 | 525 | .28253286  | R-squared = 0.4609     |  |  |
|          |            |     |            | Adj R-squared = 0.4525 |  |  |
|          |            |     |            | Root MSE = .39329      |  |  |

  

| lwage    | Coef.     | Std. Err. | t     | P> t  | [95% Conf. Interval] |           |
|----------|-----------|-----------|-------|-------|----------------------|-----------|
| marrmale | .2126757  | .0553572  | 3.84  | 0.000 | .103923              | .3214284  |
| marrfem  | -.1982676 | .0578355  | -3.43 | 0.001 | -.311889             | -.0846462 |
| singfem  | -.1103502 | .0557421  | -1.98 | 0.048 | -.219859             | -.0008414 |
| educ     | .0789103  | .0066945  | 11.79 | 0.000 | .0657585             | .092062   |
| exper    | .0268006  | .0052428  | 5.11  | 0.000 | .0165007             | .0371005  |
| expersq  | -.0005352 | .0001104  | -4.85 | 0.000 | -.0007522            | -.0003183 |
| tenure   | .0290875  | .006762   | 4.30  | 0.000 | .0158031             | .0423719  |
| tenursq  | -.0005331 | .0002312  | -2.31 | 0.022 | -.0009874            | -.0000789 |
| _cons    | .3213781  | .100009   | 3.21  | 0.001 | .1249041             | .5178521  |

Comments:

**Case 2** We can use dummy variables to represent multiple categories of a variable. Consider the relationship between law school rankings and starting salaries

$$\begin{aligned}\log(\text{salary}) = & \beta_0 + \delta_0 \text{top10} + \delta_1 r11\_25 + \delta_3 r26\_40 + \delta_4 r41\_60 + \beta_1 \text{LSAT} \\ & + \beta_2 \text{GPA} + \beta_3 \log(\text{libvol}) + \beta_4 \log(\text{cost}) + u.\end{aligned}$$

where *top10*, *r11\_25*, *r26\_40*, *r41\_60* would be equal to 1 when the variable *rank* falls into the appropriate range.

\*\* Rank below 60 would be the base case.

```
. regress lsalary top10 r11_25 r26_40 r41_60 LSAT GPA llibvol lcost
```

| Source   | SS         | df  | MS         | Number of obs = | 136    |
|----------|------------|-----|------------|-----------------|--------|
| Model    | 9.16538532 | 8   | 1.14567316 | F( 8, 127) =    | 120.15 |
| Residual | 1.2109665  | 127 | .009535169 | Prob > F =      | 0.0000 |
|          |            |     |            | R-squared =     | 0.8833 |
|          |            |     |            | Adj R-squared = | 0.8759 |
| Total    | 10.3763518 | 135 | .076861865 | Root MSE =      | .09765 |

| lsalary | Coef.    | Std. Err. | t     | P> t  | [95% Conf. Interval] |          |
|---------|----------|-----------|-------|-------|----------------------|----------|
| top10   | .5393428 | .053542   | 10.07 | 0.000 | .4333927             | .6452928 |
| r11_25  | .4716199 | .0390921  | 12.06 | 0.000 | .3942637             | .548976  |
| r26_40  | .2790977 | .0346972  | 8.04  | 0.000 | .2104383             | .3477571 |
| r41_60  | .182382  | .0283098  | 6.44  | 0.000 | .126362              | .238402  |
| LSAT    | .0060482 | .0034919  | 1.73  | 0.086 | -.0008616            | .012958  |
| GPA     | .1305893 | .0818678  | 1.60  | 0.113 | -.0314122            | .2925908 |
| llibvol | .0725522 | .0289213  | 2.51  | 0.013 | .0153221             | .1297824 |
| lcost   | .0249169 | .0283224  | 0.88  | 0.381 | -.031128             | .0809619 |
| _cons   | 8.363103 | .4457314  | 18.76 | 0.000 | 7.481081             | 9.245125 |

Comments:

