

Lecture 21
CHAPTER 9: Dummy Variable Regression Models

In the previous chapter, the dependent and independent variables in our multiple regression models have had **quantitative** meaning. For example, the salary of CEO, annual firm sales, return on equity in percent, and return on firm's stock. In each case the magnitude of the variable conveys useful information.

However, in the empirical work, we must also incorporate **qualitative factors** into regression models. The gender or race of an individual, the industry of a firm (manufacturing, retail, and so on), and the region in Thailand where a city is located (north, south, west, and so on) are all considered as the qualitative factors.

8.1 Describing Qualitative Information

Normally, qualitative factors often come in the form of binary information:

Example:

- [1] A person is female or male or female. GENDER
- [2] A firm offers a certain kind of employee pension plan or it does not. YES (1)
- [3] A firm is located nearby the dam or not. YES (1)
NO (0)

All of these examples, the relevant information can be captured by defining a **binary variable** or a zero-one variable.

In econometrics, binary variables are most commonly called **dummy variables**, although this name is not especially descriptive.

In defining a dummy variable, we must decide which event is assigned the value one and which is assigned the value zero.

Question: Why do we use the the values zero and one to describe qualitative information?

Answer: These values are arbitrary: any two different values would do. The real benefit of capturing qualitative information using zero-one variable is that it leads to regression models where the parameters have very natural interpretations.

QUANTITATIVE

Firm i	x_2 SALARY	x_3 YEAR	DUM1 (GENDER)	DUM2 (RACE)
1			0	0
2			0	0
3			1	0
4			0	1
5			1	0
6			0	0
7			0	0
8			1	1
9			0	0
10			0	0
...			1	1
...			1	1

(GENDER) $\left\{ \begin{array}{l} 0 \text{ IF } \text{♀} \\ 1 \text{ IF } \text{♂} \end{array} \right.$
 (RACE) $\left\{ \begin{array}{l} 0 \text{ WHITE} \\ 1 \text{ BLACK} \end{array} \right.$

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8.2 A Single Dummy Independent Variable

Suppose we would like to estimate the following simple model of hourly wage determination:

$$wage_i = \beta_0 + \delta_0 female + \beta_1 edu + u_i \quad n = 526$$

$$WAGE = \beta_0 + \delta_0 D_1 + \beta_1 edu + u_i$$

FEMALE IS A BINARY VARIABLE TAKING VALUE OF ONE FOR FEMALE AND VALUE OF ZERO FOR MALE

WAGE $D_1 = 1$ IF FEMALE
 $D_1 = 0$ IF MALE

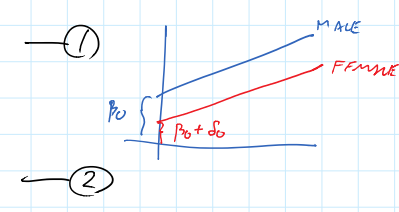
FEMALE = 1 IF THE PERSON IS FEMALE
FEMALE = 0 IF THE PERSON IS MALE

δ_0 REFLECTS WAGE DIFFERENTIAL BET. MALES AND FEMALE, GIVEN THE SAME LEVEL OF EDUCATION

IF $\delta_0 < 0$, IT IMPLIES THAT, ON AVERAGE, FEMALES EARN LESS WAGES THAN MALES, GIVEN THE SAME LEVEL OF EDUCATION.

$$E(WAGE | FEMALE = 1, EDU) = \beta_0 + \delta_0 \cdot 1 + \beta_1 edu = (\beta_0 + \delta_0) + \beta_1 edu$$

$$E(WAGE | FEMALE = 0, EDU) \text{ (MALE)} = \beta_0 + \delta_0 \cdot 0 + \beta_1 edu = \beta_0 + \beta_1 edu$$

