

## EE325 Section 2 HW 5 Answers

### 10.12

(a) **False.** If exact linear relationship(s) exist among variables, we cannot even estimate the coefficients or their standard errors.

(b) **False.** One may be able to obtain one or more significant t values.

(c) **False.** As noted in the chapter (see Eq. 7.5.6), the variance of an OLS estimator is given by the following formula:

$$\text{var}(\hat{\beta}_j) = \frac{\sigma^2}{\sum x_j^2} \left( \frac{1}{1 - R_j^2} \right)$$

As can be seen from this formula, a high  $R_j^2$  can be counterbalanced by a low  $\sigma^2$  or high  $\sum x_j^2$ .

(d) **Uncertain.** If a model has only two regressors, high pairwise correlation coefficients may suggest multicollinearity. If one or more regressors enter non-linearly, the pairwise correlations may give misleading answers.

(e) **Uncertain.** If the observed collinearity continues in the future sample values, then there may be no harm. But if that is not the case or if the objective is precise estimation, then multicollinearity may be problem.

(f) **False.** See answer to (c) above.

(g) **False.** VIF and TOL provide the same information.

(h) **False.** One usually obtains high  $R^2$ 's in models with highly correlated regressors.

(i) **True.** As you can see from the formula given in (c), if the variability in  $X_3$  is small,  $R_j^2$  will tend to be small and in the extreme case of no variability in  $X_3$ ,  $\sum x_{3i}^2$  will be zero, in which case the variance of the estimated  $\beta_3$  will be infinite.

**11.1** (a) **False.** The estimators are unbiased but are inefficient.

(b) **True.**

(c) **False.** Typically, but not always, will the variance be overestimated.

(d) **False.** Besides heteroscedasticity, such a pattern may result from autocorrelation, model specification errors, etc.

(e) **True.**

(f) **True.**

(g) **False.** Heteroscedasticity is about the variance of the error term  $u_i$  and not about the variance of a regressor.

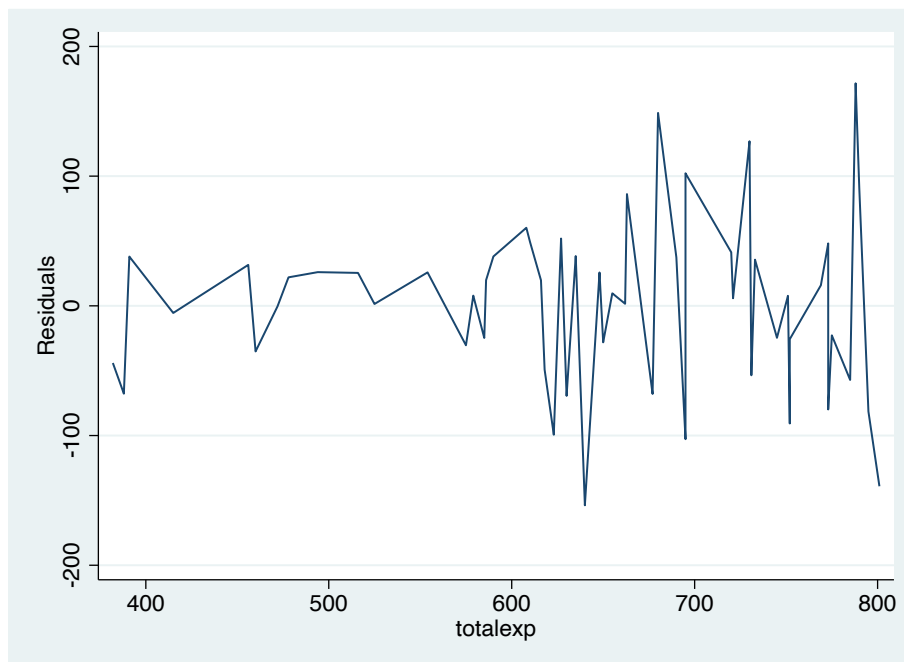
### 11.16

a.

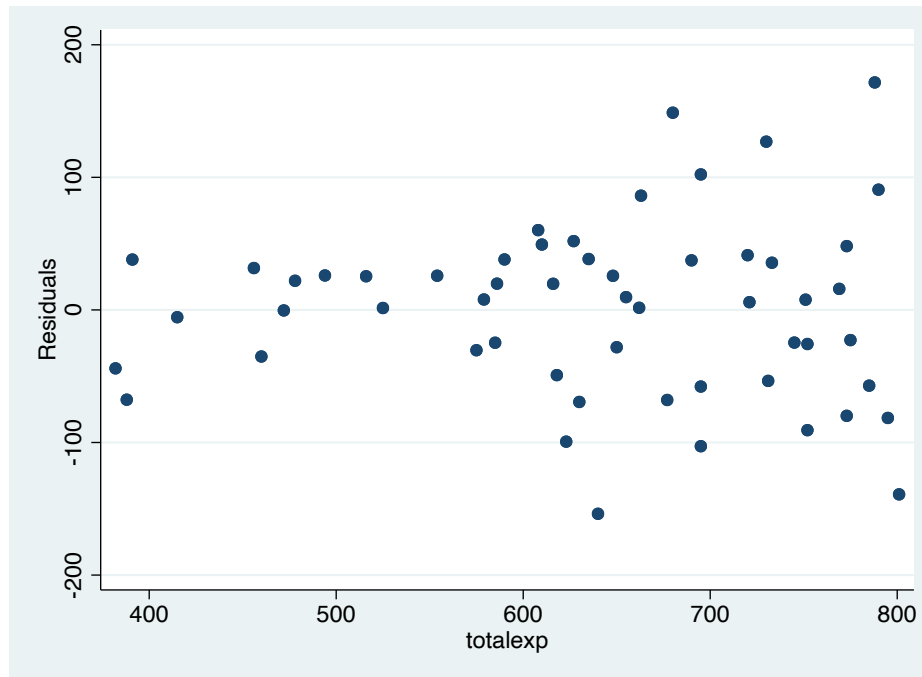
Source	SS	df	MS	Number of obs	=	55
Model	<b>139022.82</b>	<b>1</b>	<b>139022.82</b>	F(1, 53)	=	<b>31.10</b>
Residual	<b>236893.616</b>	<b>53</b>	<b>4469.69087</b>	Prob > F	=	<b>0.0000</b>
				R-squared	=	<b>0.3698</b>
				Adj R-squared	=	<b>0.3579</b>
Total	<b>375916.436</b>	<b>54</b>	<b>6961.41549</b>	Root MSE	=	<b>66.856</b>