## 6 Interactions involving dummy variables

### Case 1 Interactions among dummies

\*\* We can use interactions among dummies to account for the effect of each combination of dummies as well:

A different way to estimate eq.(8.1) is

$$\begin{aligned}\log(wage) = & \beta_0 + \delta_0 female + \delta_1 married + \delta_3 female \cdot married + \beta_1 educ \\ & + \beta_2 exper + \beta_3 exper^2 + \beta_4 tenure + \beta_5 tenure^2 + u.\end{aligned}$$

where  $female \cdot married = female \times married$ .

|                                                                                            |            |           |            |                        |                      |           |
|--------------------------------------------------------------------------------------------|------------|-----------|------------|------------------------|----------------------|-----------|
| <pre>. gen female_married = female*married</pre>                                           |            |           |            |                        |                      |           |
| <pre>. regress lwage female married female_married educ exper expersq tenure tenursq</pre> |            |           |            |                        |                      |           |
| Source                                                                                     | SS         | df        | MS         | Number of obs = 526    |                      |           |
| Model                                                                                      | 68.3617623 | 8         | 8.54522029 | F( 8, 517) = 55.25     |                      |           |
| Residual                                                                                   | 79.9679891 | 517       | .154676961 | Prob > F = 0.0000      |                      |           |
| Total                                                                                      | 148.329751 | 525       | .28253286  | R-squared = 0.4609     |                      |           |
|                                                                                            |            |           |            | Adj R-squared = 0.4525 |                      |           |
|                                                                                            |            |           |            | Root MSE = .39329      |                      |           |
| lwage                                                                                      | Coef.      | Std. Err. | t          | P> t                   | [95% Conf. Interval] |           |
| female                                                                                     | -.1103502  | .0557421  | -1.98      | 0.048                  | -.219859             | -.0008414 |
| married                                                                                    | .2126757   | .0553572  | 3.84       | 0.000                  | .103923              | .3214284  |
| female_married                                                                             | -.3005931  | .071767   | -4.19      | 0.000                  | -.4415838            | -.1596024 |
| educ                                                                                       | .0789103   | .0066945  | 11.79      | 0.000                  | .0657585             | .092062   |
| exper                                                                                      | .0268006   | .0052428  | 5.11       | 0.000                  | .0165007             | .0371005  |
| expersq                                                                                    | -.0005352  | .0001104  | -4.85      | 0.000                  | -.0007522            | -.0003183 |
| tenure                                                                                     | .0290875   | .006762   | 4.30       | 0.000                  | .0158031             | .0423719  |
| tenursq                                                                                    | -.0005331  | .0002312  | -2.31      | 0.022                  | -.0009874            | -.0000789 |
| _cons                                                                                      | .3213781   | .100009   | 3.21       | 0.001                  | .1249041             | .5178521  |

Comments:

**Case 2** Interaction between a dummy and a continuous variable

$$\begin{aligned}\log(wage) = & \beta_0 + \delta_0 female + \beta_1 educ + \delta_1 female \cdot educ + \beta_2 exper \\ & + \beta_3 exper^2 + \beta_4 tenure + \beta_5 tenure^2 + u.\end{aligned}$$