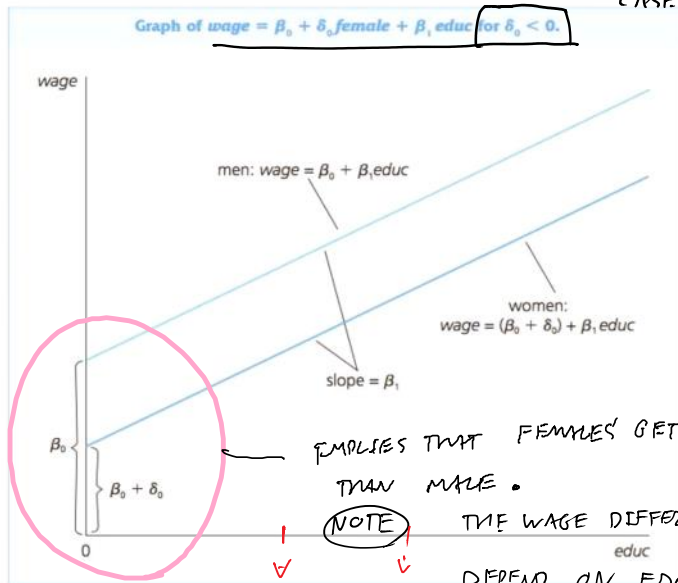


① - ②;

$$E(WAGE | FEMALE = 1, EDU) - E(WAGE | FEMALE = 0, EDU) = \delta_0$$

GIVEN THE SAME LEVEL OF EDUCATION, WAGE DIFFERENTIAL IS DUE TO GENDER ONLY!

Figure 9.1: Graph of Wage



IMPLIES THAT FEMMES GET LOWER HOURLY WAGE THAN MALE.

NOTE THE WAGE DIFFERENTIAL DOES NOT DEPEND ON EDUCATION.

THIS EXPLAINS WHY THE WAGE-EDUCATION PROFILES OF THE TWO ARE PARALLEL.

$$WAGE = \beta_0 + \delta_0 FEMALE + \beta_1 EDU + u_i$$

MRIE IS BENCHMARK GROUP

FEMALE = 1 IF FEMALE
FEMALE = 0 IF MALE

VS. FEMALE = 0 IF FEMALE
FEMALE = 1 IF MALE

THE GROUP THAT WE ASSIGN "0" IS CALLED "BENCHMARK GROUP" OR "BASE GROUP"

VS. MALE = 1 IF MALE
MALE = 0 IF FEMALE

FEMALE IS USED AS BENCHMARK GROUP

THE GROUP THAT

Lecture Note: EE 325-2/2015: Introductory Econometrics—page—148

YOU MAKE A COMPARISON

AGAINST IT

(EX)

$$WAGE = \alpha_0 + \gamma_0 MALE + \beta_1 EDU + u_i$$

MALE = 1 IF MALE
MALE = 0 IF FEMALE

FEMALE IS USED AS BENCHMARK GROUP

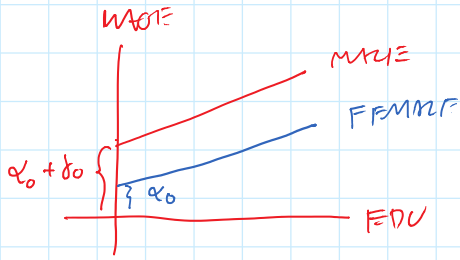
$$E(WAGE | MALE = 1, EDU) = (\alpha_0 + \gamma_0) + \beta_1 EDU \quad \text{--- ①}$$

