

1. Which of the following can cause the usual OLS t statistics to be invalid (that is, not to have t distributions under H_0)?

- Heteroskedasticity.
- A sample correlation coefficient of .95 between two independent variables that are in the model.
- Omitting an important explanatory variable.

i) Yes, violate the assumption of homoskedasticity.

ii) No, it just requires the coefficient not to be 1.

iii) Yes, violate the 4th assumption $E(u|x) = 0$ if the omitted variable is correlated with the independent variable in the model.

2. Consider an equation to explain salaries of CEOs in terms of annual firm sales, return on equity (roe , in percentage form), and return on the firm's stock (ros , in percentage form):

$$\log(\text{salary}) = \beta_0 + \beta_1 \log(\text{sales}) + \beta_2 roe + \beta_3 ros + u.$$

i. In terms of the model parameters, state the null hypothesis that, after controlling for $sales$ and roe , ros has no effect on CEO salary. State the alternative that better stock market performance increases a CEO's salary.

ii. Using the data in CEOSAL1, the following equation was obtained by OLS:

$$\widehat{\log(\text{salary})} = 4.32 + .280 \log(\text{sales}) + .0174 roe + .00024 ros$$

$$(.32) \quad (.035) \quad (.0041) \quad (.00054)$$

$$n = 209, R^2 = .283.$$

By what percentage is $salary$ predicted to increase if ros increases by 50 points? Does ros have a practically large effect on $salary$?

iii. Test the null hypothesis that ros has no effect on $salary$ against the alternative that ros has a positive effect. Carry out the test at the 10% significance level.

iv. Would you include ros in a final model explaining CEO compensation in terms of firm performance? Explain.

i). $H_0: \beta_3 = 0$

$H_a: \beta_3 > 0$

ii) the proportionate effect on $\widehat{\text{salary}} = 0.00024(50) = 0.12 = 1.2\%$.

Thus, a 50 point ceteris paribus increase in ros is predicted to increase salary by 1.2%.

iii) $H_0: \beta_3 = 0$

$H_a: \beta_3 > 0$

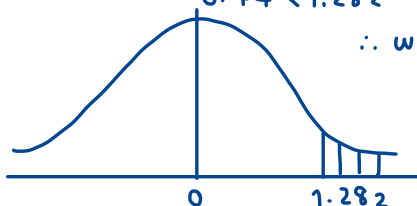
significant level 10% = 0.1; d.f. = 209 - 3 - 1 = 205

$t_{crit} = 1.282$

$t_{cal} = \frac{\hat{\beta}_3 - 0}{s.e.(\hat{\beta}_3)} = \frac{0.0024}{0.0054} = 0.44$

$0.44 < 1.282$

\therefore we can't reject H_0 at 10% significant level.



C1. The following model can be used to study whether campaign expenditures affect election outcomes:

$$voteA = \beta_0 + \beta_1 \log(expendA) + \beta_2 \log(expendB) + \beta_3 prtyst rA + u,$$

where $voteA$ is the percentage of the vote received by Candidate A, $expendA$ and $expendB$ are campaign expenditures by Candidates A and B, and $prtyst rA$ is a measure of party strength for Candidate A (the percentage of the most recent presidential vote that went to A's party).

- What is the interpretation of β_1 ?
- In terms of the parameters, state the null hypothesis that a 1% increase in A's expenditures is offset by a 1% increase in B's expenditures.

. reg voteA lexpendA lexpendB prtyst rA

Source	SS	df	MS	Number of obs	=	173
Model	38405.1096	3	12801.7032	F(3, 169)	=	215.23
Residual	10052.1389	169	59.480112	Prob > F	=	0.0000
				R-squared	=	0.7926
				Adj R-squared	=	0.7889
Total	48457.2486	172	281.728189	Root MSE	=	7.7123

voteA	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
lexpendA	6.083316	.38215	15.92	0.000	5.328914 6.837719
lexpendB	-6.615417	.3788203	-17.46	0.000	-7.363246 -5.867588
prtyst rA	.1519574	.0620181	2.45	0.015	.0295274 .2743873
_cons	45.07893	3.926305	11.48	0.000	37.32801 52.82985

iv) rewrite hypothesis $\theta_1 = \beta_1 + \beta_2 \rightarrow H_0: \theta_1 = 0$
 $H_a: \theta_1 \neq 0$

rearrange equation $\hat{vote} A = \beta_0 + \theta \log(Expend A) + \beta_2 (\log(Expend B - \log(Expend A))) + \beta_3 prtyst rA$
 when estimate equation we obtain $\hat{\beta}_1 \approx -0.532$
 s.e. $(\hat{\theta}_1) \approx .533$

then $t_{cal} = \frac{-0.532 - 0}{0.533} \approx -1 \therefore$ cannot reject $H_0: \beta_2 = \beta_1$

C6. Use the data in WAGE2 for this exercise.