|                                                                                           |            |           |            |                        |                      |           |
|-------------------------------------------------------------------------------------------|------------|-----------|------------|------------------------|----------------------|-----------|
| <code>. gen female_educ = female*educ</code>                                              |            |           |            |                        |                      |           |
| <code>. regress lwage female married female_educ educ exper expersq tenure tenursq</code> |            |           |            |                        |                      |           |
| Source                                                                                    | SS         | df        | MS         | Number of obs = 526    |                      |           |
| Model                                                                                     | 65.677852  | 8         | 8.2097315  | F( 8, 517) = 51.35     |                      |           |
| Residual                                                                                  | 82.6518994 | 517       | .159868277 | Prob > F = 0.0000      |                      |           |
| Total                                                                                     | 148.329751 | 525       | .28253286  | R-squared = 0.4428     |                      |           |
|                                                                                           |            |           |            | Adj R-squared = 0.4342 |                      |           |
|                                                                                           |            |           |            | Root MSE = .39984      |                      |           |
| lwage                                                                                     | Coef.      | Std. Err. | t          | P> t                   | [95% Conf. Interval] |           |
| female                                                                                    | -.2197774  | .1675154  | -1.31      | 0.190                  | -.5488721            | .1093172  |
| married                                                                                   | .0529779   | .0407884  | 1.30       | 0.195                  | -.0271535            | .1331092  |
| female_educ                                                                               | -.0056186  | .0130532  | -0.43      | 0.667                  | -.0312625            | .0200254  |
| educ                                                                                      | .0813472   | .0085008  | 9.57       | 0.000                  | .0646469             | .0980476  |
| exper                                                                                     | .0268542   | .005335   | 5.03       | 0.000                  | .0163733             | .0373351  |
| expersq                                                                                   | -.0005375  | .0001124  | -4.78      | 0.000                  | -.0007583            | -.0003167 |
| tenure                                                                                    | .0314803   | .0068669  | 4.58       | 0.000                  | .0179898             | .0449709  |
| tenursq                                                                                   | -.0005792  | .0002351  | -2.46      | 0.014                  | -.0010412            | -.0001173 |
| _cons                                                                                     | .389629    | .1186101  | 3.28       | 0.001                  | .1566119             | .6226461  |

Comments:

## 7 Testing for Differences in Regression Functions across Groups

- Is it reasonable to believe that the population regression function that explains the dependent variable is the same across subsamples of populations?
- For example, is it reasonable to believe that the function that explains "GPA of college athlete" is the same for male and female students?

Consider

$$\text{cumgpa} = \beta_0 + \beta_1 \text{sat} + \beta_2 \text{hsperc} + \beta_3 \text{tothrs} + u,$$

where

*cumgpa* = cumulative GPA

*sat* = SAT score

*hsperc* = high school rank percentile

*tothrs* = total hours of college courses

- If we want to test whether "male" students and "female" students have the same values of  $\beta_0, \beta_1, \beta_2, \beta_3$  we can estimate the following model

$$\begin{aligned} \text{cumgpa} = & \beta_0 + \delta_0 \text{female} + \beta_1 \text{sat} + \delta_1 \text{female} \cdot \text{sat} + \beta_2 \text{hsperc} + \delta_2 \text{female} \cdot \text{hsperc} \\ & + \beta_3 \text{tothrs} + \delta_3 \text{female} \cdot \text{tothrs} + u, \end{aligned}$$

and the null hypothesis would be

$$\begin{aligned} H_0 &: \delta_0 = \delta_1 = \delta_2 = \delta_3 = 0 \\ H_a &: \text{otherwise (at least one } \delta_j = 0) \end{aligned} \quad (8.2)$$

We can use the F-test to test this type of null hypothesis:

1. The restricted model (r)

**. regress cumgpa sat hsperc tothrs**

| Source   | SS         | df  | MS         | Number of obs = | 732    |
|----------|------------|-----|------------|-----------------|--------|
| Model    | 168.533658 | 3   | 56.1778861 | F( 3, 728) =    | 74.72  |
| Residual | 547.364897 | 728 | .751874858 | Prob > F =      | 0.0000 |
|          |            |     |            | R-squared =     | 0.2354 |
|          |            |     |            | Adj R-squared = | 0.2323 |
| Total    | 715.898555 | 731 | .979341389 | Root MSE =      | .86711 |

| cumgpa | Coef.     | Std. Err. | t     | P> t  | [95% Conf. Interval] |           |
|--------|-----------|-----------|-------|-------|----------------------|-----------|
| sat    | .0009028  | .0002079  | 4.34  | 0.000 | .0004947             | .0013109  |
| hsperc | -.0063791 | .0015678  | -4.07 | 0.000 | -.0094572            | -.0033011 |
| tothrs | .0119779  | .0009314  | 12.86 | 0.000 | .0101494             | .0138064  |
| _cons  | .9291105  | .2285515  | 4.07  | 0.000 | .4804118             | 1.377809  |