$$E(WAGE | MALE = 0, EDU) = \alpha_0 + \beta_1 EDU \quad \text{--- ②}$$

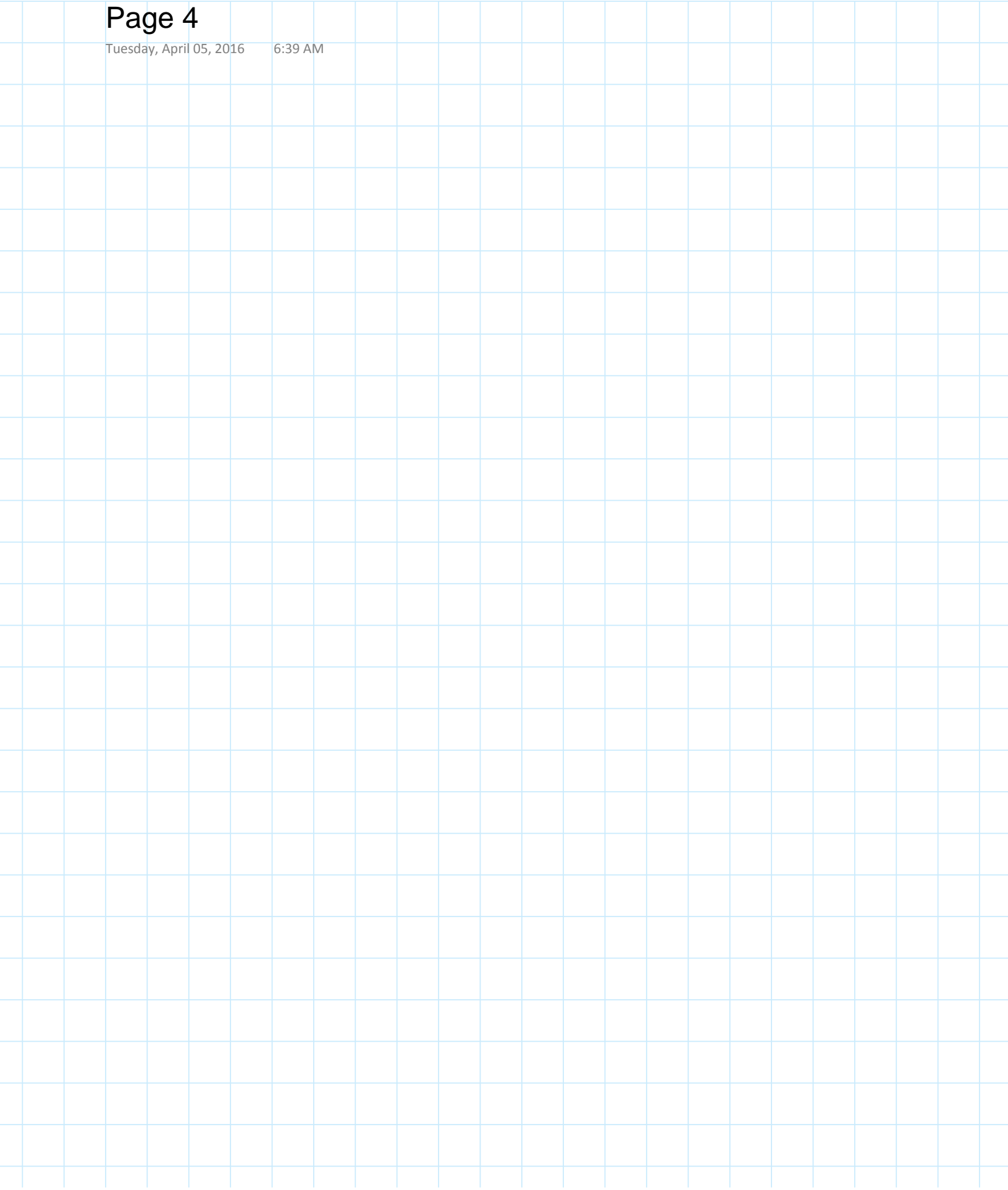
$$\text{①} - \text{②} \quad E(WAGE | MALE = 1, EDU) - E(WAGE | MALE = 0, EDU) = \gamma_0$$

$$E(\text{WAGE} \mid \text{MARRIE} = 0, \text{EDU}) = \alpha_0 + \beta_1 \text{EDU} \quad (2)$$

$$(1) - (2) \quad E(\text{WAGE} \mid \text{MARRIE} = 1, \text{EDU}) - E(\text{WAGE} \mid \text{MARRIE} = 0, \text{EDU}) = \gamma_0$$



INTERCEPT FOR MALE'S REGRESSION LINE
 $\alpha_0 + \delta_0$
 α_0
 INTERCEPT FOR FEMALE'S REGRESSION LINE



Now, we added more variables into the wage model. Taking males as the base group, a model that controls for experience and tenure in addition to education is

$$wage_i = \beta_0 + \delta_0 \text{female} + \beta_1 \text{edu} + \beta_2 \text{exper} + \beta_3 \text{tenure} + u_i$$

If edu, exper, and tenure are all relevant productivity characteristics, the null hypothesis of no difference between men and women (No wage discrimination) is:

$$H_0: \delta = 0 \quad (\text{NO WAGE DISCRIMINATION})$$

$$H_1: \delta < 0 \quad (\text{WAGE DISCRIMINATION})$$

In table 9.1, it represents the partial listing of the sample data of wage model. We see that Person 1 is female, Person 2 is female, Person 3 is male, and so on.

