

Limited Dependent Variable Models

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Part 4: Tobit model for corner solution responses, Censored and Truncated regression models

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Corner solution response

- ▶ A dependent variable is zero for a nontrivial fraction of the population, but is continuously distributed over positive values
- ▶ Example: the amount of expenditure an individual spends on alcohol in a given month
 - ▶ For some significant fraction of the population, the amount spent on alcohol is 0
- ▶ What if we use a linear model for this dependent variable?
 - ▶ We could possibly obtain negative fitted values, similar to the problems with LPM for binary outcomes
 - ▶ The distribution of y piles up at zero. So, y cannot have a conditional normal distribution as in a linear model

Tobit model

- ▶ Define an underlying latent variable:

$$y^* = \beta_0 + \mathbf{x}\beta + u, u|x \sim \text{Normal}(0, \sigma^2) \quad (1)$$

- ▶ Our observed dependent variable:

$$y = \max(0, y^*) \quad (2)$$

- ▶ y^* has a normal, homoskedastic distribution with a linear conditional mean (the classical linear model assumptions)
- ▶ y has a continuous distribution over strictly positive values

Tobit model

- ▶ The density of y given x is the same as the density of y^* given x for positive values
- ▶ For $y=0$, $P(y = 0|x) = P(y^* < 0|x) = P\left(\frac{u}{\sigma} < -\frac{x\beta}{\sigma} | x\right) = \Phi\left(-\frac{x\beta}{\sigma}\right) = 1 - \Phi\left(\frac{x\beta}{\sigma}\right)$ (3)
- ▶ For $y>0$, the density of y_i given x_i is
$$\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(y-x_i\beta)^2}{2\sigma^2}\right] = \frac{1}{\sigma} \phi\left[\frac{y-x_i\beta}{\sigma}\right]$$
 (4)
- ▶ Log likelihood is

Interpreting the Tobit estimates

- ▶ In Tobit models, we have 2 expectations of interest:
- ▶ $E(y|y > 0, \mathbf{x})$: 'conditional expectation' as it is conditional on $y > 0$
- ▶ $E(y|\mathbf{x})$: 'unconditional expectation'
- ▶ Given $E(y|y > 0, \mathbf{x})$, we can find
 $E(y|\mathbf{x}) = P(y > 0|\mathbf{x})E[y|y > 0, \mathbf{x}]$ (5)
- ▶ Note: if $z \sim N(0, 1)$, then $E[z|z > c] = \frac{\phi(c)}{1-\Phi(c)}$ for any constant c

Interpreting the Tobit estimates

- ▶ $E(y|y > 0, \mathbf{x}) = \mathbf{x}\beta + \sigma\lambda\left(\frac{\mathbf{x}\beta}{\sigma}\right)$, $\lambda\left(\frac{\mathbf{x}\beta}{\sigma}\right) = \frac{\phi(\mathbf{x}\beta/\sigma)}{\Phi(\mathbf{x}\beta/\sigma)}$ “inverse Mills ratio” (6)
- ▶ $E(y|\mathbf{x}) = P(y > 0|\mathbf{x})E[y|y > 0, \mathbf{x}] = \Phi\left(\frac{\mathbf{x}\beta}{\sigma}\right)\mathbf{x}\beta + \sigma\phi\left(\frac{\mathbf{x}\beta}{\sigma}\right)$ (7)
- ▶ Partial effects of x_j on $E[y|y > 0, \mathbf{x}]$:
$$\frac{\partial E[y|y > 0, \mathbf{x}]}{\partial x_j} = \beta_j \left\{ 1 - \lambda\left(\frac{\mathbf{x}\beta}{\sigma}\right) \left[\frac{\mathbf{x}\beta}{\sigma} - \lambda\left(\frac{\mathbf{x}\beta}{\sigma}\right) \right] \right\}$$
 (8)
 - ▶ same sign with β_j
- ▶ Partial derivative of $E[y|\mathbf{x}]$ with respect to x_j :
$$\frac{\partial E[y|\mathbf{x}]}{\partial x_j} = \beta_j \Phi\left(\frac{\mathbf{x}\beta}{\sigma}\right)$$
 (9)

Interpreting the Tobit estimates

- ▶ The partial effect at the average: evaluating the adjustment (scale) factor, $\Phi(\bar{\mathbf{x}}\hat{\beta}/\hat{\sigma})$
- ▶ The average partial effect: compute the scale factor as $\frac{1}{n} \sum_{i=1}^n \Phi(\mathbf{x}_i\hat{\beta}/\hat{\sigma})$
- ▶ Notice that the scale factor is always between zero and one: $0 < \Phi(\mathbf{x}\hat{\beta}/\hat{\sigma}) < 1$ since $\hat{P}(y_i > 0 | \mathbf{x}_i) = \Phi(\bar{\mathbf{x}}\hat{\beta}/\hat{\sigma})$
 - ▶ If there are few observations with $y_i = 0$, the scale factor will be closer to one.

Censored and Truncated Regression Models

- ▶ Censoring: missing data on the response variable
- ▶ Truncation: missing data on both the response variable and explanatory variables

Censored Regression Models

- ▶ Let y follows the classical linear model:

$$y_i = \mathbf{x}_i\beta + u_i, u_i|x_i, c_i \sim N(0, \sigma^2) \quad (10)$$

$$w_i = \min(y_i, c_i) \quad (11)$$

- ▶ We observe y_i only if it is less than a censoring value, c_i
 - ▶ Assume u_i is independent of c_i
 - ▶ This setup is right censoring, e.g. top coding data
- ▶ If using OLS regression, it will use only the uncensored observations ($y_i < c_i$), producing inconsistent estimators of β_j
- ▶ This is similar to corner solution response (in Tobit model with economic behavior resulting in zero outcomes), but here it is a data collection problem. For some reason, the data are censored!

Censored Regression Models

- ▶ Estimating by using maximum likelihood
- ▶ For uncensored observations ($w_i = y_i$), the density of w_i is the same as that of y_i : $N(\mathbf{x}_i\beta, \sigma^2)$
- ▶ For censored observations ($w_i = c_i$),
$$P(w_i = c_i | \mathbf{x}_i) = P(y_i \geq c_i | \mathbf{x}_i) = P(u_i \geq c_i - \mathbf{x}_i\beta) = 1 - \Phi[(c_i - \mathbf{x}_i\beta)/\sigma] \quad (12)$$
- ▶ Then combine these two parts, maximize the sum of these across i , with respect to β_j and σ to obtain the MLEs.

Truncated Regression Models

- ▶ Let y follows the classical linear model:
$$y = \mathbf{x}\beta + u, u|\mathbf{x} \sim N(0, \sigma^2) \quad (13)$$
- ▶ A random draw (\mathbf{x}_i, y_i) is observed only if $y_i \leq c_i$, where c_i is the truncation threshold that can depend on exogenous variables.
- ▶ To estimate β_j and σ , we need the distribution of y_i , given that $y_i \leq c_i$, and \mathbf{x}_i :

$$g(y|\mathbf{x}_i, c_i) = \frac{f(y|\mathbf{x}_i\beta, \sigma^2)}{F(c_i|\mathbf{x}_i\beta, \sigma^2)} \quad (14)$$

- ▶ $f(\cdot)$ the normal density with mean $\mathbf{x}_i\beta$ and variance σ^2
- ▶ $F(\cdot)$ the normal cdf with the same mean and variance, evaluated at c_i
- ▶ We renormalize the density by dividing by the area under $f(y|\mathbf{x}_i\beta, \sigma^2)$ that is to the left of c_i