## 2. The unrestricted model (ur)

```
. gen female_sat = female*sat
. gen female_hspc = female*hspc
. gen female_tothrs = female*tothrs
. regress cumgpa female sat female_sat hspc female_hspc tothrs female_tothrs
```

| Source   | SS         | df  | MS         | Number of obs = 732    |  |  |
|----------|------------|-----|------------|------------------------|--|--|
| Model    | 181.589407 | 7   | 25.9413439 | F( 7, 724) = 35.15     |  |  |
| Residual | 534.309148 | 724 | .73799606  | Prob > F = 0.0000      |  |  |
|          |            |     |            | R-squared = 0.2537     |  |  |
|          |            |     |            | Adj R-squared = 0.2464 |  |  |
| Total    | 715.898555 | 731 | .979341389 | Root MSE = .85907      |  |  |

  

| cumgpa        | Coef.     | Std. Err. | t     | P> t  | [95% Conf. Interval] |           |
|---------------|-----------|-----------|-------|-------|----------------------|-----------|
| female        | -1.113638 | .528539   | -2.11 | 0.035 | -2.15129             | -.0759859 |
| sat           | .0006113  | .000235   | 2.60  | 0.009 | .0001499             | .0010727  |
| female_sat    | .0011167  | .0005     | 2.23  | 0.026 | .0001351             | .0020984  |
| hspc          | -.0059675 | .0017765  | -3.36 | 0.001 | -.0094551            | -.0024798 |
| female_hspc   | .0000508  | .0041025  | 0.01  | 0.990 | -.0080035            | .008105   |
| tothrs        | .0103004  | .0010928  | 9.43  | 0.000 | .0081549             | .0124459  |
| female_tothrs | .0055599  | .0020696  | 2.69  | 0.007 | .0014968             | .009623   |
| _cons         | 1.213984  | .2648281  | 4.58  | 0.000 | .6940617             | 1.733907  |

Comments:

7.1 We can use the "Chow statistics" to test this type of hypothesis as well

- When there are many variables in the model, adding an interaction for every explanatory variable would make the regression analysis messy.
- In which case, we can use the "Chow test" or "Chow statistic" to test the hypothesis expressed in (8.2).
- Chow statistic is a type of F-statistic.

$$F = \frac{SSR_p - (SSR_1 + SSR_2)}{SSR_1 + SSR_2} \cdot \frac{[n - 2(k + 1)]}{k + 1},$$

where  $n$  is the total number of observations.

$SSR_p = SSR$  from the pooled model (include observations from both subsamples)

$SSR_1 = SSR$  from subsample 1

$SSR_2 = SSR$  from subsample 2

`regress cumgpa sat hsperc tothrs if female == 0`

| Source   | SS         | df  | MS         | Number of obs = | 552    |
|----------|------------|-----|------------|-----------------|--------|
| Model    | 89.6937042 | 3   | 29.8979014 | F( 3, 548) =    | 41.94  |
| Residual | 390.619421 | 548 | .712809162 | Prob > F =      | 0.0000 |
|          |            |     |            | R-squared =     | 0.1867 |
|          |            |     |            | Adj R-squared = | 0.1823 |
| Total    | 480.313125 | 551 | .871711661 | Root MSE =      | .84428 |

| cumgpa | Coef.     | Std. Err. | t     | P> t  | [95% Conf. Interval] |          |
|--------|-----------|-----------|-------|-------|----------------------|----------|
| sat    | .0006113  | .000231   | 2.65  | 0.008 | .0001576             | .001065  |
| hsperc | -.0059675 | .0017459  | -3.42 | 0.001 | -.0093969            | -.002538 |
| tothrs | .0103004  | .001074   | 9.59  | 0.000 | .0081907             | .0124101 |
| _cons  | 1.213984  | .2602697  | 4.66  | 0.000 | .7027359             | 1.725233 |