Table 9.1: A Partial Listing of the Wage Data.

	wage	educ	exper	tenure	female
1	3.1	11	2	0	1
2	3.2	12	22	2	1
3	3	11	2	0	0
4	6	8	44	28	0
5	5.3	12	7	2	0
6	8.8	16	9	8	0
7	11	18	15	7	0
8	5	12	5	3	1
9	3.6	12	26	4	1
10	18	17	22	21	0
11	6.3	16	8	2	1
12	8.1	13	3	0	1
13	8.8	12	15	0	0
14	5.5	12	18	3	0
15	22	12	31	15	0
16	17	16	14	0	0
17	7.5	12	10	0	1
18	11	13	16	10	1
19	3.6	12	13	0	1
20	4.5	12	36	6	1
21	6.9	12	11	4	1
22	8.5	12	29	13	0
23	6.3	16	9	9	1
24	.53	12	3	1	1
25	6	11	37	8	1
26	9.6	16	3	3	0
27	7.8	16	11	10	0
28	13	16	31	0	0
29	13	15	30	0	0
30	3.3	8	9	1	1
31	13	14	23	5	0
32	4.5	14	2	5	1
33	9.7	13	16	16	1

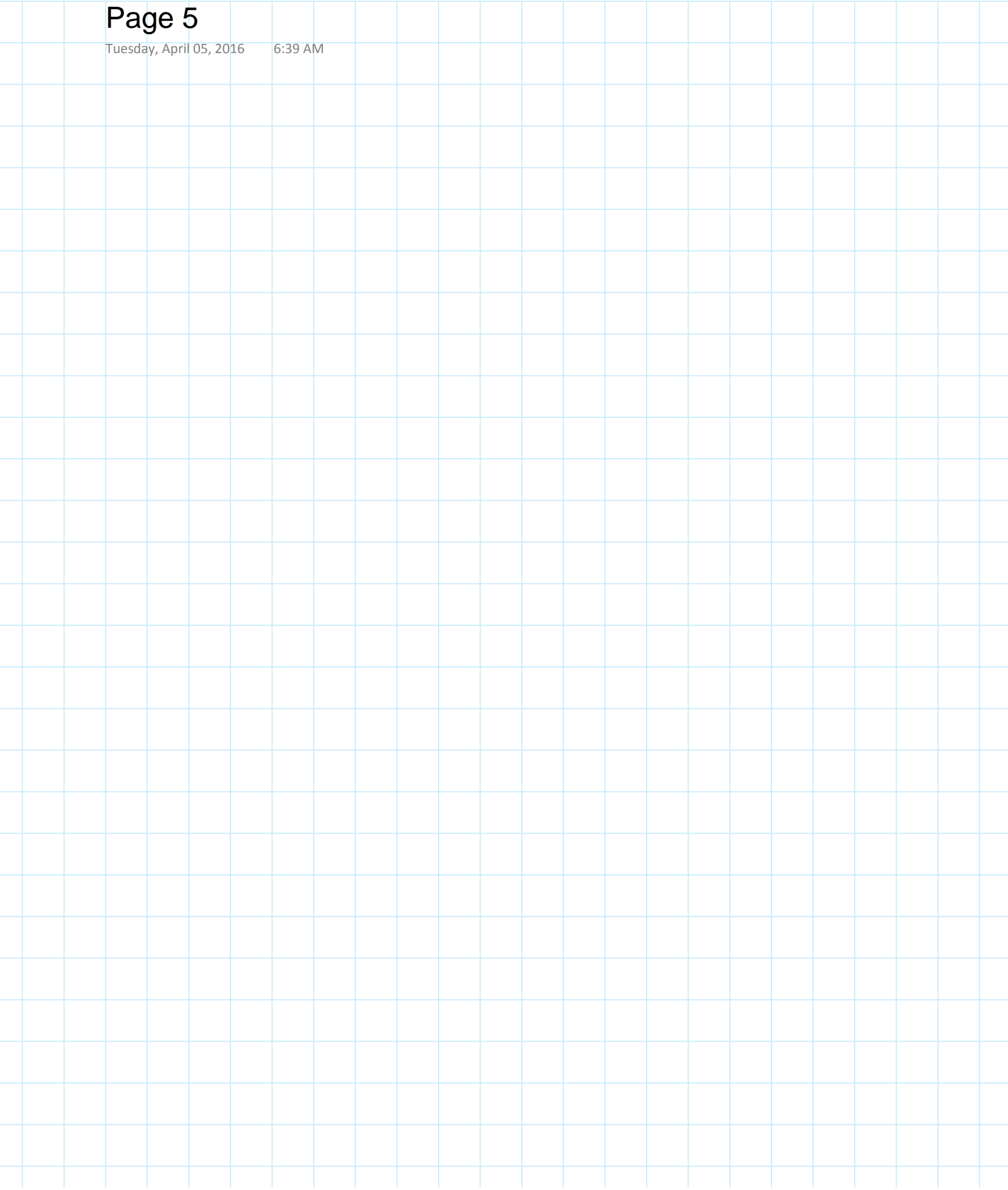
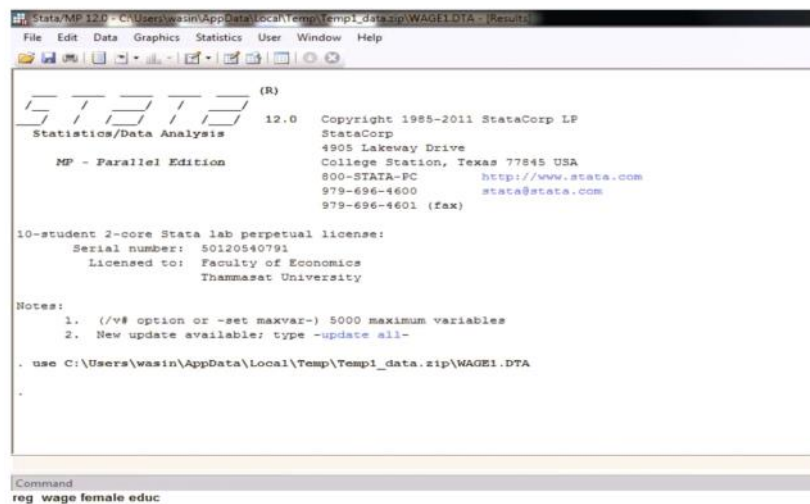


