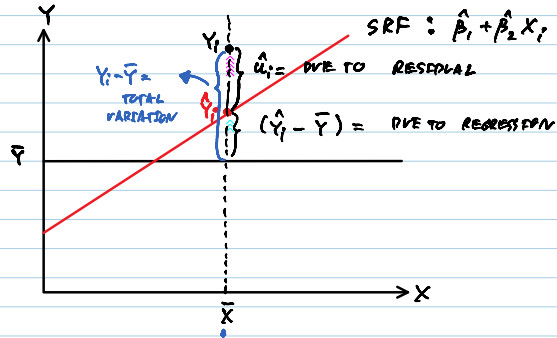


ANALYSIS OF VARIANCE



$TSS = ESS + RSS$

TSS = TOTAL SUM OF SQUARES

$= \sum (Y_i - \bar{Y})^2$

ESS = EXPLAINED SUM OF SQUARES

$= \sum (\hat{Y}_i - \bar{Y})^2 = \beta_2^2 \sum x_i^2$

RSS = RESIDUAL SUM OF SQUARES

$= \sum (Y_i - \hat{Y}_i)^2$   
OR  $= \sum \hat{u}_i^2$

$\sum y_i^2 = \sum \hat{y}_i^2 + \sum \hat{u}_i^2$

$\sum y_i^2 = \beta_2^2 \sum x_i^2 + \sum \hat{u}_i^2$

ANOVA TABLE

ANALYSIS OF VARIANCE

SUM OF SQUARES

MEAN SUM OF SQUARES =  $\frac{SS}{df}$

SOURCE OF VARIATION	SS	df	MSS
DUE TO REGRESSION (ESS)	$\beta_2^2 \sum x_i^2$	1	$\beta_2^2 \sum x_i^2 / 1$
DUE TO RESIDUALS (RSS)	$\sum \hat{u}_i^2$	n-2	$\sum \hat{u}_i^2 / n-2 = \hat{\sigma}^2$
TSS	$\sum y_i^2$	n-1	$\sum y_i^2 / n-1$

- ①
- ② (OUR OLD KNOWLEDGE)
- ③

THEREFORE,

$F = \frac{MSS \text{ OF ESS}}{MSS \text{ OF RSS}}$   
 $= \frac{\beta_2^2 \sum x_i^2}{\sum \hat{u}_i^2 / (n-2)}$   
 $= \frac{\beta_2^2 \sum x_i^2}{\hat{\sigma}^2}$

~ F-DISTRIBUTION WITH 1 df in the NUMERATOR AND (n-2) df IN THE DENOMINATOR.

DECISION RULE: IF  $F^* >$  CRITICAL F VALUE ( $F_{\alpha, 1, n-2}$ ), THEN REJECT  $H_0$ . (i.e., X COULD BE USED TO EXPLAIN Y)

IF  $F^* <$  CRITICAL F VALUE ( $F_{\alpha, 1, n-2}$ ) THEN, FAIL TO REJECT  $H_0$ .

PRACTICE

$\alpha = 0.05$	$\alpha = 0.01$
FIND $F_{\alpha, 1, 5} = 6.61$	FIND $F_{\alpha, 1, 5} = 16.3$
$F_{\alpha, 1, 8} = \dots$	$F_{\alpha, 1, 8} = \dots$
$F_{\alpha, 2, 5} = \dots$	$F_{\alpha, 2, 5} = \dots$

$$F_{\alpha, 2, 5} = \dots \quad F_{\alpha, 2, 5} = \dots$$

$$F_{\alpha, 3, 5} = \dots \quad F_{\alpha, 3, 5} = \dots$$

BACK TO OUR WAGE - EDU EXAMPLE:

$H_0: \beta_2 = 0$  (EDU CANNOT BE USED TO EXPLAIN WAGE)

$H_1: \beta_2 \neq 0$  (EDU CAN BE USED TO EXPLAIN WAGE)

$n = 13$

**CALCULATION**

SOURCE OF VARIATION	SS	df	MSS
ESS	95.4255	1	95.4255 (1)
RSS	9.6428	$n-2 = 11$	0.8811 (2)
TSS	105.1183	$n-1 = 12$	

$$F^* = \frac{\text{MSS OF ESS}}{\text{MSS OF RSS}} = \frac{95.4255}{0.8811} = 108.3026$$

CRITICAL F VALUE ( $F_{0.05, 1, 11}$ ) = 4.84,

SINCE  $F^* >$  CRITICAL F, THEN WE REJECT  $H_0$ .