`regress cumgpa sat hsperc tothrs if female == 1`

| Source   | SS         | df  | MS         | Number of obs = | 180    |
|----------|------------|-----|------------|-----------------|--------|
| Model    | 83.4816253 | 3   | 27.8272084 | F( 3, 176) =    | 34.08  |
| Residual | 143.689727 | 176 | .816418902 | Prob > F =      | 0.0000 |
|          |            |     |            | R-squared =     | 0.3675 |
|          |            |     |            | Adj R-squared = | 0.3567 |
| Total    | 227.171352 | 179 | 1.2691137  | Root MSE =      | .90356 |

| cumgpa | Coef.     | Std. Err. | t     | P> t  | [95% Conf. Interval] |          |
|--------|-----------|-----------|-------|-------|----------------------|----------|
| sat    | .0017281  | .0004642  | 3.72  | 0.000 | .0008119             | .0026442 |
| hsperc | -.0059167 | .0038895  | -1.52 | 0.130 | -.0135927            | .0017594 |
| tothrs | .0158603  | .0018485  | 8.58  | 0.000 | .0122122             | .0195085 |
| _cons  | .1003465  | .4810947  | 0.21  | 0.835 | -.8491105            | 1.049803 |

Comments:

## 8 A Binary Dependent Variable (y variable): The Linear Probability Model

- So far, our Y variables are continuous.
- What if we are interested in explaining a qualitative Y variable (that is, Y is a dummy variable)?

Consider

$$\begin{aligned}y &= \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + u \\E(y|\mathbf{x}) &= \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k,\end{aligned}$$

where  $\mathbf{x}$  denotes all of the explanatory variables  $(x_1, \dots, x_k)$ .

```
. regress inlf nwifeinc educ exper expersq age kidslt6 kidsge6
```

| Source   | SS         | df  | MS         | Number of obs = 753    |  |  |
|----------|------------|-----|------------|------------------------|--|--|
| Model    | 48.8080578 | 7   | 6.97257969 | F( 7, 745) = 38.22     |  |  |
| Residual | 135.919698 | 745 | .182442547 | Prob > F = 0.0000      |  |  |
|          |            |     |            | R-squared = 0.2642     |  |  |
|          |            |     |            | Adj R-squared = 0.2573 |  |  |
| Total    | 184.727756 | 752 | .245648611 | Root MSE = .42713      |  |  |

  

| inlf     | Coef.     | Std. Err. | t     | P> t  | [95% Conf. Interval] |           |
|----------|-----------|-----------|-------|-------|----------------------|-----------|
| nwifeinc | -.0034052 | .0014485  | -2.35 | 0.019 | -.0062488            | -.0005616 |
| educ     | .0379953  | .007376   | 5.15  | 0.000 | .023515              | .0524756  |
| exper    | .0394924  | .0056727  | 6.96  | 0.000 | .0283561             | .0506287  |
| expersq  | -.0005963 | .0001848  | -3.23 | 0.001 | -.0009591            | -.0002335 |
| age      | -.0160908 | .0024847  | -6.48 | 0.000 | -.0209686            | -.011213  |
| kidslt6  | -.2618105 | .0335058  | -7.81 | 0.000 | -.3275875            | -.1960335 |
| kidsge6  | .0130122  | .013196   | 0.99  | 0.324 | -.0128935            | .0389179  |
| _cons    | .5855192  | .154178   | 3.80  | 0.000 | .2828442             | .8881943  |

where

*inlf* = 1 if the woman reports working for a wage outside the home at some point during the year, zero otherwise.

*nwifeinc* = husband's earnings (in thousands of dollars)

*educ* = years of education

*exper* = past years of labor market experience

*age* = age

*kidslt6* = number of children less than 6 years old

*kidsage6* = number of kinds between 6 - 18 years old