Table 9.2: The command function to estimate the wage model in STATA program



The screenshot shows the STATA software interface. The title bar indicates the file path: C:\Users\wasin\AppData\Local\Temp\Temp1_data.zip\WAGE1.DTA - [Read-Only]. The menu bar includes File, Edit, Data, Graphics, Statistics, User, Window, and Help. The main window displays the STATA logo and version information: STATA (R) 12.0, Copyright 1985-2011 StataCorp LP, StataCorp, 4905 Lakeway Drive, College Station, Texas 77845 USA, 800-STATA-PC, http://www.stata.com, 979-696-4600, stata@stata.com, 979-696-4601 (fax). It also shows a 10-student 2-core Stata lab perpetual license for Faculty of Economics at Thammasat University. A notes section lists: 1. (/v# option or -set maxvar-) 5000 maximum variables, 2. New update available; type -update all-. Below this, a command window shows the command: . use C:\Users\wasin\AppData\Local\Temp\Temp1_data.zip\WAGE1.DTA. At the bottom, a Command window shows the command: reg wage female educ.

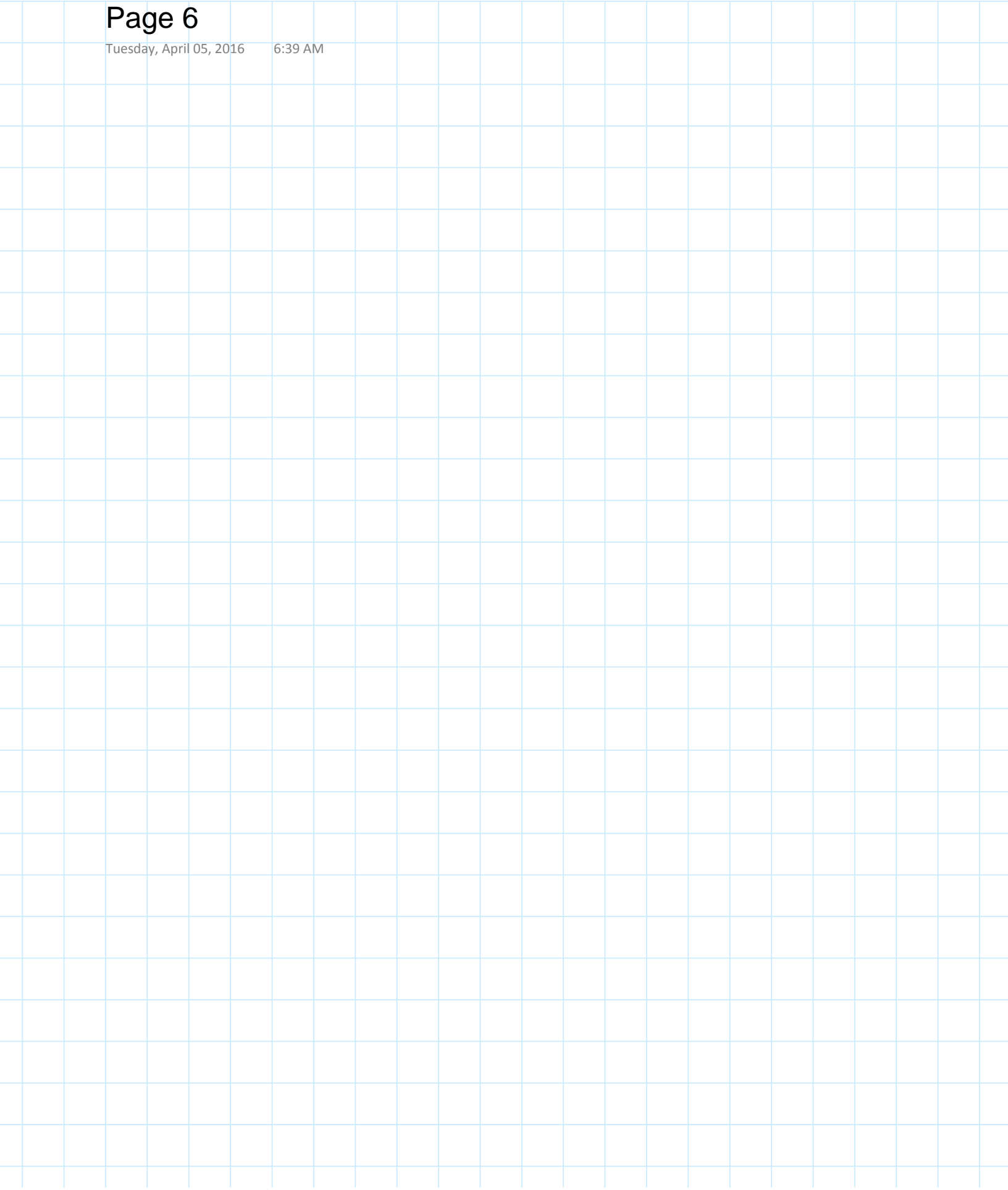


Table 9.3: $wage_i = \beta_0 + \delta_0 \text{female} + \beta_1 \text{educ} + \beta_2 \text{exper} + \beta_3 \text{tenure} + u_i$