- Consider the standard wage equation

$$\log(wage) = \beta_0 + \beta_1 educ + \beta_2 exper + \beta_3 tenure + u.$$

State the null hypothesis that another year of general workforce experience has the same effect on $\log(wage)$ as another year of tenure with the current employer.

- Test the null hypothesis in part (i) against a two-sided alternative, at the 5% significance level, by constructing a 95% confidence interval. What do you conclude?

i) $H_0: \beta_2 = \beta_3$

ii) d is significant at a 95% of confidence interval, thus we can't reject H_0 . we can conclude that one additional year of general workforce experience has the same effect on $\log(wage)$ as another year.

. reg lwage educ exper tenure

Source	SS	df	MS	Number of obs	=	935
Model	25.6953242	3	8.56510806	F(3, 931)	=	56.97
Residual	139.960959	931	.150334005	Prob > F	=	0.0000
				R-squared	=	0.1551
				Adj R-squared	=	0.1524
Total	165.656283	934	.177362188	Root MSE	=	.38773

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	.0748638	.0065124	11.50	0.000	.062083 .0876446
exper	.0153285	.0033696	4.55	0.000	.0087156 .0219413
tenure	.0133748	.0025872	5.17	0.000	.0082974 .0184522
_cons	5.496696	.1105282	49.73	0.000	5.279782 5.713609

- Estimate the given model using the data in VOTE1 and report the results in usual form. Do A's expenditures affect the outcome? What about B's expenditures? Can you use these results to test the hypothesis in part (ii)?

- Estimate a model that directly gives the t statistic for testing the hypothesis in part (ii). What do you conclude? (Use a two-sided alternative.)

$$\begin{aligned} i) \Delta vote A &= \beta_1 \Delta \log(Expend A) \\ &= (\beta_1 / 100) [100 \times \Delta \log(Expend A)] \\ &\approx (\beta_1 / 100) [\% \Delta \log(Expend A)] \end{aligned}$$

$\therefore \beta_1 / 100$ is ceteris paribus percentage point change of vote received when campaign expenditure by candidate A increase by 1%.

ii) $H_0: \beta_2 = -\beta_1$
 $H_a: \beta_2 \neq -\beta_1$

iii) Regression model in usual form:

$$vote A = (45.1) + 6.08 \log(Expend A) - 6.62 \log(Expend B)$$

C8. The data set 401KSUBS contains information on net financial wealth (*nettfa*), age of the survey respondent (*age*), annual family income (*inc*), family size (*fsize*), and participation in certain pension plans for people in the United States. The wealth and income variables are both recorded in thousands of dollars. For this question, use only the data for single-person households (so *fsize* = 1).

- i. How many single-person households are there in the data set?
- ii. Use OLS to estimate the model

$$nettfa = \beta_0 + \beta_1 inc + \beta_2 age + u,$$

and report the results using the usual format. Be sure to use only the single-person households in the sample. Interpret the slope coefficients. Are there any surprises in the slope estimates?

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. reg nettfa inc age if fsize ==1
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Source	SS	df	MS	Number of obs	=	2,017
Model	544916.989	2	272458.495	F(2, 2014)	=	136.46
Residual	4021048.06	2,014	1996.54819	Prob > F	=	0.0000
				R-squared	=	0.1193
				Adj R-squared	=	0.1185
Total	4565965.05	2,016	2264.86361	Root MSE	=	44.683

nettfa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
inc	.7993167	.0597307	13.38	0.000	.6821762 .9164572
age	.8426563	.0920169	9.16	0.000	.6621982 1.023115
_cons	-43.03981	4.080393	-10.55	0.000	-51.04204 -35.03758

i) 2.017 observations

ii) $\hat{\beta}_1$ can be interpret as a \$1,000 increase in income correspond to a \$799 increase in net financial wealth we interpret $\hat{\beta}_2$ as a 1 year increase in age correspond to a \$842 increase in net financial wealth.

iii) $\hat{\beta}_0 = -43.04$

This is an individual's net financial wealth when their income is \$0 and their age is 0. In other word, a net financial wealth of new born babies.

iv) t-statistic = $(0.843 - 1 / 0.092) \approx -1.71$

p-value = $P(T < -1.71) \approx 0.044$

Against the one-sided alternative $H_1: \beta_2 < 1$

\therefore reject the null hypothesis at the 5% significance level but not at the 1% significance level.

iii. Does the intercept from the regression in part (ii) have an interesting meaning? Explain.

iv. Find the p -value for the test $H_0: \beta_2 = 1$ against $H_1: \beta_2 < 1$. Do you reject H_0 at the 1% significance level?

v. If you do a simple regression of *nettfa* on *inc*, is the estimated coefficient on *inc* much different from the estimate in part (ii)? Why or why not?