IT IMPLIES THAT EDU STATISTICALLY INFLUENCE WAGE.

EXTENSIONS OF THE TWO-VARIABLE LINEAR REGRESSION MODEL.

(CHAPTER 6)

AGENDA: (I) REGRESSION THROUGH THE ORIGIN (R-T-O) OR INTERCEPTLESS MODEL:

$$Y_i = \beta_2 X_i + u_i$$

(II) SCALING AND UNITS OF MEASUREMENT

(III) FUNCTIONAL FORMS OF REGRESSION MODELS

(I) REGRESSION THROUGH THE ORIGIN (R-T-O)

$$Y_i = \beta_2 X_i + u_i \quad (\text{PRF})$$

EX: CAPM MODEL

CAPITAL ASSET PRICING MODEL

WE OBTAIN  $\hat{Y}_i = \hat{\beta}_2 X_i$  VIA OLS

OLS ESTIMATOR FOR  $\beta_2$

$$\hat{u}_i = Y_i - \hat{Y}_i$$

$$\hat{u}_i = Y_i - \hat{\beta}_2 X_i$$

$$\sum \hat{u}_i^2 = \sum (Y_i - \hat{\beta}_2 X_i)^2 \rightarrow \text{SUM OF SQUARED ERROR.}$$

$$\sum \hat{u}_i^2 = \sum (Y_i - \hat{\beta}_2 X_i)^2 \rightarrow \text{SUM OF SQUARED ERROR.}$$

$$\text{MIN}_{\hat{\beta}_2} \sum \hat{u}_i^2 = \sum (Y_i - \hat{\beta}_2 X_i)^2$$

"OUR OBJECTIVE FUNCTION"

FONC : FIRST-ORDERED NECESSARY CONDITION

$$\frac{\partial \sum \hat{u}_i^2}{\partial \hat{\beta}_2} = 2 \sum (Y_i - \hat{\beta}_2 X_i)(-X_i) = 0$$

$$\sum \hat{u}_i X_i = 0$$

FROM

$$2 \sum (Y_i - \hat{\beta}_2 X_i)(-X_i) = 0$$

$$\sum (Y_i - \hat{\beta}_2 X_i)(-X_i) = 0$$

$$\sum X_i Y_i - \hat{\beta}_2 \sum X_i^2 = 0$$

$$\hat{\beta}_2 = \frac{\sum X_i Y_i}{\sum X_i^2}$$

\*\*\* R-T-O MODEL

$$\hat{\beta}_2 = \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2}$$

MODEL W/ INTERCEPT

Q: IS  $\hat{\beta}_2$  UNBIASED?

A: CHECK IT OUT...

$$\hat{\beta}_2 = \frac{\sum X_i Y_i}{\sum X_i^2}$$

$$\hat{\beta}_2 = \frac{\sum X_i (\beta_2 X_i + u_i)}{\sum X_i^2}$$

$$\hat{\beta}_2 = \frac{\sum X_i \beta_2 X_i}{\sum X_i^2} + \frac{\sum X_i u_i}{\sum X_i^2}$$

$$\hat{\beta}_2 = \beta_2 + \frac{\sum X_i u_i}{\sum X_i^2}$$

$$E(\hat{\beta}_2) = \beta_2$$

$$\because E\left[\frac{\sum X_i u_i}{\sum X_i^2}\right] = 0$$