. reg wage female educ exper tenure

Source	SS	df	MS			
Model	2603.10658	4	650.776644	Number of obs =	526	
Residual	4557.30771	521	8.7472317	F(4, 521) =	74.40	
Total	7160.41429	525	13.6388844	Prob > F =	0.0000	
				R-squared =	0.3635	
				Adj R-squared =	0.3587	
				Root MSE =	2.9576	

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-1.810852	.2648252	-6.84	0.000	-2.331109	-1.290596
educ	.5715048	.0493373	11.58	0.000	.4745802	.6684293
exper	.0253959	.0115694	2.20	0.029	.0026674	.0481243
tenure	.1410051	.0211617	6.66	0.000	.0994323	.1825778
_cons	-1.567939	.7245511	-2.16	0.031	-2.991339	-.144538

Table 9.4: $wage_i = \beta_0 + \delta_0 \text{female} + u_i$

reg wage female

Source	SS	df	MS			
Model	828.220467	1	828.220467	Number of obs =	526	
Residual	6332.19382	524	12.0843394	F(1, 524) =	68.54	
Total	7160.41429	525	13.6388844	Prob > F =	0.0000	
				R-squared =	0.1157	
				Adj R-squared =	0.1140	
				Root MSE =	3.4763	

wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
female	-2.51183	.3034092	-8.28	0.000	-3.107878	-1.915782
_cons	7.099489	.2100082	33.81	0.000	6.686928	7.51205