foodexp	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
totalexp	<b>.4368088</b>	<b>.0783226</b>	<b>5.58</b>	<b>0.000</b>	<b>.2797135</b>	<b>.593904</b>
_cons	<b>94.20878</b>	<b>50.85635</b>	<b>1.85</b>	<b>0.070</b>	<b>-7.796134</b>	<b>196.2137</b>

b) The residuals obtained from this regression look as follows:



Plotting residuals (R1) against total expenditure, we observe



It seems that as total expenditure increases, the absolute value of the residuals also increase, perhaps nonlinearly.

c)

White's general test statistic : 7.374513 Chi-sq( 2) P-value = .025

H0= Homoscedasticity

H1= Otherwise

White's general test statistic : 7.374513

Critical Chi-sq(2) = 5.9914 (5% significance level)

Reject H0

There is enough evidence to say that there is heteroscedasticity

d) White's heteroscedasticity-consistent standard errors

Compared with the OLS regression results given in (a), there is not much difference in the standard error of the slope coefficient. although the standard error of the intercept has declined.

Linear regression

Number of obs = 55  
 F(1, 53) = 34.60  
 Prob > F = 0.0000  
 R-squared = 0.3698  
 Root MSE = 66.856

foodexp	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]	
totalexp	.4368088	.0742544	5.88	0.000	.2878733	.5857442
_cons	94.20878	43.26305	2.18	0.034	7.434094	180.9835

## 12.26

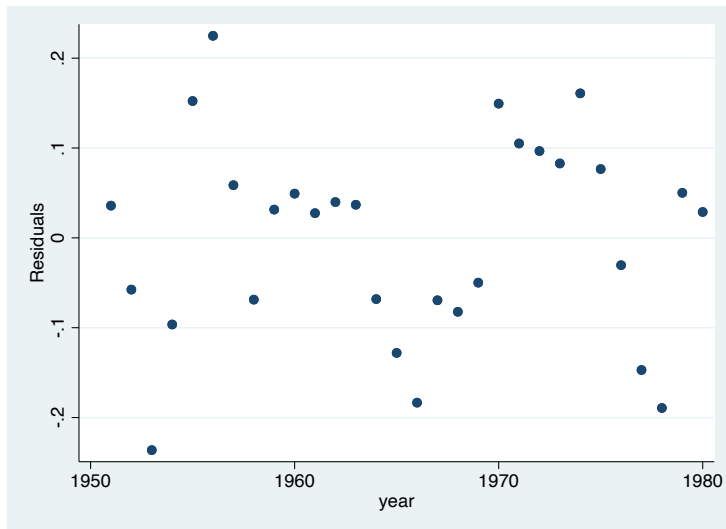
a)

Source	SS	df	MS	Number of obs =	30
Model	5.42774163	4	1.35693541	F(4, 25) =	91.54
Residual	.370572985	25	.014822919	Prob > F =	0.0000
Total	5.79831462	29	.199941883	R-squared =	0.9361
				Adj R-squared =	0.9259
				Root MSE =	.12175

lnc	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
lni	.4675086	.1659868	2.82	0.009	.1256524	.8093647
lnl	.2794423	.1147258	2.44	0.022	.0431602	.5157244
lnh	-.0051515	.142947	-0.04	0.972	-.2995564	.2892534
lna	.4414491	.1065083	4.14	0.000	.222091	.6608071
_cons	-1.500441	1.00302	-1.50	0.147	-3.5662	.5653184

Interpret the meaning ..... (Hint: Double log regression.... Must use the correct unit of analysis to receive full credits)

b)



c)

$H_0$  = No positive autocorrelation

$H_1$  = Otherwise

$H_0^*$  = No negative autocorrelation

$H_1^*$  = Otherwise

Durbin–Watson d-statistic (5, 30) = .9549404

From Durbin–Watson table

dL=1.071

dU=1.833

4-dU=4-1.833

4-dL=4-1.071

Reject  $H_0$  = No positive autocorrelation

There is enough evidence to say that there is positive first-order autocorrelation.