Example: the Hourly Wage Equation:

$$\begin{aligned} \widehat{\text{wage}} &= -1.5679 - 1.8109 \text{ female} + 0.5715 \text{ edu} + 0.025 \text{ exper} + 0.141 \text{ tenure} \\ &= (0.7246) \quad (0.2648) \quad (0.0493) \quad (0.0116) \quad (0.0212) \end{aligned} \quad (\text{Eq.32})$$

$R^2 = 0.3635 \quad n = 526$

Interpret the model:

The intercept: -1.5679 IS NOT SO MEANINGFUL AS NO ONE HAS ZERO OR NEGATIVE WAGES IN THE SAMPLE.

The coefficient on female -1.8109 => SHOWS THE AVERAGE DIFFERENCE IN HOURLY WAGES BETWEEN MEN AND WOMEN, WITH THE SAME LEVEL OF EDUCATION, EXPERIENCE, AND TENURE!

[EDU, EXPER, TENURE ARE CONTROLLED VARIABLES] COMPARED W/ MEN (BASE GROUP), WOMEN RECEIVE 1.81 \$ LESS PER HOUR THAN MEN, ON AVERAGE.

It is informative to compare the coefficient on female in the above equation to the estimate we get when all other explanatory variables are dropped from the equation:

$$\begin{aligned} \widehat{\text{wage}} &= 7.0995 - 2.5118 \text{ female} \\ \text{se} &= (0.2100) \quad (0.3034) \end{aligned} \quad (\text{Eq.33})$$

$R^2 = 0.1157 \quad n = 526$

MODEL ①

MODEL ②

-2.51183 = AVERAGE WAGE DIFFERENTIALS BETWEEN MALES AND FEMALES

$$E(\text{WAGE} | \text{FEMALE} = 1) = 7.0995 - 2.51183 \cdot (1) = 4.59 \text{ \$ PER HOUR}$$

$$E(\text{WAGE} | \text{FEMALE} = 0) = 7.0995 - 2.51183 \cdot (0) = 7.0995 \text{ \$ PER HOUR}$$

(MALE)

SO $E(\text{WAGE} | \text{FEMALE} = 1) - E(\text{WAGE} | \text{FEMALE} = 0) = 4.59 - 7.0995 = -2.51183 \text{ \$ PER HOUR!}$

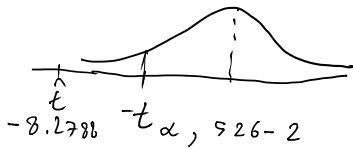
LET'S DO HYPOTHESIS TESTING

$H_0: \delta_0 = 0$ (NO WAGE DISCRIMINATION)

$H_1: \delta < 0$ (WAGE DISCRIMINATION)

- COMPUTE \hat{t} : $\hat{t} = \frac{\hat{\delta}_0 - \delta_0}{SD} = \frac{-2.51183 - 0}{0.3034} = -8.2788$

- FIND t_{CRITICAL} : $t_{\text{CRITICAL}} = t_{0.05, 524} = 1.645$



- COMPARE \hat{t} w/ t_{CRITICAL} :

SINCE $-\hat{t} < -t_{\text{CRITICAL}}$, THEN WE REJECT THE NULL HYPOTHESIS.

WAGE DIFFERENTIAL BET. MALES AND FEMALES IS STATISTICALLY SIGNIFICANT AT 95%.

IN PARTICULAR, COMPARED W/ MALES, FEMALES EARN 2.51 \$ LESS, ON AVERAGE.

MODEL 1

$$\text{WAGE} = f(\text{FEMALE}, \text{EDU}, \text{EXPER}, \text{TENURE})$$

$$\hat{\delta}_0 = -1.81$$

↓
MODEL (I) IS THEN MORE RELIABLE

MODEL 2

$$\text{WAGE} = f(\text{FEMALE})$$

$$\hat{\delta} = -2.51$$

↓
HIGHER B/C WE DO NOT CONTROL FOR DIFFERENCES IN EDU, EXPER, TENURE

